

**MORPHOMETRIC ANALYSIS OF THE BRACHYDONTOMYCARUS (Copepod: Calanoidae) AN  
ASSIGNMENT OF PHYSIOLOGICAL CLITCH BASED SIZE, SEX, AGE, AND STAGE OF DEVELOPMENT**

**By**

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To Charles, Boris, and Chris,

with whom it has been wonderful to share the last thirty years.

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*Abstract of Dissertation Presented to the Graduate School of the University of Florida in Partial Fulfillment of the Requirements for the Degree of Doctor of Philosophy*

**MECHANISTIC ANALYSIS OF THE BROAD-SPECTRUM CHAMAIN (CLAMOR) CRYPTIDAE IN A  
RECONSTRUCTION OF INDIVIDUAL CLAMOR BODY SIZE, SEX, AGE, AND AGE-RELATED GROWTH**

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Three mechanistic questions are tested in the current morphometric study of the broad-spectrum chamain (*Clamora laminae*). Is there a significant morphometric variation related to *clutch effort*, *clutch size*, *sex*, *age*, *size of origin*, and *environment*? Can the groups to which a specific individual belongs be *predicted*? And last, if it is possible to efficiently predict an individual's age by using all morphometric variables measured, is it possible to reduce the number of variables without significantly losing efficiency? Cranial and body morphometric variables were measured from captured and wild samples. MANOVA, linear discriminant analysis with stepwise selection (PCA analysis) select regressions were used in order to answer the questions above. Exploratory studies of relative growth and efficiency of reproduction were used to develop models of age estimation, sexual dimorphism, and clutch discrimination.

Log transformation of the significantly improved group description progresses through the reduction of the scaling effort among variables. Except for body mass, log transformation did not significantly improve data analysis or relative growth studies. Ratios between variables were consistently less efficient at detecting differences between groups than direct measurements. Principal axis morphometric models are consistently more efficient than separate growth curves for determining age of animals. Sexual dimorphism in cranial morphology of broad-spectrum-chamain can be detected even in young individuals. However, it is frequently impossible to obtain greater than edge measures of variation, such as *clutch size*, *body*

ness, sense of touch, and orientation). There is a positive correlation between female body size and longevity rate, headlength/body size, and clutch size. There is a significant negative correlation between female body size and the relative clutch mass (clutch mass / female body mass). Clutch discrimination based on clutch effort and female body effort (perceived energy) can be used to establish pairings among individuals in small wild populations and captive breeding programs. There is a significant overall complementarity variation among ovaries from different study sites. This may be related to some level of reproductive isolation, which may be related to habitat fragmentation. Captive ovaries present relatively broader mass, narrower, round tip, and rounded apex, and relatively longer medullary cytoplasm than wild individuals.

## CHAPTER 1 INTRODUCTION

The Inland emerald cormorant (IEC) is a small seabird South American cormorant. The species exhibits morphological associated with the Paraná and Río Paraguay River systems in Brazil, Paraguay, Bolivia, Uruguay, and Argentina, and after the small coastal near changes from the northernmost limit to southern Uruguay (Knowledge 1983, 1987). In some subpopulations, the name *myzomela alba* *pygmaea* ("yellow throat cormorant") in Brazilian Portuguese and *pygmaea alba* *pygmaea* ("small cormorant") in the Spanish language cormorants share it here. As in English and Spanish cormorants species, its most conspicuous characteristic is the black head, relatively the smallest among all cormorants (King and Butler 1987).

The State of Rio Paulo, Brazil, comprises the central region of the bird geographic distribution. However, it has lost approximately 95% of its original ecosystems (Rauer 1973). The Tied River – the main river system of Rio Paulo – has been virtually completely drained for the construction of large electric dams and to have its main upstream tributaries (Rauer 1984). The original wetland wetlands of Rio Paulo – now severely reduced – has possibly served as the linkage between northern and southern populations of the species. The local extinction of IEC in Rio Paulo could have a deleterious effect for the species as a whole. On the other hand, cormorant apparent capacity to tolerate man-made habitats may be significantly important for their conservation (Vanbelle and Lawrence 1986) as in Costa Rica (Rauer 1987). Long-term studies of this species might help us to understand how cormorants respond to environmental changes and how can we deal with similar situations in the future.

As Condit (1984) elegantly stated, conservation biology has two main subjects: the declining population paradigm which deals with the cause of localities and extinction, and the small-population paradigm which deals with the effect of smallness on the persistence of a population. The first seems to be the best approach to the study of the extirpation process of IEC in our available habitats. The cause of extinction is unknown at present, but its effects on the persistence of the populations are of major concern. How does the extirpation process affect distributional strategy of the species like independence and nesting system? How is the population

structure and dynamics) affected during the restoration process in terms of sex ratio, age structure, and sexual maturity? All these variables would affect the persistence of the population and thus require specific solutions.

Traditionally, long-term field studies are required to understand the factors of a population structure and dynamics (Loren 1987). Time and budget resources are therefore often overwhelming issues. Molecular techniques have been recently developed for yield sampling and referring the results not only to genetics but also to behavioral ecological studies of wild species (Arlow 1994). However, although they were not decreasing, they are still in preliminary when large sample sizes are used. In addition, expensive equipment and specialized personnel are needed.

Establishing the sex and age of an individual is the first step in the definition of the sex ratio and the age structure of a population (Chambers and Pollock 1984). The recent development of sex (usually) statistical software has permitted considerable advances in the use of multivariate statistical analysis in morphometric studies (Reynard et al. 1996). Intraspecific analysis of sexual dimorphism and morphometric structure using such techniques can be extremely useful in field studies of small populations. When secondary sexual characters are not visible, the use of morphometric analysis of sexual dimorphism can be essential in the determination of the population sex ratio (Giles and Giles 1983). Similarly, morphometric and allometric analysis of growth can be used for the development of age-structure models for individual animals.

Length measurements can be used by direct observation of the sex organs in the elasmobranchs (Chambers 1980, Ruvinsky 1988, Adams and Laing 1990). Sexual dimorphism can also be used in the sex discrimination of preserved resources (Walt and Pollock 1985). Sex is sometimes determined by the temperature of incubation (Porgans and Loren 1982, Walt and Laing and Anderson 1994) with an sex-temperature based in cytological studies (Chen and Chen 1978). Temperature of incubation may also affect embryonic development (Loren and Porgans 1987), incubation period (Walt et al. 1987), post-hatching growth pattern (Gordon et al. 1987, Walt and Cooper-Petersen 1989), regeneration, and down-regulatory behavior of hatchlings and young (Zlotnick and Porgans 1989).

However, there seems to be a pronounced variation caused by the clutch of eggs in the pattern of the temperature-size determination called "clutch effect" by Laing and Anderson (1994) and possibly in the growth pattern of hatchlings and young. If not, these animals from different clutches would exhibit a significant variation not only in body size but also in body shape as such

a way that offspring from different families could be distinguished through the analysis of body size and shape of the individual specimens. The clutch effect in this case would compare not only genetic but also phenotypic components such as maternal or environmental and maternal behavior and rearing status. Analogously, relatively isolated small populations could be distinguished by cluster analysis. Although transverse growth rates are commonly variable even within single populations, age estimates from growth curves can be useful for gross stratigraphic analysis (Plafon 1967). Microchemical analysis must be used for more detailed work, such as the study of age specific feeding or when post-mortem contamination is feasible (Shogren 1984).

Morphometric, analytical and relative growth models are proposed here as tools for the study of FWC, clutches and populations. Discrimination and individual's body size, sex, and age estimation. Once increasing clutches and populations might help us to understand dispersal patterns and the colonization capacity of the species. Individuals' body size, age, and sex measurement – even from preserved materials – might help researchers understand the structure of the target populations. The low cost and relative simplicity of morphometric analysis might contribute to an increasing set of alternative methods for the study of small populations.

## CHAPTER 2 LITERATURE REVIEW

Morphometrics have been defined in different ways. The use of multivariate statistical analysis in morphometrics made Bookstein and Rymer (1981) call it a “Mathematical Morphometrics” and define this field as “the application of multivariate statistical analysis to biological variability as morphological characters”. Goswami (1990) extended this definition a little and called it like “the measurement of biologically relevant forms and patterns in ways that allow their quantitative handling”. In opposition to the multivariate statistical methods, a geometrical approach was brought to the scene by Bookstein (1991) who defines morphometrics as the “mathematical theory of geometry with biology”, “the study of the geometrical forms of organisms” (Bookstein 1994), or “the statistical study of biological shapes and shape changes” (Bookstein 1991). Complementarily to him, Bookstein (1991) defines this field as “the quantitative description, analysis, and interpretation of shape and shape variation. In response to the “geometrical school”, Rymer (1991) states that multivariate morphometrics is like “application of multivariate statistical analysis to biological variability as morphological characters, and also may include the shape variation” (Goswami 1990b).

Later graphically morphometrics can be defined as “the measurement of external form of living organisms”. This word derives from the Greek *morphe* (form) and *metron* (measure). “*Morphe*” can be defined as a standard of measurement, whereas the *inflections of forms* is less straightforward “shape information of something”, “external appearance of a body”, “outline and dimensions are defined by contour” or as “external aspect of a thing” (Oxford English Dictionary 1994). More than defining the exact role of shape and shape variation and the best approach to its analysis, what seems to really intrigue morphometricians is the old question of how to describe the nature of the forms of living organisms.

## 2.1 Wilhelm Drieschner: The Search for the Essence of the Fungus

The issue of living organisms has been preoccupying Drieschner as both scientist and layman. From a broad to narrow and narrow to broad, he went on more than eight and equally, attempts for doing scientific history. Human is individuality, capacity of observation, the great variety of the forms of living organisms. Linked to the simple necessity of recognizing their production and proper maintenance from forests on rocks, humid conditions on valleys, living up complex networks, on an infinite journey in search for the essence of the form of living organisms.

Pythagoras of Samos believed that all things could be described by numbers, which means that their extension and movements could be determined by finding the numerical relations corresponding them. Essence of form was, however, the key to solve the ideas of Plato. According to him, all of the things that were generated with three points, including living organisms, seemed to be superior copies of created from the moving, timeless and eternal. Although lacking Plato's abstract of perfectness, Aristotle perceived that all living beings develop from an imperfect to perfect state, after which they again decay and finally die, leaving imperfect themselves (Baker 1925). This idea left a great essence of the living organism is form around which evolution can be seen as more sophisticated. Including enough or simple influences the way humans already living organisms.

Linnaeus employed Aristotle's system and to create his classification system. Operating within a logical and systematic framework, he could naturally find species specific good characters by which a species could be distinguished (Jain 1984-1989). Goethe explained the idea that the essential form of an organism is related to its function. In many ways he was ahead because the teleological Linnaean approach and the "genetic" (which organism) approach developed by Huxford (Bilalwan 1911).

The great variety of living organisms drives the naturalist Charles Darwin. However, he did not see variation as imperfections of the ideal organism. Instead, he considered variation as the potential for evolution – and – consequently – the origin of new forms – by natural selection (Darwin, 1859). In summary, an evolutionary theory did not fulfill Linnaeus' requirements more than superficially. There variation among species – primarily when he provided a basis to be seen as the result of his evolutionary process. However, the idea of quantitative information as the distinctive systematic of Linnaeus led to the development of an alternative system: the Numerical Taxonomy (Sokal and Rohlf 1969; Rohlf and Sokal 1973), which is still in dispute as taxonomic system.

Charles Darwin not only perceived the relationship between form and function but also between form and behavior of body organs too. For example, different feeding behavior would be related to different food sources and shape of beak. In this period not the general adaptive would be adapted to differences between the two body systems of form (anatomical) and behavior (Darwin (1871).

The evolutionary approach has contributed significantly to the origin of a new angle on the study of behavior is field called *ethology*. Although not inventors the relationship between form and behavior, the early ethologists' studies eventually reflected the idea that form and behavior would be the result of adaptation and evolution. The ethological-evolutional steps of Ivan Pavlov (1913) the fish-baiter steps of Niko Tinbergen (1951), the bearded-chickens of Edward Ross (1946), the integrated birds of Konrad Lorenz (1941), and the dancing bees of Karl von Frisch (1947), all showed evidence morphological constraints associated with behavior.

Quoting the observations of Signeuv Povel, "instinct is destiny" (Povel 1941 (1944)). The model for such a constant but preprogramed behavior laid for the relationship between form and behavior. Either a one-way way change his destiny – destiny set at its inevitable flow, but something not substantially brings from the past that deeply influences one's future. As an analogy, the interaction between form and behavior would be like a river. One could never against it go with it.

However, biologists are just biologists and they are ultimately had some false conclusions. This seems to have happened when some people in the 19th Century stated behavior the the shape of the beak's skull would be a strong predictor of their behavior. Even the prominent biologist Alfred R. Wallace (1884) was influenced by the belief that the construction of the human skull was reflective of mental faculties and character. the study of which was they called *Phrenology*. Although considered today as only a method of reading character from the shape of the skull the ideas of Thomas J. Gall, Johann G. Spurzheim and George Combe were well accepted as an experimental science with the system's assumption that mental phenomena first entered consciousness which could be observed (Darwin 1884).

From studying individuals based on their skull shape to inferring more easy visible characters is also they. Hanna moved phrenology was used and abused to such extent that defending the superiority of white over blacks and Indians as white was phrenologically called *Cranioscopy* another historical approach for biological Anthropology (Dewar 1914). Outgroup and reading different types of forehead to create systems of race literature was in the 19th but not widely accepted in England in the 19th Century. Although some of them cannot



had a universal focus. They still kept the idea of the extent of the form as a standard of perfection.

In his book *On Growth and Form*, D'Arcy W. Thompson (1877) introduced the idea that living organisms are shaped by physical constraints. His central argument was that surface-to-volume ratios were decisive in organisms grown in size. Therefore, according to him, small animals should be the product of surface forces whereas large organisms should be the product of gravitational (voluntary) forces. Here and elsewhere constraints to describe shape relationships between organisms. This could be called the primacy of the context and all concepts will influence the way biologists deal with size and shape of living organisms.

## 2.2 Contemporary Trends in the Describing of the cell shape

Current researchers, whether are named after D'Arcy's own system of processing mathematical points as a two-dimensional plane based surface distance from two adjoining lines, called axes. There are in the study of biological form transformations in which he coined as D'Arcy Thompson but it stems from the mathematician of (Gömbösi, 1933). Although D'Arcy Thompson did not present mathematical substantiation, his intuitive method of exploring shape transformations in living organisms as derivatives of Cartesian coordinates as self-consistent almost by every biologist. Only decades later, from the late 1970's to the late 1990's, Thompson's transformation grids were specialized in a statistical way. Geometric studies were generally stimulated by his *On Growth and Form*, especially the concepts of differential growth as introduced by Huxley (1924 and 1932).

Huxley was probably the first to systematically study differential growth rates; this leads organisms how they grow into place as different axes in different functions and a different points. He analyzed many examples of the geometrical specific organs as related to the growth of the body. He described the relationship within morphometrical expressions

$$F = aL^b$$

where  $F$  is the area or weight of the differentially growing organ,  $F$  is the standard weight of the animal (denoting the size or weight of the organ as given), and  $a$  and  $b$  are constants.

Although using a different approach than the generation used by Thompson, Huxley also worked with size and shape transformations as living organisms. The relative growth of the organs studied was within the ratio as the animal's body. In other words, the proportion of the organs (or other parts of the body) is similar to the animal's body changed during the growth process.

Q is, the concept it was originally different from (e.g.). For this reason, the mathematical expression was called an *affine transformation* (from the Greek, affo, "affluent") i.e. equivalent to the above-mentioned (from the Greek, iso, "isolate") growth, which was similar, but deeper than (e.g.).

James and Huxley (1943) recognized that when authors have perceived and believe that certain organisms have noticeable change of relative size with increase of body size. Ponder (1910) described a so heterophasic growth and Osipov (1941) called it *dynamic growth*. However, these terms were unsatisfactory because of ambiguity in their own meanings (Huxley and Teague (1943). In addition, these studies were not quantitative analysis of differential growth as first proposed by Huxley (1910) (1912).

The studies of Thompson and Huxley led to a great number of biological studies concerning size and shape transformations of living organisms. They said the Author thus brought, if the concept of a single measurement as independent length, width, mass and constant specialized living organisms forms changed? the concepts of shape and size are same one to be not exactly this. Huxley used single length measurements to body size, although he knew body size could be more adequately defined as weight. Thompson, by his own proposal, were geometrical approach to the study of biological forms. However, living organisms do not fit well to triangles, circles, rectangles or squares. Therefore, mathematicians (Germany) could not be applied effectively to describe biological forms. This probably was the reason that Thompson failed in proposing comprehensive methods for pure and transformations. There were no comprehensive available for this purpose at that time! Huxley (1943) and Richards and Kermack (1943) were probably the first to try to accomplish these, not by themselves in a book called *Growth and Form: Simple Principles of Growth* (Richard Clark and Huxley (1943).

Huxley (1943) suggests four essential stages in the analysis procedure of shape transformations as follows as a function of age. "The treatment of each figure, and each set of measurements, etc. and finally, as a separate analysis object, the defining of a transformation between each each figure and an arbitrary chosen initial figure by means of a single function with different parameter values, the reduction of the parameters to their efficient number and the replacement of this parameter by a continuous function of a single variable parameter standing in age" (Quoted in the original). According to this procedure, each member of the series figure would be quantitatively through the comparison of a single value. Richards and Kermack (1943) went further and proposed the utilization of (single) parameter concepts like

length, was not evident in manuscript to incorporate the concept of allometric growth of body size to transformation, growth of Thompson.

Thompson's original development of body size used only two variables at a time. This is means needed to still efficiently capture the relationship between the variables involved, but are necessary among other variables not included. Multivariate statistical analysis developed during the 1930's and 1940's (Pearson 1904, Mahalanobis 1946, Rao 1948) could be useful on the problem of biological characteristics through the calculation of (Mahalanobis) distance among (one-way) populations, based on frequency variables. The *z*-value Pearson (1944) helped, he also believed that the statistical images could be broken down into two parts, size and shape.

According to Pearson, the shape distance would be related to the variation among differences between mean values of any two populations (which he called *z*-value), whereas size distance would be related to the ratio of the deviations of all characters studied. Thus, when the relationship between two objects or organisms was seen, they would, just vary in terms of size, but not in terms of shape. Pearson recommended the way to treat shape variation could be analyzed, but yet, insufficient help to clarify their vague concepts.

The allometric relationships between organism growth and physiological processes particularly metabolism was established by Huxley (1932) and was actually modified by Klander (1933, 1940, 1941) and developed extensively by McNair (1943, 1946, 1970, 1971) and others (Thomson 1986, Solazzi-Nelson 1984). These studies linked the morphological effects of changing in body size to physiology and ecology.

Kempner, Rat and Lewontin (1968) studied the biological significance of the allometry on the allometric equation of Huxley (in rat & others). They concluded that the body presents a general field of growth potential and that changes in proportions are the result of the varying expression of this potential. They also suggest that the size of organs instead of absolute dimensions would become efficient to express variations on the intensity of this growth potential. Thus it is clear, although not constant reference to the differentiation of size ("absolute dimension") and shape ("proportional"), but they are not yet well defined. Only when Johnson and Johnson (1971) used principal components analysis on the study of size and shape variation in the gelatinous turfs or that size and shape could be clearly just statistically support.

The importance of principal components analysis was also developed by Karl Pearson (1901). According to Mady (1994-95), Pearson "expressly believed that this was the correct solution to some of the problems that arose of nature's 'transformation' at that time, although he did not propose a practical method of calculation for mean distances or their variables". The

statistical properties of principal-component were described later by Hotelling (1933). A principal component analysis tries to explain the variance covariance structure of a group of variables through a few uncorrelated linear combinations of them. Geometrically, these linear combinations represent the solution of a new coordinate system obtained by rotating the original system (Johnston and Wilks, 1970).

Using three lifetime variables of growth curve's response, Johnson and Wilks (1954) found out that the first principal component corresponds to a variation of size (i.e., growth rate), whereas the second and third principal components correspond to a variation in shape of the response structure (i.e., shape variation). In other words, if there is only one variation among individuals'  $x_i$ , if individuals are similar or if they keep the same shape during the growth process, there is only one principal component representing for the total response, and all of its coefficients present the same sign. On the other hand, whenever there is also a shape variation, other principal components result from the analysis, and they present positive and negative coefficients, which means that variation between lifetime variables also provokes a lateral movement of response. Variance among them the original between different sets of variables can be interpreted as variation of shape variation. Rao (1973) gives a mathematical proof for this argument.

The new approach seemed to be so interesting that Tremblay (1983) and Johnson (1983) proposed the multivariate generalization of the allometric equation based on the first principal component coefficients. Tremblay proposed the use of the correlation matrix of the log<sup>10</sup> transformed lifetime data because it tends to make principal components independent from the scale of magnitude and the scale of measurement of the variables. Correlation matrix is the covariance matrix for the standardized variables, by comparison where the covariance of variation between a response and a predictor is variables scaling could cause mathematical variance inflation, however, Johnson advocated that log<sup>10</sup> data transformation generally worked up to reduce growth data levels as well as make the growth curves independent from magnitude and scaling. Thus, his proposal was the best manner to generalize the allometric equation as the first principal component of the covariance matrix of log transformed lifetime data. Hopton (1984) and Igusa (1984) argued that this procedure was not completely satisfactory unless you could not correct the units of the remaining components, a well known to have been done by Johnson (1984) in a study of salmon run in the North American states (Atlantic salmon). However, in spite of log result statistically a significant step was done in this way. The study of biological relation growth curve's definitive multivariate approach

A few years before, Spearman (1904) had proposed a generalization of the allometric equation through the use of multiple linear regression. In this procedure, too, some problems, at least, do obstruct the correlation among variables. It may be affected by the nature of the movement, and it differs not only from all the others (as it depends on the number of responses), which prevents it from representing the various points in all sorts of relative steps between all pairs of variables.

The evolutionary, taxonomic, and physiological implications of allometric relations are extensively discussed by Huxley & Gould (1960). His criticism of linearity as a formulae cannot be resolved in any one form of mathematical expression such as the present function, as the study of proportion changes included wide variations in rate of value theoretical expression in the particular circumstances. In points not too regular may be morphological, physiological or chemical, and these differences may appear in analogy, phylogeny or in some comparison of related forms differing or not. He also suggests the use of factor analysis in the multivariate generalization of allometric relations.

Factor analysis is a complex multivariate statistical analysis. It also demands a set of variables or items of a small number of subjects or factors and then helps in elucidating the common relationships among these variables. However, all statistical analysis can never elucidate (Meady 1989). Factor analysis was developed by Charles Spearman (1904). He studied the contribution between two or more of various types and found that many degree of variations could be accounted for by a single model for the system, which accounts a significant relationship when Spearman, Karl Pearson and others used their procedure to define and measure intelligence (Johansen and Nielsen 1982).

Although apparently successful in the separation of size and shape components in factors, multivariate analysis were often successful in their definition. The problem possibly, neither Alfred Womersley's nor Spearman's approach was developed for the intermediate goals of Huxford. This was what Smith (1963) pointed. He proposed the following steps for the completion of the diagrams: most values of corresponding points are marked and degrees are marked in order to give the best possible fit. This would give measures of size and shape difference. The displacement of each point and its correspondence on the other diagram are then analysed in terms of linear, quadratic, cubic or higher transformation analysis. Corresponding points are called as homologous points, in the sense of the operational terminology used by Huxford (Smith and Smith 1963, Smith and Sokal 1971), which does not mean that they descend from a common ancestor. Although somewhat related to what was proposed by Huxford

(1943), this procedure was inspired from geologists who developed *areal surface analysis* and computer programs to study the regional and local trends of *contoured maps* (Crawford and Croftall 1981 and others cited by Smith).

Although *areal surface analysis* indeed improved the analysis of *point – not area* independent – shapes, its basic limitations for the study of biological phenomena that this procedure is restricted to *individuals' shapes*. The way around was made for a geologist, for whom a *single map* of a certain area may be enough (geographic area not shape!). However, biologists generally work with populations of individuals which vary and therefore require sampling, which is its own complex statistical analysis. Geometric analysis of *individuals' shapes* may be insufficiently *descriptive* for biologically continuous values that is looking for clusters of typical organisms. This could be claimed by any biologist who worked with growth or any other kind of quantitative variables in the late 1940's and 1950's. However, as a certain point, researchers used to look for clusters of typical organisms before the late 1980's and they kept doing so during the 1970's.

The use of computers and statistical software stimulated the use of *mathematical techniques* for the study of living organisms forms, which is that time already would be called as *morphometrics* (Bookchin 1984). Mathematical techniques were often *efficiently replace* statistical among individuals and continuous among variables. These procedures it for seemed to be efficient in deriving flat and shape components of biological measurements. However, they did not present any information about the geometry of the species studied, which was exactly the strength of Thompson's transformation grids. *Morphometrics* alone were sufficient to handle statistical variables.

The first possible attempt to combine *multivariate statistical analysis* with the *areal surface analysis* was done by Liemant (1949). He proposed the use of *canonical discriminant transformations* which seemed to him less more efficient to stimulate the use of *statistical analysis* itself than its application on Thompson's grids. *Canonical correlation analysis* is efficient for identifying and quantifying the associations between two sets of variables. It was initially developed by Hotelling (1935, 1936) as a technique to relate *withematic speed* and *power* to *studying spatial and power*. Its possible ecological applications are extensively discussed by Gower (1979). This technique relates *dimensionality* between two sets of variables into *subspace* of *canonical variables* (Johnson and Wichern 1992). The idea of Gower was that this technique could *efficiently study* the relationships between size and shape variables.

Other attempts to mathematically define (and test) sharpness were motivated by the idea of the multivariate generalization of the silhouette aspect, Wernersson (1975) defines sharpness vectors of size variables and then stated that sharp is actually sharp related to a variety of size variables. Gould (1971) states that the coefficients of the silhouette equations (see above) but is size-independent. Different between regressions as believed in the similarity of regression differences in size between comparable animals of similar shape. He also suggests that when the coefficient is significant the two regressions is evidence of geometric similarity and he extracted from the two variables.

The concept of geometric similarity identified by Gould was fundamentally influential future comparative studies between different organisms, although not exactly the way he proposed it is not restricted to the 11 variable silhouette equation. If many biologists take from similar organisms are the idea of similarity (or as between) forms is simple. Homology forms convey the most important information in evolution, phylogeny and other fields of biology. Increasing the number of variables and would intercorrelated measurements for comparison between different forms. Once again, multivariate statistical analysis would be the way to use how different (or how similar) they are most things do not really are. The problem Malvern was then added to the growing field of Morphometrics (Blackith & Rayman, 1971).

Initially Blackith (1965) describes morphometrics as a synthesis of geometrical measurement "a quantitative approach to measurement derived from the Pythagorean value that is geometrical relations by the concept of measuring". Later (however Blackith and Rayman (1971) suggest the statistical property and morphometrics could be considered as changing branches of the same field, possibly denominated "quantitative similarity". In addition they suggest that "another approach was required all those separate branches of morphometrics and to develop moderate changes of shape and size to filling out the provision of discriminatory biology (which clearly shows many of the techniques are based, directly or indirectly on the discriminant function" (Blackith and Rayman (1971:122).

They referred to the discriminant function analysis recently developed by Fisher (1936) to separate individual characteristics between or more known populations. Considering that this group of biologists are not motivated by morphological variations that have interest of taxonomic, but also to morphological variation (e.g. sexual dimorphism, ontogeny and many others). Blackith and Rayman were not wrong. The problem is that apparently they did not separate adequately the technique from the field (for biologists including human races).

multivariate statistical analysis is a set of methods that can be used to help detect or better understand biological processes that affect shape and size of living organisms. For a researcher interested in a specific group of morphological traits such as for a wildlife biologist interested in the sexual dimorphism of building skeletons. Practically, this was probably the major contribution of this branch of biological morphology: the detailed description of the most useful multivariate statistical analysis, even before described in the normal subjects literature (see Jones (Payson, Eklund, and Campbell 1987)).

Eklund and Raymond introduced size and shape as components of the form of living organisms, where transformation could be explained by principal components analysis as proposed by Johnson and Menzies (1981). They were typically skeptical in relation to the use of outlines or shape descriptors. They were advised that through the use of total surface analysis of transformations, point-to-point length deformations will represent the outline of a solid not that of a dissection, or Thompson's "stained" (Eklund and Raymond 1971: 108). However, not all biologists agreed with this advice. An analysis of lateral or transverse specimens were done, the center of typical gross can be easily seen, and therefore used to identify species. In this case, the outline of the profile itself could be the difference between species.

As an example of the above, Christopher and Waters (1974) proposed the use of Fourier maps to quantitatively describe complex shape. Fourier series (also called "Fourier analysis" or "harmonic analysis") is a mathematical expression of size and shape which can be applied to rough and irregular outlines. An abstract map and procedure was previously described in engineering mathematics texts (e.g., Page 1946) and others cited by Christopher and Waters. Other attempts present the use of Fourier analysis as used in various ways previously introduced by the development of new methods of computerized image sequences (Hild and Jander 1984 and Farris et al., 1982).

However, there was something missing in the concept of shape as just the outline of an organism and mainly on the concept that multivariate statistics could fully and only describe the geometry of shape (Humphries et al., 1981). According to Fred Bookstein (1991) it did not. According to him, both concepts lack the idea of landmarks, landmark structures used as points of variation including other structures (i.e., for specific comparison the use of Fourier analysis see Bookstein et al., 1982). The replacement of this simple concept revolutionized the field of morphometrics as no more than a decade.

According to Bookstein (1997a), shape and shape change should not be treated by discrete variables, vectors of great distances, sizes or other quantities, or, in short, as it used to be



by multivariate statistical analysis, instead according to him, shape measurements should be functional in order to provide geometric information about organisms form. He insisted that this was exactly what Percy Thompson proposed in 1917. However, he was mistaken that Percy Thompson did not develop a mathematical framework for this (Bookstein 1977b) the great failure this initiative, and his first attempt was to represent Thompson transformation by what he called *homologous grids* (Bookstein 1977a).

Homologous grids could express orthogonal even after transformation, contrary to Thompson's grids. These grids formally generalized the forms of classical geometric analysis in the context of two-dimensional Euclidean and Riemannian geometry. Shape change could be represented as a tensor's inner field upon every point of an object or organism, which means that shape change could be related geometrically to differential change in form (Bookstein 1978).

However, although homologous grids "provided shape comparison as a statistical coordinate system" described "new set-coordinates with regularizations of transformations" in any sub-region of mapping surfaces for the former *shapings*" (Bookstein 1978:12). In other words, this method was efficient to represent individual shape change, but it did not deal efficiently with variation among individuals in a statistical way. His first attempt to solve this problem was the use of triangles (Bookstein 1977a,b, 1979a,b, 1982) and true measures (Brown and Bookstein 1982, Bookstein et al., 1983) to describe shape and shape change. However, according to him these methods were not able to solve statistical problems concerning distribution theory (Bookstein 1982:29).

According to Bookstein, only when landmark locations were taken as the data themselves, it was possible to apply full statistical analysis to geometrical studies of biological shape, to which he called "morphometric systems" (Bookstein 1983). Goodall (1982) derived the equivalent  $P$  and  $W$  metric principal component analysis. Taking the "supershape" concept, i.e., Bookstein (1984b) introduced the shape coordinates for triangles and showed how shape differences can be weighted by homologous  $P$ . Contemporarily to them, Goodall (1984) extended the global shape space to which Goodall's and Bookstein's methods had separately applied to statistical matters in shape space. Their further publications (Bookstein 1986, Goodall 1985, Goodall 1987) reflected the convergence of their their approaches as a single foundation for the study of landmark data.

As is very further, Bookstein (1984a) proposed the use of thin-plate spline functions for landmark data to fit the differences in the positions of landmarks to that negative relative to their position in number. This means that after spline curves is made by using a spline that is fit to a

shape like or exactly like actual phenomena or objects whose size remains small is quantified via its "binding energy" (another analogy). The latter maps would be comparatively lower points, called principal maps, "which are geometrically independent, allow free deformation of progressively smaller geometrical scales" according to his vision.

After this, there were significant developments of new methods and software programs with a strongly geometric approach for morphogenesis analysis (Koball and Boudry 1998; Boudry 1999; Marois et al. 1999). Multiscale morphogenesis became "fractalized" morphogenesis" (Marois 1999), as reference to the "non-morphogenesis" (Koball and Marois 1999) or "geometric morphogenesis" (Laporte et al., 1999). Both approaches became increasingly more sophisticated than the well-known maps (Boudry et al., 1991) and a considerable flexibility with geometry became a main understanding of the new program (Boudry 1999).

According to Boudry (1999), one should not feel too inadequate if one is having trouble with geometric morphogenesis because "no matter how much experience you have in mathematics (statistical approaches, you will not think that geometric morphogenesis being going the first time you understand it, and even the second time – and perhaps the third time" (Appendix 10). The positive reason for this is the redundancy of there were about twelve mathematical methods used for professional computer.

Beyond this argument, there is still the same controversy about how to define the shape of living organisms. The so-called geometric morphogenesis was then the like so-called mathematics ("fractalized") simple curves (and – at least in part – diagrams, some division of typical organisms) (Boudry 1999: 114, with expressive definition of "non-math-shape" (Boudry 1999). Boudry, from the simplest geometrical shape – a triangle – can be extremely difficult to be defined by nonexperts of mathematics who, with background as Figure (see Marois 1999). A simple story of water-lapping from a shore, looks at looking off creating a flow (fractalized shape as complex than the calculation of the shape alone, according to Robert Shaw (1995) – a physicist specialist in chaos – as a "non-of-the-art computer technology" (Clack 1998).

Geometry morphogenesis is based to two principles: *fractalization* and *refinement* (Boudry 1999). However, the most geometric form of both depend on their superordinate. Approaching more mathematics (like fractal geometry) a theoretical infinite capacity to do so, it would be possible to see that theoretically it always proceed. A proofreading for this was developed by David Mandelbrot (1977) when he substituted and others "how long is the coast of Britain" and creating developing what is called *Fractal Geometry*. According to Mandelbrot, the answer

for the question: *Depends on the scale of the map* (i.e. getting more complexity) (How many dimensions it includes). According to (Garcá, 2007) "Muchelton found that as the scale of representations becomes smaller the required length of a contour that without bends, loops and protrusions resembling more smaller outlines and subcontours – at first decrease – starts to rise, where the picture then finally returns to zero. Perhaps <sup>10</sup> is also called to improve study representations of biological shapes, giving two complementary required lines drawn."

Exhibits to Reinventive Geometry: How there is an eye-see that they have arrived in French Germany

Geometric models are usually far more complex than real-world ones. However, it does not mean that they capture all reasons of being organisms: size and shape is what has been being called the structure of the form. Models and studies are useful and important because they help people to see what they have difficulty to visualize (Chernomskiy and Vrubel's, 1992). However, even the best "details of a typical organism" is just a sketch, which by definition may be simpler than the real organism. Otherwise, it would be much more to see the real organism itself being simpler means being an approximation of reality: as simple as possible, as complex as necessary. Multivariate morphometric models can be simple enough to understand the basis of morphological variation and complex enough to provide different descriptions of different groups or segments of organisms. This is the general goal of the present study.

## CHAPTER I MATERIALS AND METHODS

The present study relies on the development of morphometric models that can be applied to the assessment of the following: (1) individual identification (which sex, age, size of snout, and environment (aquatic or wild)) (for this purpose, body and head measurements were taken from 244 captives and 25 wild animals).

### 1.1 Study Sites

The captive animals were housed at *Escuela Superior de Agricultura - Univ. de Quito*, University of Quito, Pichincha, State of Quito, Pichincha, Ecuador. Information about their age, sex, date of birth, and perhaps an individual on the database exposed (stockbook) of the species (Ortiz and Jennings 1980; Vértiz and Molina 1995; Vértiz and Ramírez-Pérez 1991; and Vértiz and Berlin, in press). The wild animals were captured on small wetlands associated with tributaries of Tiro River in San Cayetano (Quito State) (Figure 1, Table 1).

Morphometric differences between wild and captive animals are evaluated in order to understand how exactly the captive environment influences anurals morphology. A similar study is being conducted by Myers (1996) with the American alligator, whereas his experimental conditions show the captive environment could be key element in crocodilians in reproduction and conservation programs.

### 1.2 Field Techniques

Field studies were carried out from October 1994 to May 1996. Captive individuals consisted of approaching the animals by hand at night under spotlight. Animals (>10 m total length) were captured by hand, similar to the method described by Wells (1987) (Handling as described by Chelovsky (1993), was used occasionally for adults). The wild animals were not easy and usually submerged before the netter was in place, similar to what has experienced by Webb and Menzel (1977) with *Craugastor paranaensis* in Australia and Hutton et al. (1987) in





Figure 1. Partial distribution of BSC in coastal South America (shaded area). It includes the whole basin of the Pacific (black area) and the area where mite samples were captured for the present study. Adapted from Noriega-Vargas (1992).

Table 1. Study site description. Approximate population size, maximum number of mite samples used during night-survey, Sample size, number of mite samples captured.

<b>A. Yoffe Island</b>	
Where was recorded	Paracutuba Bay
Latitude	12°49'41" S
Longitude	67°11'11" W
Area (ha)	~ 32
Habitat type	coral (reefward)
Common aquatic plants	Fan: <i>Cyrtosira</i> , <i>Cyrtosira</i> sp. Fan: <i>Halimnaphra</i> , <i>Halimnaphra</i> <i>brachycauda</i> Fan: <i>Proteridionaceae</i> , <i>Proteridion</i> sp. 1 Fan: <i>Schizothauma</i> , <i>Schizothauma</i> <i>reticulifolia</i> and <i>S. curvata</i>
Approximate population size	0
Sample size	1
<b>B. Punta de Arena</b>	
Where was recorded	Paracutuba Bay
Latitude	12°29'40" S
Longitude	67°04'51" W
Area (ha)	~ 366
Habitat type	lagoon
Common aquatic plants	Fan: <i>Stylocostion</i> , <i>Stylocostion</i> <i>marginatus</i> and <i>S. marginatus</i> Fan: <i>Thalassia</i> , <i>Thalassia</i> <i>capitata</i> Fan: <i>Laminidionaceae</i> , <i>Laminidion</i> sp. 1 Fan: <i>Myrtilidionaceae</i> , <i>Myrtilidion</i> sp. 1 and sp. 2 Fan: <i>Hydrocharitaceae</i> , <i>Hydrocharis</i> sp. Fan: <i>Schizothauma</i> , <i>Schizothauma</i> <i>reticulifolia</i> and <i>S. reticulifolia</i>
Approximate population size	20
Sample size	8

Table 1 (cont.): Study semi-desert sites. Apparent population size: maximum number of animals seen during night counts. Sample size: number of animals captured

#### C. Fennell

Watercourse associated	Arava Creek
Latitude	33° 24' 38" N
Longitude	117° 13' 17" W
Area (ha)	~ 10
Habitat type	Agave
Common species plants	Fam. Typhaceae: <i>Elyda virginica</i> L. Fam. Malvaceae: <i>Alcea rosea</i> L. Fam. Boraginaceae: <i>Physalis</i> sp.2 Fam. Nymphaeaceae: <i>Nymphaea</i> sp.1
Apparent population size	22
Sample size	3

#### B. C. Longquads

Watercourse associated	Arava Creek
Latitude	33° 50' 20" N
Longitude	117° 07' 20" W
Area (ha)	~ 2
Habitat type	agave semidesert scrub
Common species plants	Fam. Typhaceae: <i>Elyda virginica</i> L. Fam. Cyperaceae: <i>Eleocharis</i> sp. Fam. Poaceae: sp.1 Fam. Nymphaeaceae: <i>Nymphaea</i> sp.1
Apparent population size	6
Sample size	3

#### T. Davidson

Watercourse associated	Pedernales Creek
Latitude	33° 26' 37" N
Longitude	117° 12' 22" W
Area (ha)	~ 3
Habitat type	agave semidesert scrub
Common species plants	Fam. Malvaceae: <i>Alcea rosea</i> L./Fam. Boraginaceae Fam. Lamiaceae: <i>Origanum</i> sp.2 Fam. Solanaceae: <i>Solanum</i> Fam. Poaceae: sp.1 Fam. Cyperaceae: <i>Cyperus</i> sp. and <i>C. brevifolius</i>
Apparent population size	26
Sample size	12

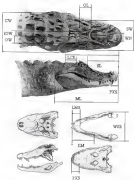


Figure 1. Morphometric variables. Dorsal and lateral view of *Cricetus lemmings* (adult). See Table 1 for description of variables. (adapted from Wernisch and Mennert (1991:311, Fig.126) after Kutschera (1949, *Atlas der der Wildtiere*, Teil 5, 535))





The morphometric variables used in this study were adopted from Jordanley (1973) and Hall and Poston (1994). They are based on three distances between landmarks (body and head size) resulting in ratios between measurements (size variables). Hall and Poston used three ratios relative growth analysis. Relative growth expresses change of proportions in body size increases. The study of relative growth has been characterized by Gould (1966) as the study of size and its implications in ontogeny and phylogeny. However, describing growth processes and size implications, these ratios express two (area, shape) variables in the sense that they represent relative length and width instead of absolute values.

The use of ratios presents several disadvantages (Hall and Smith 1971). According to them, ratios tend to be relatively insensitive, are normally distributed and heterogeneous. However, ratios ratios are still used by some authors (e.g., Hall and Poston (1994) they have been selected and discussed in this present study for comparison purposes.

### 3.4 Exploratory Analysis

Growth curves were created for each through two year-old capture intervals for all measured size variables. Univariate and multivariate three source analysis and body and head size and size variables are compared in regressions. Age of wild animals are estimated based on morphometric methods developed for capture-recapture of known age.

Allometric relations between morphometric variables and body size were determined. These relations may have a useful application in the context of body size based on parts of the body (for examples concerning predation, see Niko and Møller 1973 and Hagenstein 1982). Allometric relations of reproduction were also determined in order to investigate the degree investment in their size (Kruuk 1999).

### 3.5 Methodology

Three correspondence analyses are based on the present study, concerning the following dependent variables: clutch, sex, age, area of origin, and environment (capture or wild) of individuals (Figure 1). The first and simplest one can be stated as: Is there a significant morphometric variation between groups: (a) between individuals from different stretches, between males and females, and between animals from different ages, (area of origin is environment)? In other words, do animals from different groups vary morphometrically? Or yet, are the mean vectors of these statistical populations equal? (Aitchison and Rippen, 1977) It is better to

investigate those questions such as: what is the analysis of variance (ANOVA) a means of testing (shape, size, type, etc. of angles and moments about (response or wild) as responses (dependent variables) and level size and other variables as quality (independent variables)

The basic difference between parametric Analysis of Variance (ANOVA) and MANOVA lies that for the first, an individual observation consists of only one single value or number whereas for the second multivariate is represented by a set of values, each one corresponding to a specific variable. In other words, in MANOVA and other kinds of multivariate statistical analysis, an individual can be represented as a series of one but not as many values as variables in the model. In ANOVA, an individual could serve as a matrix of not just only few values and only values. The steps for MANOVA and ANOVA are similar. However, whereas ANOVA, uses simple statistical algebra, MANOVA and other multivariate statistical analysis use matrix algebra, where operations, although essentially the same, are more complex and time consuming than their statistical counterparts.

If growth affects them differently in the groups above, i.e., the answer for the questions above is positive, a related question can be stated as: Can the group to which a specific individual belongs, be predicted? In order to answer this question, researchers study not only testing the groups above as response (dependent variables) and level size multivariate variables as predictor (independent variables)

The main purpose of discriminant analysis is to find linear combinations of the variables that maximize differences among pre-existing statistical populations (Nelson and Bryant 1979). Cross-validation is used to compare the asymptotic expected error rate which is the proportion of misclassified observations. Asymptotic error rate leads to the optimum because the data being classified are the same data used to build the classification function. This procedure starts by first observation from the data set, develops a classification function using the remaining observations that classifies the initial observation. Next, it removes the first observation to the data set, enters the second observation, and repeats the same process. It continues to do a similar work all observations in the data set (Mumtaz 1996). This process is similar to jackknife procedure (Folger 1971).

The results of the linear discriminant analysis presented in this study include the classification matrix (also called "confusion matrix") and the linear discriminant functions (Mumtaz 1996). The squared discriminant of a specific observation ( $x$ ) to the mean (center) of a group (also given by the general formula below) with the assumption of equal within group covariance

$$d^2(\mathbf{a}) = (\mathbf{a} - \mathbf{m})' \mathbf{B}_1 (\mathbf{a} - \mathbf{m}_0)$$

where

$\mathbf{a}$  is a row vector of length  $p$  containing the values of the predictors for the observation

$\mathbf{m}_0$  is a column vector of length  $p$  containing the means of the predictors calculated from the data in group 1

$\mathbf{B}_1$  is the covariance matrix calculated from the data in group 1

An observation  $\mathbf{a}$  is classified into group 1 if the squared distance of  $\mathbf{a}$  to group 1 is the smallest. Thus, the formula above expands to

$$d^2(\mathbf{a}) = (\mathbf{a} - \mathbf{B}_1 \mathbf{x} - 0.1 \mathbf{m} - \mathbf{B}_1 \mathbf{x})' (\mathbf{a} - \mathbf{B}_1 \mathbf{x})$$

where the term in square brackets is a linear function of  $\mathbf{x}$ , and is called the linear discriminant function for group 1. Here group 1 is the group with the smallest squared distance (or the largest linear discriminant function). Thus, the parent the classification of new observations (Mardia 1964)

Principal components analysis (PCA) were one of the commonest means of analysing multivariate and time series data matrix of log-transformed raw and ratio variables. Correlation matrix is the most usual matrix for the multivariate variables. It measures the similarity of variables and it is appropriate when differences in variables – scaling could reasonably avoid variance. PCA makes a correlation matrix, reducing correlations among variables to zero. Analogously, PCA creates a covariance matrix reducing covariances among variables to zero. In both cases, however, variance is preserved (Mardia 1964)

The results of the principal components analysis presented in this study are the eigenvalues (i.e. the variance of the principal components), the proportion and cumulative proportion of the total variance explained by each principal component, and the coefficients for each principal component. These coefficients are unique weights for a sign change in all coefficients of the eigenvalues are retained and not zero. If no eigenvalue is retained, then the “spare component” by all the principal eigenvalues vectors corresponding to the same eigenvalue is unique but the individual vectors are not. Therefore, the coefficients that our statistical package presents may not agree with the coefficients presented by another (Mardia 1964)

PCA reduces dimensionality and helps interpretation of the often crowding relationships among variables that were not previously suspected (Johnson and Wichern 1992). Some key shape elements can be identified via PCA (Johnson and Wichern 1992). The first principal component corresponds to size variation, although log transformation of data may result in the split of some shape variables on 2 (Shapiro 1964; Jensen 1988). In other words, if there is only

that variation among individuals  $\rightarrow$   $n$ , they are similar (see Thompson 1977) – there will be only one principal component of untransformed data accounting for the total variation. On the other hand, when there is also a residual shape variation, select principal components result from the analysis (Jensen 1973).

At last, if it is possible efficiently predict the group to which an individual belongs using all variables, it is worth determining which subset of variables are the most important sources of variation for the model. The answer for this question is more vague than for the question above, but not less important since it might have method or measurement, in addition, it is usually desirable to work with as few variables as possible without significantly lowering efficiency. Some steps should be taken in order to answer the question. The first one is to try to reduce the number of variables used in the present study was a least-squares regression (Jensen 1974), where the “best subset” was chosen based on the highest coefficient of regression ( $R^2$ ) for an arbitrary number of variables. Variable selection in this method is based on the fact that the subset would very virtually maximize the regression coefficient and predict better regression with smaller variance than the full model using all predictors (Jensen 1974).



Figure 2. Questions and statistical procedures of the present study.

Careful is necessary when using a statistical procedure of variable selection because the selection does a purely statistical point of view may not be the best from a subsequent perspective (Cody and Smith 1987). However, if  $n$  is as big as in this case – the study purpose is to statistically describe the statistical population, it is practically there not matter whether or not the

variables make biological sense (Jolliffe 1999). After reducing the number of variables, MANOVA, discriminant analysis with cross-validation, and PCA were repeated with the best subset in order to test for its efficiency.

Univariate analyses: Log transformed concentrations and ratios are separately analyzed for comparative purposes with their untransformed counterparts. Log-transformation is a simple device that may save and requires interpretation, and statistical designs must of the effect of body size on other attributes (Peters 1983). An example of this situation is presented and discussed. All statistical analyses were done in Minitab for Windows (Minitab 1994) and these procedures are chosen when adequate.

## CHAPTER 4 RESULTS AND DISCUSSION

### 4.1 Variables, Distributions, Correlation, and Reliability

Descriptive statistics of all variables are presented in Appendix A. The majority of the variables are either normally distributed or positively skewed but not very skewed. Some variables need to be not normally distributed but may lead to spurious correlations (Cortality 1978; Anthony et al. 1976) but these distribution effects of the use of ratios did not significantly occur in the present study. Probability distributions of variables were interestingly bipartite-like patterns of fit seen (Fyfe and Jones 1976), which is similar to Shapiro-Wilk Test (Shapiro and Wilk 1965). The results are presented in Appendix A.

Although the capture and handling sample distributions is skewed in young animals, the relatively large sample size (20-400 individuals) including 100 adults (3 year old or older) counterbalances skewness as it is previously proved the skewness distribution effects for age estimation results were grouped by age, from one to five years old. In this case, distribution was normal or approximately normal for virtually all morphometric variables in all developed age classes.

On the growth curves and on the allometric studies of capture animals (pages 4-11, 1 and 6-11) respectively, underrepresented data on age classes could influence regression equations or models. There is a common problem with any biological applications of regression equations (for discussion see Peters 1983 [14-15]). A visual examination in Figure 1-8 suggests that this problem may be happening on the growth curves of capture animals. The absence of animals six and eight years old and the presence of a "big" one year old and a "small" one year old single individuals may erroneously suggest that the whole population would tend to decrease in body size from seven to two years of age. Mathematically speaking, this could lead us to represent the growth curves of these animals with a cubic instead of a linear or quadratic regression as presented in Table 4 and discussed later on item 4.2.3.

Capture and handling are health-related adverse wild animals are study biased. There may have had some influence on the results of the studies of differences between wild and

regular animals, as presented and discussed later in sect. 4.2. Similar is the trend through age classes of captive animals. This may have had some influence on the results of the overall demographic study of captive animals, as presented and discussed later in sect. 4.2.4.2.

A correlation analysis presented in Appendix B. Body- and head-size variables are highly correlated in wild and captive non-breeding animals. Correlation coefficients varied from 0.742 to 0.949 in wild animals and from 0.894 to 0.957 in captive non-breeding. Morphology presented a considerably lower correlation between body- and head-size variables, varying in absolute values from 0.606 to 0.709. Correlation coefficients among rate-variables and between size and head-size variables were in general considerably lower, varying in absolute values from 0.407 to 0.740 in wild animals, and from 0.611 to 0.785 in captive non-breeding. Similarly to the pattern discussed above, morphology presented a considerably lower correlation among rate variables and between distinct head size variables, varying in absolute values from 0.460 to 0.692.

The present study is exploratory, in the sense that one of its goals is the determination of which set of morphometric variables are more affected by climate, genetic growth, age, size of origin, and environment. A practical significance of these high correlations levels among variables for future morphometric studies is that most of the data to be added from each year. Consequently, it is probably unnecessary to take them all every time, which would be time-consuming. However, in some circumstances, if independent variables are correlated, then the partial regression coefficients associated with them may not be accurate to reflect the dependence of response variable that exists in the population. This correlation among independent variables is known as multicollinearity, non-independence in the relationship among variables. In practice this is of little consequence if it is slight, but, if multicollinearity is substantial then consequences can be serious (Joh 1994-1997). In the present study these high correlations coefficients were detected for most of the dependent variables only (head- and body-size variables). The response multicollinearity was limited for independent variables (climate, size, age, size of origin, and environment).

Clustering of variables based on similarity levels may be helpful to classify variables into groups, where the groups are initially not known. Agglomerative clustering begins with all variables separate, each forming its own cluster. In the first step, the two variables closest together are joined. In the next step, either a third variable joins the first pair, or two other variables join together from a different cluster. Each step usually is done less often than the step before until, at the end, all variables are combined as one cluster. Given two variables are



combined in a cluster. Day may join with other variables, but day will always remain together (Figure 4)

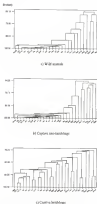


Figure 4: Approximate hierarchical clustering of variables (Distance procedure, Day = 1, Hierarchical = 1, Cluster Variable)

Body- and head-size variables are correlated apart for body width and capture rate handling animals whereas the pattern is not followed for headlength. Headlength predicts most of their size except for standard together with three related head-size variables. Thus BL and BL:ST are correlated together (the same happening with SHL and SH:SL, LH and PLL:SL, LH:SL and BL:SL, SHL and SH:SL and PLL and PL:SL). This difference between headlength and other variables is possibly caused by headlength's smaller variance in body-size.

In total, to test the relationships between morphometric variables need to be independent of body size. The pattern is related to the lower correlations between variables chosen by headlength most variables are used in manner distance between clusters. Correlation and redundancy between variables and the influence of body size on the relationship between morphometric variables is discussed later with the alternative relations presented in table 4.2.1.

#### 4.2 Relative Growth

##### 4.2.1 Growth Curves of Captive Animals

There are three general approaches in the analysis of growth data. One approach involves measurements that are independent dimensions or the size, is related to each other obtained approaches are required if the same individuals are measured repeatedly over time. The third approach is to fit growth data to models that assume particular functional relationships between size and time (Anderson 1982). The first approach above is used in the present study. Growth curves are presented here as regression equations with age being the independent variable ( $X$ -axis) and generally a single measurement related to body size (e.g., body mass or snout total length) being the dependent variable ( $Y$ -axis).

Growth models have been developed in order to provide mathematical treatment to the process of biological growth. "Von Bertalanffy's asymptotic growth model" (Bertalanffy 1938) and "Bathurst's family of growth models" (Bathurst 1981) are the most well-known. Richard's family includes von Bertalanffy and the following other widely used models as special cases, logistic, sigmoidal, logistic or asymptotic, and Gompertz. The basic elements that differentiate these models are the growth rate and the point of inflection of the curve (Bryant and Richard 1983).

In a range of studies on the present study in its growth curves of capture animals in age growth model because the models utilized were based on continuous growth and their growth curves are of little utility in studying wild populations. Polysplacental organisms are presented in explanatory data for deeper understanding the relationships in growth curves of each

equation beyond its limitation is proposed, which could be considered as “functional size” by quoting Akçaya et al. (1994). However, the relatively large sample size and the availability of records from a limited age spectrum (from under one year-olds) may be useful for preliminary analysis of growth pattern for the species.

Although mostly studied separately, cell, individual, and population growth are intimately-linked parts of the same ecological process. An individual can be seen as a population of cells that individually grow, making up a very complex system resulting in the growth of the individual as a whole. In turn, the upper levels of the individual (growth rates) – birth, mortality, and death – are intricately linked to the upper levels of population dynamics. Whole-body growths are used respectively for increase and decrease in the population size, individual mortality is simply related to dispersal and reproduction, which is also necessary for the population self-repopulation or decrease, and even has consequences for geographic distribution of the species.

It is inherently noted the tremendous difficulty in mathematically modeling organ growth. The emphasis that growth is mainly dependent on external factors such as nutrition, temperature, tide space, as well as internal factors such as hormones, progressive differentiation, change of water content, and age (literaturely, 1933). The emphasis, however, examples from physics and physiology in which complex phenomena are sometimes be described by statistical laws and suggested that such procedures could apply to organ growth. In the present study, age analysis was performed and discussed in context of variation for the growth of individuals. Since in this study, the influence of environmental factors such as clutch, maternal heritage (also called kinship heredity), sex of origin and environment, on the clammer growth is discussed.

Table 1 presents the analysis of variance (ANOVA) of sex and age in relation to the morphometric variables taken. ANCOVA here is used to compare male and female growth trends. While significant ( $P$ -value  $< 0.100$ ) separate growth curves are presented. When sex ( $P$ -value  $> 0.100$ ) analysis and therefore are pooled together. Six of eight “length” variables (TL, EL, BL, CL, LCL, and PCL) and three of six “width” variables (CW, SW, and WBL) showed differences between males and females. This pattern can be an indication that old males tend to be longer than females both in terms of body and head length which is compatible to what has described by Hall (1971). No clear variable presented differences between males and females.

Five of eight variables related to the upper jaw (BL, EL, CL, LCL, and PCL) and one of four variables related to the lower jaw (WBL) showed differences between

males and females. Crustaceans most often use generally one/less visual displays whereas the lesser jaw is kept between the water surface while the upper jaw and the top of the head are exposed above the detached males see Garouk *et al.*, 1990; Agassiz *et al.* 1993; Lamy 1993, and Vlier 1993).

Table 2. Analysis of covariance: age as covariate of age on growth curves ( $P$  values)

Body-size variable	Age	Sex	Independent variables	Age	Sex	Body-size variable	Age	Sex
TL	0.000	0.070	OC	0.000	0.003	SL	0.000	0.000
SW	0.000	0.000	CW	0.000	0.000	SLST	0.000	0.000
HW	0.000	0.000	BL	0.000	0.000	SC	0.000	0.000
HW	0.000	0.000	SW	0.000	0.000	SCST	0.000	0.000
			OL	0.000	0.000	SLST	0.000	0.000
			OW	0.000	0.000	SL	0.000	0.000
			OCW	0.000	0.000	SLST	0.000	0.000
			LCW	0.000	0.000	SLST	0.000	0.000
			OCW	0.000	0.000	SLST	0.000	0.000
			LCW	0.000	0.000	SLST	0.000	0.000
			OCW	0.000	0.000	SLST	0.000	0.000
			LCW	0.000	0.000	SLST	0.000	0.000
			OCW	0.000	0.000	SLST	0.000	0.000
			LCW	0.000	0.000	SLST	0.000	0.000
			OCW	0.000	0.000	SLST	0.000	0.000
			LCW	0.000	0.000	SLST	0.000	0.000
			OCW	0.000	0.000	SLST	0.000	0.000
			LCW	0.000	0.000	SLST	0.000	0.000

Model procedure: (a) = ANOVA in General Linear Model

Dependent (response) variable: body-size variable

Model (predictor) variable: Age

Covariate: Sex

The performance of upper jaw and carapace variables on the visual description can indicate a second reference associated with the social behavior of crustaceans. Thus, the more exposed regions of a male body would function in the more prominently than on the female. There is a great number of unique examples not being both and, conversely, besides the "reference" periods full and low state (for exposed structures and examples see Gould and Gould 1988; Andersen 1994; and Blett and Blett 1994). Sexual dimorphism and the occurrence of individual gender-based behavioral asymmetries is discussed here in more detail.

In the present study growth curves of 1–10-year-old captive animals are presented for all size and shape variables (Tables 4–5, Figures 2–5). As explained above, males and females growth curves are presented together unless for the variables with a significant effect ( $P$  value < 0.05) of gender on growth observed by the analysis of covariance (Table 1). With the exception

of body mass,  $\ln(\cdot)$  transformation did not improve regression equations. Thus, all other variables are presented without transformation.

Table 4. Growth curves of captive animals. Body and head mass (g) vs. age.

#	Sex	Y	a	b	c	d	R-squared	P	n
1	male	Y <sub>1</sub>	12.0787	43.4754	-11.1236	0.1936	0.988	0.001	25
2	female	Y <sub>1</sub> <sup>a</sup>	14.4613	23.1388	-10.7407	0.088	0.989	0.079	88
3	male	Y <sub>2</sub>	-7.4529	36.8483	-7.8888	0.1386	0.988	0.007	38
4	female	Y <sub>2</sub> <sup>a</sup>	4.0867	37.4878	-6.4948	0.088	0.989	0.076	92
5	male	LogBM	-1.0964	1.0948	-0.0388	0.6108	0.988	0.003	18
6	female	LogBM	-1.0388	0.4779	0.0008	0.088	0.989	0.088	10
7	male & female	BM	-12.0494	1.0486	0.088	-0.0004	0.988	0.004	114
8	male	ICL	0.4743	0.0478	-0.001	0.0783	0.988	0.088	34
9	female	ICL <sup>a</sup>	0.1478	0.0488	-0.0017	0.088	0.989	0.071	82
10	male	ICW	-0.0483	0.0383	-0.0004	0.088	0.989	0.088	34
11	female	ICW <sup>a</sup>	0.0478	0.0388	-0.0008	0.088	0.988	0.088	82
12	male & female	ICL	0.0388	0.0478	-0.0008	0.088	0.988	0.088	118
13	male	ICW	-0.0478	0.0388	-0.0008	0.088	0.988	0.088	34
14	female	ICW <sup>a</sup>	0.0478	0.0388	-0.0008	0.088	0.988	0.088	82
15	male & female	ICL	0.0478	0.0388	-0.0008	0.088	0.988	0.088	118
16	male & female	ICW <sup>a</sup>	0.0478	0.0388	-0.0008	0.088	0.988	0.088	118
17	male & female	ICW	0.0388	0.0478	-0.0008	0.088	0.988	0.088	118
18	male	LCI <sup>a</sup>	1.071	0.071	-0.0008	0.088	0.988	0.088	28
19	female	LCI <sup>a</sup>	0.071	0.071	-0.0008	0.088	0.988	0.088	82
20	male & female	LCI <sup>a</sup>	0.071	0.071	-0.0008	0.088	0.988	0.088	118
21	male	ICL	0.0478	0.0388	-0.0008	0.088	0.988	0.088	34
22	female	ICL <sup>a</sup>	0.0478	0.0388	-0.0008	0.088	0.988	0.088	82
23	male & female	ICL	0.0478	0.0388	-0.0008	0.088	0.988	0.088	118
24	male	ICW	0.0388	0.0478	-0.0008	0.088	0.988	0.088	34
25	female	ICW <sup>a</sup>	0.0388	0.0478	-0.0008	0.088	0.988	0.088	82
26	male & female	ICW <sup>a</sup>	0.0388	0.0478	-0.0008	0.088	0.988	0.088	118

Y<sub>1</sub> =  $0.0001 \times \text{age}^2 + 0.007 \times \text{age} + 0.0001$

Multiple predictors: Best fit Regression → Final Least Squares (2-predictor Regression)

While the equations of BM, ICL and ICW were not considered because the values of asymptote were smaller than transformation did not improve results.

(a)  $\ln(\cdot)$  instead BM was included in the equation (d = 0) whenever significant (P-value < 0.05).

Quadratic equation (a) was included in the equation, or (b) otherwise either quadratic or cubic equation were significant (P-value < 0.05).

Male and female presented separately when ANCOVA for sex was significant (P-value < 0.05). See Table 3 for P-values.

<sup>a</sup> Although cubic equation considered asymptotic equation, the quadratic equation or transformation (logarithm) did not present model structure decreasing values of Y<sup>2</sup> or positive values of age<sup>3</sup> which indicates cubic curve or biology.

<sup>b</sup> Growth curves for *Uta stansburiana* were developed only for the ages of one through five years.

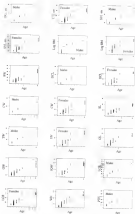
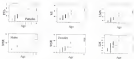


Figure 3. Plots of growth of captive hermits. Body- and larval size variables. Age in years. Males and females plotted together unless stated otherwise. See Table 2 for growth differences between sexes (and Table 4 for regression equations).



**Figure 5.** (cont.) First growth of captured animals (Body- and Total-size variables). Age in years. Males and females presented together within same otherwise. See Table 3 for growth differences between ages. See also Table 4 for regression equations.

**Table 3.** Growth curves of captured animals (size variables).

	Sex	T	1	2	3	4	P-value	r	N
1	males & females	52.37	0.1365	0.0434			0.000	0.768	134
2	males & females	55.17	0.1568	-0.0549	-0.0042		0.000	0.895	134
3	males & females	57.91	0.2561	-0.0468	0.0643		0.000	0.927	134
4	males & females	60%	-0.1363	-0.0333	0.0131		0.000	0.788	134
5	males & females	62.76	0.0365	-0.1357	0.0767	-0.0047	0.000	0.875	134
6	males & females	65.61	0.0404	0.0094	-0.0031	-0.0001	0.000	0.902	134
7	males & females	68.39	0.0559	-0.0709	-0.0015		0.000	0.774	134
8	males & females	71.23	0.2143	-0.0199	0.0044	-0.0001	0.000	0.948	134
9	males & females	74.04	0.0889	0.0001			0.007	0.864	134
10	males & females	80.54	0.2644	-0.0173	0.0022		0.000	0.944	134
11	males & females	83.58	0.2033	-0.0753	0.0133		0.000	0.920	134

$$Y = a + bx + cX^2 + dX^3 \quad X = \text{age} \quad \text{growth}$$

(Blank position, last  $\rightarrow$  Regression  $\rightarrow$  Fitted Line Plot (Polynomial Regression))

Values shown (N) are calculated by the regression  $\rightarrow$  (N) = number of significant (P-values  $\leq 0.05$ )

Quadratic values (C) are included in the equation ( $\rightarrow$   $\rightarrow$  N) where quadratic position  $\rightarrow$  values shown were significant (P-value  $\leq 0.05$ )

Males and females presented together between ANCOVAs for sex was not significant (P-value  $\geq 0.05$ ). See Table 3 for P-values.

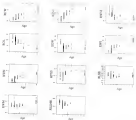


Figure 8. Effect of growth of capture volume. Growth variables, Age in years, Weight and Length, presented together (see Table 3 for significance of var. and age (see Table 3 for significant response).

Both real and virtual variables' growth curves presented high coefficients of determination ( $r^2 \geq 0.8$ ), the proportion of the variance of one variable determined by the variance of the other. Significance of binary growth equations presented ( $r^2 \geq 0.99$ ) whereas the others presented  $0.800 < r^2 < 0.99$  with only one exception (34) ( $r^2 = 0.403$ ). This means that the models proposed are efficient to explain most of the variance presented. Models presented higher coefficient of determination than females possibly because of their smaller sample size (Table 3, Figure 7).

Early variables presented lower and more variable coefficients of determination, varying from virtually zero to 0.55 to 0.95 to 0.97 (Table 3). This means that age is not a constant value of variance like the former is it is for the later, which can be easily visualized in Figure 8.



The data means that *SVL* variables are generally less age-dependent than body- and head size variables.

Many (1994) suggest that growth data for reptiles and amphibians and reptiles are best fitted by logistic-4p, von Bertalanffy, or Gompertz, whereas data for larger reptiles without limbing the von Bertalanffy equation. Doherty (1992) suggest that both the Richards and the von Bertalanffy equation model described the growth of captive male and female alligators equally well, but the Richards model explained significantly more variation for wild alligators. Holland-Parker (1994) suggests that both methods are less efficient than Brody's method (Brody 1945) for growth data for vertebrate invertebrates. Adamowicz (1992, 1998) suggests that either the von Bertalanffy or the full Richards model may often be insufficiently for summarizing the growth of amphibians. The former being more appropriate to compare results with other published data, and the latter possibly better to specify individual cases.

The SVL growth curve of the present study is surprisingly similar to the one presented by Magagnoli and Sazdovitch (1992) for *Crotalus cerastes* collected in Central Amazonia, Brazil. On the other hand SVL growth curves for *BNC* in several environments of Rio de Janeiro, São Paulo State, Brazil, presented a much slower growth rate possibly because of lower light intensity rate (Silva et al., in press).

Using both Richards and von Bertalanffy growth models, Mlynarski et al. (in press) suggest asymptotic SVL from 10 to 11 cm with the age of female first reproduction ranging from 11 years (all cm SVL) to the growth of 12 to 13 years (11 cm SVL). The present study suggest a period of 1 to 10 years for the female first sexual maturity (10 cm SVL), which is compatible to the maturation of females at sexual maturity in Santa Fe, Argentina (Lamas, pers. comm.) and to the sexual age at first reproduction (from 14 years) observed in captivity in Rio Paulo, Brazil (Doherty, in press).

Females sampled in this study showed a clear tendency of reaching asymptotic values of most body and head size variables and some rate variables from 7 to 10 years of age. As explained above in case of 1 sample cases may explain the best fit of cubic rather than quadratic equations for females in some variables (see Table 1). In these cases, quadratic equations were preferred because the use of cubic equations would lead to a decrease in the residual sum after a certain age, which would not make biological sense in this case.

### 4.2.3 Age Determination

Many techniques are used to determine age in teleosts. Some are simple and easily employed without special training, but for some species sophisticated techniques are required. According to Lorenz and Tabor (1983), the criteria for ideal age determination should include: following characteristics: "independence from irregular seasonal and physiological variations", "value independent of age, time or past climate without subjective judgement", "suitability for being subject of all ages" and "ease of application by non-related technicians". Although some standard techniques exist determinants for some species, Lorenz and Tabor consider that they do not. For that reason, they propose that not so much of the ideal criteria may be compromised depending on the degree of accuracy required, in order to obtain the advantages needed.

While pharyngeal osteoichthines are generally easier to determine age in fish (Birstein and Pelson (1996), the standard techniques for post-smolt (or age determination in salmon involve lower' natural rings (Birstein, 1984), growth rings (Frost (1973), otoliths which already is long known' post-smolt (van Birstein and Birstein, 1987), tooth eruption replacement and wear (Lorenz and Tabor 1983), and growth layers group counts in bone (Morris and *et al.* 1994). Lorenz and Tabor (1983) and Birstein and Pelson (1996) present extensive reviews on aging techniques for fish and mammals in the traditional multiple measurement approach.

Age determination of fish is essential in fisheries management decisions (Purdy 1987). Although recent studies have demonstrated the great age of some species can result in serious underestimates (Birstein and McFarlane 1987) the following standard techniques are generally used depending on the species life span: daily growth rings (Purdie 1971, 1974), annual growth bands (Cavallazzi 1987), length-frequency and weight progression analysis (Purdy 1987), and age-growth related models (Purdy 1987). Bar (1994) presents a comprehensive review on the age determination of fishes.

Purdie and Vernal (1987) subject five methods of age determination in capelin and report: "captives of known individuals, comparisons from size-frequency data, otolith microstructure, and tooth analysis (for some information)". They question the accuracy of otolith microstructure and report strongly correlated otolith microstructure and age, and conclude that only two of the methods above are reliable: otolith microstructure and mark-recapture.

Otolith microstructure is based on the observation that growth in capelin and capelin is periodic and that post-smoltification is essential in various loach (Birstein 1988), and in the euphrates river of teleosts (Cage 1988). The technique usually consists of counting annual

logarithmic growth (Janssen 1981). Mark-recapture techniques are based on the growth record of subsets of known age and the application of growth models that best fit their data (Anderson 1982).

Annual growth-rings in phalanges were used to determine age of three European hogs (Collins and McCarthy 1982). Databank technology was more efficient than loggers by using growth model-based open mark-recapture analysis rather than constant left open and the relationship between age and age in log-linear relationships (Anderson and Anderson 1988) in a database. Alpha, Canada (Raman et al. 1984). Jensen (1984) presents a technique to correct for the age determination and longevity of ungulates.

Mark-recapture, interval records, and longevity are also proposed by Collins (1984) as the primary methods of aging ungulates, however, databank technology was proposed as the most precise method of age determination in ungulates (Collins 1984). Examples concerning ungulates (Collins 1984), are hinds (Jag et al., 1984; Klinger and Mowat 1985; Reynolds and Jag 1985) hinds (Cheney et al. 1985), and muskoxen (de Gooijer 1985; Reynolds 1985; Harris 1985, 1987) constitute the system.

Although databank technology is an excellent post-mortem technique its application is being restricted as well as its frequency (Jensen 1987). The applicability of its use as being limited has prevented its use for open mark-recapture systems such as the RSC. Therefore mark-recapture experiments and the subsequent application of growth models such as von Bertalanffy and Richards are still widespread (Jensen 1988). However, growth-curve based on single body size measurements are usually considerably superior to age predictors for several years because of their great variation in growth patterns within single populations (Jensen 1987).

Deriving from single growth curves (so-called "models") are generally not considered as representative of little biological meaning. On the other hand, the description of relative growth as a multivariate system, with several sets of variables, would be more efficient to capture deviations from the growth curves when they are found to be associated. This would prevent loss of biological information (Jensen and Madsen 1988). Wiley et al. (1984) and Wiley and Davis (1987) propose that "because MANOVA, unlike other than logistic, considers among characters, it is the correct statistical test for evaluating overall group differences" in systematic studies.

In the present study, an alternative multivariate approach is proposed for the study of age determination in RSC. Age classes (one-, two-, three-, four- and five-year old animals) are considered distinct groups and the hypothesis presented in Figure 1 are tested. Univariate and

multivariate two-sample analysis of body- and head-size variables was compared to age predictors. Finally, more explanatory variance for head shape (multiple regression) developed for age through the two-old experimental groups retained for the construction of wild animals' age.

#### 4.1.3.3 Age discrimination of equine animals (1–5 years old)

There is a significant difference ( $F$ -value = 9.902) between age classes (one-, two-, three-, four- and five-year old animals) in terms of the whole set of head size variables (log-transformed head size variables, and size variables for Withers Length, Height, and Volume) from MANOVA. This means that the variance among groups is bigger than the variance into groups, which increases our discrimination (see Figure 1).

The second question tested was: Can the age class in which a specific individual belongs be predicted? In order to answer this question, linear discriminant analysis with stepwise selection was for untransformed head size variables (Table 2), log-transformed head size variables (Table 3), and size variables (Table 4). The overall proportion of correct classification was similar for the three kinds of variables: varying from 8.7% (34.3%) for untransformed variables to 9.7% (38.4%) for untransformed head size variables.

There was, however, a significant variation among age classes in terms of the proportion of correct classification for the three types of variables: it varied from 8.3% to 1.00% for one-year old, from 0.14% to 1.44% for two-year-old, from 0.00% to 0.12% for three-year-old, from 0.00% to 0.20% for four-year-old, and from 0.00% to 0.67% for five-year-old animals.

Untransformed head size variables presented consistently better results than the untransformed log-transformed variables. The one-year old class was consistently better discriminated than the others, whereas the four-year old class presented consistently the worst results. Most of the four-year-old animals were misclassified as three-year old. This probably can be explained by the equal distance between groups. Age classes three and four were the closest of all. This suggests that during the period few comparative variations occur. On the other hand, the one-year-old class was the most "distinct" one. This means that there is a consistent ontogenetic variation from the first to the second year of age, which is easy to use even with a little classification weak confidence. This pattern was also reported by Block (2003).

PCA for the whole set of untransformed head size variables (Table 2), log-transformed variables (Table 3), and size variables (Table 4) and their plots (Figure 7) are useful to visualize how much the different age classes are clustered apart. For both untransformed and log

important first-order variables the majority of the variance is included in the first principal component (31.6% and 33.3% respectively) (see equations in Tables 4 and 5). This means that the majority of variance between age classes is covered by overall first size change.

Table 3: Linear discriminant analysis: Response signal against absolute predictors of first size variables (sample size: 18 individuals (Groups: 1 = 3 years-old).

a) Summary of classification with new variables

Group	1	2	3	4	5
1	15	4	0	0	0
2	0	32	2	0	0
3	1	1	7	4	1
4	0	1	4	1	0
5	0	0	0	0	3
Total N	16	37	13	5	3
Percentage	13	30	1	1	3
Discrim.	0.877	0.892	0.598	0.380	0.892

N = 68    Misclassified = 27    Percentage Correct = 41/68

b) Linear discriminant function for group

	Constant	PC1	PC2	PC3	PC4	PC5	PC6	PC7	PC8	PC9
1	1.2716	1.22	0.24	-0.55	-0.18	4.38	0.058	1.98	0.32	2.12
2	1.2250	3.00	-0.97	-0.50	-0.56	0.77	0.044	-0.02	0.88	2.34
3	1.8776	1.78	-0.68	-0.58	-0.15	7.58	0.73	0.58	0.88	1.88
4	0.6110	0.88	-0.41	-0.94	-0.58	8.38	18.25	0.17	0.30	0.20
5	0.0114	7.77	-12.37	7.34	2.27	0.17	14.97	11.28	0.09	2.61

	SE	1-SE	t-stat	p-val
1	0.86	1.10	-0.82	0.42
2	0.72	-0.62	-0.89	0.38
3	0.11	-0.11	-0.52	0.60
4	0.05	0.05	-0.73	0.47
5	-0.20	-0.04	1.28	0.21

The coefficients of discriminant variables for the first principal component are quite similar, ranging in absolute terms from 0.248 to 4.370 (see PC1 in Table 3), whereas log-transformed variables principal coefficients, varying from 0.102 to 0.300 (see PC1 in Table 5). This means that log-transformations showed effects of some specific morphometric variables during independent variables during aging process. Some length (SL) and bone-related metrics (OPB) were the variables that contributed the most to variance, according to the PC1, of log-transformed variables. This very result does suggest that during the first three years of age the morphogenetic

and the distance between eye gaze trigger have influenced relative width the whole head.  
Allometric relations will be discussed later in section 4.2.3

**Table 4: Linear discriminant analysis: Response: age of experimentalists. Predictors: all log-transformed head-face variables. Sample size: 59 subjects (groups: 1-3 years old)**

**a) Summary of discriminant analysis results**

Group	1	2	3	4	5
1	50	1	0	0	0
2	0	14	0	0	0
3	0	0	0	1	1
4	0	2	0	0	0
5	0	0	1	0	1
Total (N)	50	17	11	1	2
Not Classified	0	0	0	0	1
Discriminant	0.899	0.978	0.944	0.999	0.987

(N = 59)    Wilks' Lambda = .79    Population Correlation = 0.92

**b) Linear discriminant function by group**

	Constant	PC1	PC2	PC3	PC4	PC5	PC6	PC7	PC8	PC9
1	1083.9	2733.9	180.1	1792.1	-489.1	354.9	542.9	160.1	244.9	119.1
2	-2023.2	5419.1	494.9	3894.9	-4949.1	284.2	447.1	15.2	289.1	-189.1
3	2911.9	3841.1	288.9	5475.1	-841.2	169.4	307.1	19.1	322.1	-479.2
4	-1941.9	1012.9	184.1	3609.1	487.9	381.1	212.1	89.1	167.1	882.9
5	1199.9	1894.9	175.9	1199.9	-479.1	173.1	101.9	54.9	171.1	129.1

\*\*\*

	PC1	PC2	PC3	PC4	PC5
1	18.1	912.9	-487.1	-429.1	19.1
2	18.1	148.9	-491.1	191.9	184.9
3	-1.2	849.9	-188.9	881.2	173.2
4	4.9	1987.1	107.4	894.1	-184.2
5	4.2	1112.1	429.1	894.1	-11.9

The second principal components for both unrotated and log-transformed head-face variables contributed with 1.44% (1.4%) and 1.61% (2.4%) of the total variance, respectively (see PC1 in tables 4 and 5). The simplification of variables for the second principal component of unrotated variables was the orbital width (PC1), which correlates with the other two eye measurements (PC2 and PC3) in magnitude and sign. This pattern captures a narrowing of the eye orbits on the head during the first four years of age. They become relatively longer (PC2 with positive age) and narrower (PC3 with negative age), while the distance between the eyes increases (PC4 with positive age).

Table 7: Linear discriminant analysis. Response: age of capture animals. Predictors: all size-measured. Sample size: 50 animals. Groups: 1 = 5 years old

*a) Summary of discriminant function coefficients*

Group	1	2	3	4	5
1	66	2	0	-2	0
2	2	50	2	2	0
3	0	2	0	4	0
4	0	2	0	1	0
5	0	2	0	0	1
Total N	10	44	10	2	0
N Correct	10	22	1	1	1
Proportion	1.000	0.502	0.100	0.500	0.500

$N = 66$      $N$  Correct = 32     $\text{Proportion Correct} = 0.485$

*b) Linear discriminant function for group*

Group	BL	FL	HL	SL	SW	BL/FL	BL/HL	BL/SL	BL/SW
1	1044.0	-606.0	1071.0	1129.2	4123.2	0.91	101.6	-1302.0	111.2
2	2033.0	-489.0	2081.0	1050.0	4307.0	0.95	491.2	-1070.0	187.0
3	1071.0	-440.0	1090.0	1043.0	4096.0	0.91	104.2	-1064.0	117.7
4	1054.0	106.0	1047.0	1145.0	4228.0	0.94	601.0	-1048.0	129.1
5	1738.0	602.0	1051.0	1071.0	4057.0	0.91	449.0	-1043.0	205.0

	BL/FL	BL/SW
1	104.0	101.1
2	103.0	140.0
3	111.0	176.0
4	111.0	210.0
5	112.0	201.0

The inter-ventral width (SW) is the variable that contributes most to the variance contained in the second principal component of the log-transformed PCA. There are clear contrasts between 4 leading morphometric variables (BL, FL, HL, and SL) in both magnitude and sign. This pattern possibly captures the widening of the eye region (but not necessarily the eye itself) of the head as relative to the mandible.

The PCA of size variables showed, as would be expected, a completely different pattern. The first principal component accounted with less than half of the total variance, while the second principal component accounted for almost two thirds of the total variance (see regression in Table 10). The various shape components of size variables made the PCA defined in groups. There seems to be a contrast between the relative width and the relative length of snout (compare signs and magnitudes of BL/FL and SW/SL in PC1 in Table 10). There also seems to be a contrast between the relative ventral width (BL/SW) and the relative snout-ventral width (BL/SL) (compare their signs and magnitudes in PC1 in Table 10). These two contrasts

show all the major sources of variation for the first principal component. There may be a “size” component in the context described above, as the shape variables herein are not completely size-independent (see Figure 1) and the dimensions are otherwise redundant (except 4,2,3), but this is rather complex and subjective.

The second branching (BWH) and BCPW is the major source of variation in the second principal component (see PC2 in Table 10). The pattern may also be interpreted as rearrangement of the spine region during independent change in the first three years of age, as the sense that decreased truncal breadth (BWH) with post- or age2 and the spine width leaving narrower (BCPW with age2 or age3).

On Figure 9a, it is possible to see the distribution of animals against the values for the two principal components of the whole set of nontransformed, log-transformed and ratio variables, respectively. One year old animals are clustered apart, as all of them, which is comparable to the pattern derived by the linear discriminant analysis. Assuming PC1 as the “size” component and PC2 as the more “shape” component, and considering the distributions of individual variables, it is possible to infer that there are consistent morphometric variations both in just tail shape during the first year of age. Two year old animals are increasingly put apart, with both intermediate and long-transformed, but not with ratio variables. On the other hand, animals from three to five years of age present a less evident separation.

**Table 1. Principal components analysis. Eigenvalues of the correlation matrix of all best-size variables. Sample size: 19 animals. Only the first principal component was presented**

Variables	PC1	PC2	PC3	PC4	PC5	PC6
Age (years)	0.040	0.036	0.005	0.000	0.000	0.000
Age (days)	0.040	0.036	0.005	0.000	0.000	0.000
Concavity	0.040	0.036	0.005	0.000	0.000	0.000
Circle	PC1	PC2	PC3	PC4	PC5	PC6
BC1	-0.476	0.138	0.118	0.000	-0.000	-0.000
CV	-0.390	0.138	0.000	-0.000	0.000	-0.000
SL	-0.270	0.070	-0.070	-0.000	-0.000	-0.000
SW	-0.270	0.138	0.000	0.000	0.000	-0.000
SL	-0.360	0.070	0.000	-0.000	-0.000	-0.000
SW	-0.360	0.138	0.000	-0.000	-0.000	-0.000
BWH	-0.000	0.036	-0.000	0.000	-0.000	-0.000
LCS	-0.000	0.036	0.000	0.000	-0.000	-0.000
WTH	-0.000	0.036	0.000	0.000	-0.000	-0.000
PCW	-0.000	0.036	0.000	0.000	-0.000	-0.000
SL	-0.000	0.036	0.000	0.000	-0.000	-0.000
LAC	-0.000	0.036	0.000	0.000	-0.000	-0.000
WTH	-0.000	0.036	0.000	0.000	-0.000	-0.000
SW	-0.000	0.036	0.000	0.000	-0.000	-0.000



Table 9: Principal component analysis: Eigenanalysis of the covariance matrix of all log-transformed head size variables. Sample size: 58 subjects. Only the first principal component are presented

Variable	PC1	PC2	PC3	PC4	PC5	PC6
Procrustes	0.988	0.001	0.001	0.000	0.000	0.000
Condylar W	0.756	0.285	0.180	0.180	0.000	0.000
Variable	PC1	PC2	PC3	PC4	PC5	PC6
BL	-0.147	-0.040	0.181	0.180	-0.000	-0.000
CV	0.256	-0.100	-0.003	-0.004	0.117	0.000
SL	-0.120	0.001	0.000	-0.004	-0.000	0.000
SW	-0.290	-0.160	0.101	-0.004	-0.000	-0.000
BL	-0.000	-0.000	0.100	0.004	-0.000	-0.000
CV	-0.000	-0.170	-0.100	0.100	-0.000	0.000
SW	-0.000	0.000	-0.003	0.000	0.000	-0.000
CV	-0.000	0.000	0.100	0.000	-0.000	0.000
BL	-0.000	-0.000	0.000	0.000	-0.000	-0.000
CV	-0.000	-0.000	0.000	0.000	-0.000	-0.000
BL	-0.000	-0.000	0.000	0.000	-0.000	-0.000
CV	-0.000	-0.000	0.000	0.000	-0.000	-0.000
BL	-0.000	-0.000	0.000	0.000	-0.000	-0.000
CV	-0.000	-0.000	0.000	0.000	-0.000	-0.000
BL	-0.000	-0.000	0.000	0.000	-0.000	-0.000
CV	-0.000	-0.000	0.000	0.000	-0.000	-0.000

Table 10: Principal component analysis: Eigenanalysis of the covariance matrix of all size variables. Sample size: 58 subjects. Only the first principal component are presented

Variable	PC1	PC2	PC3	PC4	PC5	PC6
Procrustes	0.987	0.001	0.001	0.000	0.000	0.000
Condylar W	0.880	0.100	0.000	0.000	0.000	0.000
Variable	PC1	PC2	PC3	PC4	PC5	PC6
BL	0.100	0.000	0.000	-0.000	0.000	-0.000
CV	0.000	-0.000	-0.000	0.000	0.000	-0.000
BL	0.000	0.000	0.000	-0.000	0.000	0.000
CV	0.000	0.000	0.000	0.000	-0.000	0.000
BL	0.000	0.000	0.000	0.000	0.000	-0.000
CV	0.000	0.000	0.000	0.000	0.000	0.000
BL	0.000	0.000	0.000	0.000	0.000	0.000
CV	0.000	0.000	0.000	0.000	0.000	0.000
BL	0.000	0.000	0.000	0.000	0.000	0.000
CV	0.000	0.000	0.000	0.000	0.000	0.000
BL	0.000	0.000	0.000	0.000	0.000	0.000
CV	0.000	0.000	0.000	0.000	0.000	0.000
BL	0.000	0.000	0.000	0.000	0.000	0.000
CV	0.000	0.000	0.000	0.000	0.000	0.000

Table 11 presents the best values of variables for age discrimination. Untransformed head size variables present the best results in terms of the significant regression coefficients of discriminators. Log-transformations of the variables in duplicate of two different variables (CV and BL) instead of BL and SL. The best values of size variables showed a preponderance of relative width ratios (CV, BL/BL and CV/BL) whereas length and width variables were shared equally by the best values of untransformed and log-transformed variables.

Table 11. Age regression equations based on best subsets of morphometric variables.

Type of variable	Regression equation	$r^2$	F-value (probab)
Total size	$\text{Age} = 0.175 + 0.002 \text{ LW} + 0.00018 \text{ SL} + 0.00019 \text{ HW} + 0.00017 \text{ PSL}$	0.979	0.000
Log transformed	$\text{Age} = 0.001 + 0.0000001 \text{ LW} + 0.0000001 \text{ SL} + 0.0000001 \text{ HW} + 0.0000001 \text{ PSL}$	0.999	0.000
Ratio	$\text{Age} = 0.001 + 0.0000001 \text{ LW} + 0.0000001 \text{ SL} + 0.0000001 \text{ HW} + 0.0000001 \text{ PSL}$	0.999	0.000

Type of variable	Best subset	F-value
Total size	LW	0.999
	SL	0.995
	HW	0.991
	PSL	0.996
Log transformed total size variables	LW	0.999
	SL	0.997
	LCH	0.974
	PHS	0.991
Ratio	SLHW	0.991
	SLST	0.999
	HWST	0.997
	PHS	0.997

Statistical properties: Best = Regression in Best Subsets

Sample size: 153 caprine animals

All variables chosen are located in the cranium and none in the mandible. Craniometric features morphometric made of endocranial, living in both aquatic and terrestrial habitats and exhibiting many behaviors in the water surface or underwater (Long 1987). These include stereotyped sexual dimorphism observed and described by humans and by the position of the mandible beneath the water surface (whereas the top of the head protrudes still above water) a large skull (see Garrod et al., 1978; Apstein et al. 1981 and Viter 1989). Although some morphological correlates may be added, this pattern stresses the possible relationship between skull behavior and ontogenetic changes. The regression equations presented should of course not be extrapolated beyond the limits of the ages studied.

There was significant difference between age classes (one-, two-, three-, four- and five-year old animals) in terms of the best subset of total size variables (log-transformed total size variables and ratio variables) ( $F$ -value < 0.0001 for Wilks, Levene-Hartigan, multivariate Tests of MANOVA). This means that, even when considering all variables, the variance among groups is still bigger than the variance into groups, which crosses the first step of the interpretation of the present study (see Figure 5).

Linear discriminant analysis with linear variables are presented for the best subset of total size variables (Table 12), log-transformed variables (Table 13), and ratio variables (Table 14). The overall properties of correct classification is significantly better than for the

while all of variables varying from 0.04 for rate variables to 0.003 for log-transformed variables, with an intercepting value for nontransformed variables.

Table 12. Linear discriminant analysis. Response: age of capture controls. Predictors: first values of head-size variables. Sample size: 100 animals. Groups: 1–5 years old.

*all summary of classification with given relations*

Group	1	2	3	4	5
1	10	12	0	0	0
2	1	29	0	1	0
3	0	4	7	2	11
4	0	7	1	1	0
5	0	0	0	0	0
Total N	10	42	8	3	11
N Correct	10	36	7	1	0
Proportion	0.007	0.076	0.104	0.100	0.000

N = 100    N Correct = 54    Proportion Correct = 0.540

*all linear discriminant function for group*

	Constant	L.W.	NL	L.S.	P(S)
1	1.036	1.767	-1.544	-0.148	0.001
2	1.920	1.174	-0.409	0.280	0.001
3	-0.768	0.054	-0.004	0.000	0.000
4	-0.070	0.000	-0.000	0.000	0.000
5	-0.000	0.000	0.000	0.000	0.000

Table 13. Linear discriminant analysis. Response: age of capture controls. Predictors: first values of log-transformed head-size variables. Sample size: 100 animals. Groups: 1–5 years old.

*all summary of classification with given relations*

Group	1	2	3	4	5
1	10	0	0	0	0
2	1	48	2	0	0
3	0	4	8	2	1
4	0	1	7	2	1
5	1	0	0	1	2
Total N	11	53	17	5	4
N Correct	11	43	0	3	2
Proportion	0.007	0.076	0.000	0.000	0.000

N = 101    N Correct = 57    Proportion Correct = 0.564

*all linear discriminant function for group*

	Constant	L.W.	NL	L.S.	P(S)
1	-0.000	0.000	0.000	0.000	0.000
2	-0.000	0.000	0.000	0.000	0.000
3	-0.000	0.000	0.000	0.000	0.000
4	-0.000	0.000	0.000	0.000	0.000
5	-0.000	0.000	0.000	0.000	0.000

Table 14 Linear discriminant analysis. Age-based age of capture groups. Predictive test subset of test variables. Sample size: 160 animals (Groups: 1 = 2 years old)

a) Number of discriminants with age-variables					
Group	1	2	3	4	5
1	17	26	6	0	0
2	3	39	3	3	0
3	0	4	2	1	0
4	0	9	4	2	1
5	0	0	1	0	4
Total N	20	78	13	3	5
N Correct	16	69	9	2	4
Proportion	0.800	0.885	0.692	0.667	0.800
N = 111    N Correct = 64    Proportion Correct = 0.576					

b) Linear discriminants function by group					
	Constant	W <sub>1</sub> /W <sub>2</sub>	W <sub>2</sub> /W <sub>3</sub>	W <sub>3</sub> /W <sub>4</sub>	W <sub>4</sub> /W <sub>5</sub>
1	164.9	-100.3	328.1	147.5	-205.3
2	-176.2	121.9	1144.4	843.1	-230.1
3	-632.5	154.9	2040.4	943.4	-266.4
4	-172.1	-491.7	2039.4	466.4	-104.0
5	76.4	1473.3	364.9	853.9	-184.0

Randomly in the pattern shown by the discriminant analysis of the whole set of variables there were significant variation among age-classes in terms of the proportion of correct classifications for the different types of variables. It varied from 0.473 to 0.933 for two-year-olds from 0.429 to 0.728 for three-year-olds, from 0.358 to 0.538 for three-year-olds, from 0.383 to 0.488 for four-year-olds and from 0.488 to 0.667 for five-year-olds (Table 14).

Log transformed variables provided slightly better results than their untransformed counterparts. Two-year-olds animals were the most distinct group from all elsewhere. Two-, three- and four-year-old groups were considerably close to each other. Two-year-old animals were also close to the five-year group. This means that the variation in the number of variables resulted in a slightly lower efficiency in discriminating animals from intermediate ages, although the same efficiency is reached in put the extreme age-classes assigned apart.

PCA, for the test subsets of untransformed (and raw) variables, log transformed raw variables and age-variables (Table 15) and their plots (Figure 15, 16, 17 respectively) are useful to visualize how much the different age-classes are clustered apart after the reduction in the number of variables. PC1 is responsible for most of the variance for the three kinds of variables, varying from 0.731 for raw to 0.974 for untransformed variables, which is around a-year similar to the result for the log transformed variables. This means that the overall size variation,

in the stepwise regression change for the age classes analysed, could be captured by the best subset of variables as well.

Similar results pattern observed by the PCA for the whole set of untransformed and log-transformed variables, the PC1 coefficients were more variable for the later than for the former (see PC1 in Table 1 for testing). This means that log-transformations is still more sensitive to the role of specific variables on the overall size change of the lipid area with a reduced number of variables. Index of total weight (ICW) is still the major source of variance in age class.

Table 11 Age distribution of system samples. Principal components analysis of best subsets of variables. Sample size: 141 animals

A) Eigenvalues of the eigenvalue matrix of best subsets of best size variables				
Eigenvalue	1 0000	0.0416	0.0034	0.0001
Proportion	0.954	0.042	0.003	0.001
Cumulative	0.954	0.996	0.999	1.000
Variable	PC1	PC2	PC3	PC4
ICW	-0.364	0.385	0.136	0.799
BL	-0.066	0.336	-0.005	-0.358
SW	0.002	0.015	0.236	-0.176
PC3	-0.009	0.007	-0.132	0.636

B) Eigenvalues of the eigenvalue matrix of best subset of log-transformed variables				
Eigenvalue	0.1441	0.00404	0.00134	0.00011
Proportion	0.944	0.007	0.001	0.001
Cumulative	0.944	0.991	0.992	0.993
Variable	PC1	PC2	PC3	PC4
ICW	-0.007	-0.109	0.196	0.447
ICW	-0.038	0.716	-0.001	-0.007
LCB	0.005	-0.107	0.400	-0.793
PC3	-0.009	-0.007	-0.006	0.006

C) Eigenvalues of the eigenvalue matrix of best subset of size variables				
Eigenvalue	0.001668	0.000498	0.0001444	0.000004
Proportion	0.793	0.237	0.025	0.004
Cumulative	0.793	0.944	0.970	1.000
Variable	PC1	PC2	PC3	PC4
ICW	-0.166	-0.108	-0.007	0.001
BL/ST	-0.002	-0.076	-0.000	-0.006
ICW/ST	-0.002	-0.062	-0.001	-0.001
SW	0.005	-0.003	0.000	0.001

PC1 extracted for 0.954 and 0.971 of the total variance the untransformed and log-transformed variables respectively. A contrast between PC3 and the other variables is clear (1.0

PC1 of the biplot (see PC1 in Table 1b). This pattern suggests a rearrangement of the relative part of the score contributed to the end of the biplot, so the score that first appears becomes orthogonally distinct (PC2 only appears right). PC2 of first subset of log-transformed variables is its first cluster a contrast between IQW and PC2. This may be relative to the ordering of the non-optimal division and the anterior region of the score, which is compatible with the results of the PCA for the first subset of transformed variables.

Once again, PC3 of total variables showed a complex pattern. There is a contrast in PC1 between the relative width of snout (SNWT) and the relative jaw-chord width (JWC). There may be a "nose" component in this contrast since it is related to two relative width indices, but no biological meaning is apparent. PC2 (see Table 1b) accounts for 8.3 % of the total variance contained in the PCA for total variables. It presents all coefficients with the same sign. SNWT and JWC are the major sources of variation again. However, in this case, they may represent a "vertical" shape change relative to the many variables described above, but that is also other complex and subjective.

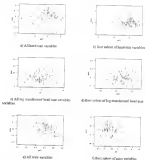
On Figure 7b, d, and f it is possible to see the distribution of variable against values for the two principal components of the first subset of untransformed, log-transformed, and total variables, respectively. It is also possible to compare these results with the ones presented by simple plots of the whole set of variables. Finally, it is possible to see the correspondence of transforming data and working with the same amount of raw measurements.

Log transformation of total variables did not significantly improve the results of any of analyses above. Total variables showed complex and subjective pattern in the principal components analysis of both the whole set and the first subset of variables. In addition, they were always related to the metric variation. One-year old animals were efficiently distinguished with both the whole set and the first subset of variables. Five- and three-year old animals, however, were classified more efficiently with the first subset of variables than with the whole subset. Three-year old group remained the same along the reduction of the number of variables, whereas two-year old metric presented more results with less values.

Table 14 presents correspondence between univariate and multivariate discriminant analysis of age in captive animals. Proportions of correspondence are presented for different combinations of variables and single variation for first discriminant characterizing

There are statistical experiments of the efficiency of both univariate and multivariate discriminant analysis when age classes are pooled together. Pooling three- and five-year old in either improved the results significantly better than pooling two- and three-year groups.

together. However, this results were obviously inconsistent, therefore, we did not join old records together. Considering the previous results, the procedure can be justified.



**Figure 7:** Age distribution of egyptus animals based on head measurements. Plots of egyptus animals'  $\log_{10}(\text{mm})$  against values for the four principal components.





0.014 (0.004) to 0.043 (0.01) for two-, three-, and four-year old groups together. Shortest snout lengths (SVL) also varied from 0.071 to 0.082 according to the age distributions and body mass varied from 0.006 to 0.077 under same circumstances. Multivariate discriminant analysis presents the proportion of correct estimates of age varying from 0.743 (all log-transformed and mass correlated to 0.776 (all head size variables with SVL) and from 0.714 (all head size variables with or without SVL) to 0.643 (all log-transformed head size variables) for two-, three- and four year old groups together.

The inclusion of SVL does not significantly improves the model. On the contrary, its efficiency slightly decreases when three- and four-year old animals are pooled together. This explains why growth curves of single variables are not efficient to predict age. Shortest snout length and body mass, the most robust body size measurements for anurans, are less efficient to predict age than for anurans based on snout length (SL), snout length (HL), or snout width (CW), which is surprising, considering that few studies have speculated would propose age estimation models with these variables. In many old with SVL, we find:

These results show that sexual dimorphism may explain more efficiently morphometric changes from when there is no consistent variation in body size. In other words, sexual dimorphism can be more related to development rather to growth, as the means that developmental index is the first coefficient for coefficient of discrete processes, while growth index is biological processes (body size).

The practical application of such concepts in the context is the capacity to discriminate well from young individuals individuals in what amounts to separate individuals from age-dependent morphological processes. External factors that influence growth rate, such as nutrition, environmental and others, would apparently be irrelevant to sexual dimorphism due to body size. If true, morphometric analysis of age discrimination of anurans should be more efficient to estimate age of wild animals than single measurements related to body size. This hypothesis is tested in the next sections as an exploratory exercise, using as wild animals with knowledge were available in the Poço, Brazil, at the time of the present study.

### 4.2.1.3 Age estimation of wild animals

The age of wild animals were estimated through the best linear regression of untransformed and log-transformed head size variables, and mass variables presented in Table

1) Repeating procedure for wild animals in the same or next for the capture months  $q = \text{age} + 1$  year = 1, 1 = age + 2 = 2, 2 = age + 3 = 3, 3 = age + 4 = 4 and 4 = age + 5 = 5.

Lasso regression analysis with cross-validation was run for the non-transformed head size variables (Table 17), log-transformed head size variables (Table 18) and mass variables (Table 19). The estimated/rounded age of wild animals was the response (dependent variable) and the best subset of variables selected for capture months were the predictors (independent variables).

The two-year old group presented a maximum (1.000) proportion of correct estimate with all kinds of variables. The three-year old group also presented a maximum proportion of correct estimate with the log-transformed variables. The three-year old animals were the only ones to present cross-validation results both with log-transformed and mass variables.

Table 17. Lasso regression analysis: Response: estimated age of wild animals. Predictors: best subset of head size variables. Sample size: 25 animals. Groups: 2, 3 years old.

a) Summary of classification with cross validation		
Group	1	2
1	13	6
2	6	17
Total N	19	23
N correct	13	17
Proportion	1.000	1.000
N = 25    N correct = 24    Proportion correct = 1.000		

b) Least dependent feature in the group					
	Common	CV	SE	FW	Padj
1	-0.456	1.027	-0.143	1.000	1.000
2	-0.432	0.232	0.004	1.000	1.000

The overall proportion of correct classification, even with cross-validation procedure, varied from 0.750 for the log-transformed variables to 1.000 (exactly 100%) for the non-transformed variables, with mass variables presenting an intermediate value of 0.880, which is surprisingly high. Age-classes of two and three years were found with non-transformed and mass variables, whereas age-classes of two, three and four were found by the best subset regressions of log-transformed head size variables.

The differences between non-transformed and log-transformed head size variables are primarily due to the different subset of variables employed. Most of the variables (17 of 25) received the same age classification, except of the kind of variable used (transformed head

one variable was possibly the last variable added because its last entry appears to have the highest coefficient of determination.

Table 18: Least-squares analysis: Response: estimated age of wild animals. Predictors: last subset of log transformed head size variables. Sample size: 29 animals. Groups: 2–3 years old

*A) Summary of classification with cross-validation*

Group	1	2	3
1	0	0	0
2	0	10	0
3	0	0	0
Total N	0	10	0
N Correct	0	10	0
Proportion	1.000	1.000	1.000

N = 29    N Correct = 21    Proportion Correct = 0.724

*B) Least-squares function for group*

Constant	CV	CVW	PCA	PAB	
1	-104.7	1194.2	0.813	796.0	-0.019
2	1029.7	1470.8	-100.1	1003.7	0.021
3	1233.2	1490.4	70.04	1070.0	-0.011

Table 19: Least-squares analysis: Response: estimated age of wild animals. Predictors: last subset of scale variables. Sample size: 29 animals. Groups: 2–3 years old

*A) Summary of classification with cross-validation*

Group	1	2
1	0	0
2	0	10
Total N	0	10
N Correct	0	10
Proportion	1.000	1.000

N = 20    N Correct = 20    Proportion Correct = 1.000

B) Least-squares function for group				
	Constant	CV	CVW	PCA
1	1428.0	1617.0	404.4	1007.1
2	1338.8	1761.0	404.8	1020.0

PCA of untransformed and log transformed head size and mass variables (Table 20) and their plots (Figure 1) are useful to see how much these estimated age groups are clustered apart. PC1 accounts for much of the variation both for untransformed and log-transformed variables (see PC1 in Table 20a and b), varying from 81.3% on the left to 81.5% on the right. That means

that size change in the major crania of variation between different age classes is well most apparent for temporal variance. PC1 coefficients are positive in magnitude and sign for anatomical variables (see PC1 in Table 20a) varying unidirectionally from less < 40% (FNS) to > 50% (CW). PC1 coefficients for log transformed variables (see PC1 in Table 20b) are – have even more variable in magnitude, ranging in absolute terms from > 10% (LCR) to < 2% (CW).

PC2 accounts for > 11% for anatomical variables and for > 80% for log transformed variables (see PC2 in Table 20a and b). PC2 of anatomical variables shows a clear contrast, both in terms of magnitude and sign, between FNS and the combination of CW and LCR. The biological meaning of this may be a possible reorganization of the anterior part of the cranium in relation to the width of the cranium and the orbit, or the sense that the face becomes uniformly broader (CW and LCR with posterior eye) in relation to the brain (FNS with anterior eye).

**Table 20. Age variation of wild animals. Principal components analysis of four subsets of variation. Sample size: 20 wild animals**

**a) Eigenanalysis of the covariance matrix of four subsets of raw variables**

Variables	1 (51.0)	2 (19.0)	3 (9.7)	4 (38.4)
Proportion	0.198	0.010	0.000	0.000
Combustion	0.170	0.000	0.000	0.000
Variable	PC1	PC2	PC3	PC4
CW	-0.080	0.080	0.000	0.000
SL	-0.080	0.000	-0.000	0.000
LCR	-0.080	0.000	0.000	0.000
FNS	-0.000	-0.000	0.000	-0.000

**b) Eigenanalysis of the covariance matrix of four subset of log transformed raw variables**

Variables	1 (51.0)	2 (19.0)	3 (9.7)	4 (38.4)
Proportion	0.198	0.010	0.000	0.000
Combustion	0.170	0.000	0.000	0.000
Variable	PC1	PC2	PC3	PC4
CW	-0.080	0.080	-0.000	-0.000
SL	-0.080	0.000	0.000	0.000
LCR	-0.080	0.000	-0.000	0.000
FNS	-0.000	-0.000	0.000	0.000

**c) Eigenanalysis of the covariance matrix of four subset of raw variables**

Variables	1 (51.0)	2 (19.0)	3 (9.7)	4 (38.4)
Proportion	0.198	0.010	0.000	0.000
Combustion	0.170	0.000	0.000	0.000
Variable	PC1	PC2	PC3	PC4
CW	-0.080	0.080	-0.000	-0.000
SL	-0.080	0.000	-0.000	-0.000
LCR	-0.080	0.000	-0.000	-0.000
FNS	-0.000	-0.000	0.000	0.000

PC2 of log-transformed variables above, on the form, a direct contrast between FWH and KWF, is the sense that the more robust variable increases relative to the anterior part of the rostrum (KWF with positive sign and FWH with negative sign). This is compatible to the pattern described above, and reinforces the idea that morphologic changes occurring during the normal and possibly final years of age lead to a widening of the anterior portion of the rostrum in relation to the anterior part of the orbit. This pattern is in contrast from the pattern observed during the first year of age (the pattern described by Block, 1954) that means year old animal was captured, according to the values of age determinations used in this study.

Factor variables were used principal coefficients of different ages on PC1, tested the same age on PC2. The biological interpretation of this pattern is that again robust morphology and subjective PC1 measures for 1/3 of the total variation, whereas PC2 is responsible for 2/3 of KWF and the major source of variation on PC1, whereas KWF was the major source of variation on PC2. KWF is statistically more influenced by body-size than FWH (Table 2) (Figure 10).

Figure 8 shows the plot of wild animals, measured age against values for the two principal components of the first subsets of untransformed head-size variables (Figure 8a), log transformed head-size variables (Figure 8b), and size variables (Figure 8c). In all of these subject separation between age classes is visible.

On Figure 8a and b it is possible to consider change classes can be shown separately separated by PC1. This means that overall size change must be the main source of variation between the age-classes tested. This seems to be compatible with the more pronounced tendency of head sizes and impressions that wild animals experience. On Figure 8c, as the other head size-classes was only separated by a combination of values of PC1 and PC2. It is necessary to compare the results above with similar analysis for captive animals (Figure 9) and also with analysis of sexual dimorphism (Figure 11) and tests of origin of wild animals (Figure 12).

#### 4.2.2 Allometric Relations of Captive and Wild Animals

Allometric relations can be studied in the comparison of body size from isolated measures of parts of the body. Population monitoring of zooplankton usually involve single measurements usually only the length of animals are recorded. Thus, the relationship between length of head and total length is usually employed to establish size class distributions for the target populations. An

for example, Cholewicki (1964) suggested that the distance between the eye and the tip of the snout is taken as a measure of the total length of *Chirocentrus maculipinnatus* in fish. Champman and White (1977) propose a photographic method to measure total length of *Cynoscion pomona* from lateral photographs. It is not to improve these methods (Mogensen (1987) suggests that a sample of animals should be captured and measured. Thus, relationships between measured and actual measures – differences could be established and observed – bias could be corrected. The interesting point of this method is that it permits quantification of the actual observed – bias.

Although allometric equations may lead to convenient and useful generalizations, there are important limits concerning their use. Schmidt-Nielsen (1984) establishes the following points concerning them:

“1. Allometric equations are descriptive; they cannot be biological laws.”

“2. Allometric equations are useful for showing how a variable quantity is relative body size, all other things being equal; but usually, actually, they are not!”

“3. Allometric equations are relationships because they only connect principles and measurements that otherwise neither connect.”

“4. Allometric equations are useful not from the convenience and not merely because there is general pattern. Such deviations may be due to nature or may reveal a significant underlying agent.”

“5. Allometric equations are useful for measuring unexplained variation in the same variable, as weight is a function, but a poor body size.”

“6. Allometric equations cannot be used for interpolation beyond the range of the data on which they are based.”

In the present study, allometric relations of capture and wild BIC are presented (Table 2) showing the analysis of covariance (ANCOVA) of sex and age of capture animals in relation to the allometric equations between teleostomere number and snout-vent length (SVL). In this context, ANCOVA is used to compare males and females allometric equations. It is also useful to separate age class body size effect on the teleostomeres, sex effect analyzed. In later context, it is used to compare adult and sub-adult effect on allometric relations of bodylength and environmental effect on the allometric relations of wild and capture animals of slender body size.

All body and head size variables and all but three mass variables (BW, LBN and BPW) are significantly affected by body size ( $P$ -value  $< 0.100$ ), or in other words, they can be considered size-dependent. One body size (BW), skeletal size (CL, SL, OL, CPW, PCL and WBL) and one size variable (BLA) are significantly affected by age ( $P$ -value  $< 0.100$ ), so they can be considered age-dependent. At last one body size (BW), two head size (CL and CPW), and two mass variables (PCW, ELST, BCL, BCPW and BPW) are significantly affected by gender ( $P$ -value  $< 0.100$ ) (see Table 1).

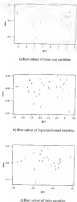


Figure 8. Plots of total values estimated by year (year) against values for the two principal components.

Whole and Mixed (WT) export a geographic record description as *Chenopodium* species involving interstitial soils, which is not present in the recent study. Half and Full export the final record description for 11 of 14 soil attributes, including DCL, SL, FOL, CW, OL, BOW, WCL, WY, and WSL. However, their results are possibly optimistic because they would not include age as a covariate of body size in their study of the same growth of *Chenopodium* arrangement. Some variation actually caused by age (independent of size) may be erroneously associated with differences between sites, or actual phenology. A moderate approach for the study of sexual dimorphism is presented in Section 4.2.4.

Table 20. Analysis of variance: Age and size as covariates of WY. (R-squared)

Variable	WY	Age	Sex	Variable	WY	Age	Sex
SL	0.000	0.000	0.000	SLW	0.000	0.000	0.000
FW	0.000	0.000	0.000	SLST	0.000	0.000	0.000
DCL	0.000	0.000	0.000	BWST	0.000	0.000	0.000
CW	0.000	0.000	0.000	BOL	0.000	0.000	0.000
OL	0.000	0.000	0.000	BOW	0.000	0.000	0.000
OW	0.000	0.000	0.000	WCL	0.000	0.000	0.000
CL	0.000	0.000	0.000	WST	0.000	0.000	0.000
OW	0.000	0.000	0.000	WY	0.000	0.000	0.000
BOW	0.000	0.000	0.000	WSL	0.000	0.000	0.000
LCL	0.000	0.000	0.000				
WY	0.000	0.000	0.000				
FOL	0.000	0.000	0.000				
SL	0.000	0.000	0.000				
LWS	0.000	0.000	0.000				
WY	0.000	0.000	0.000				
WY	0.000	0.000	0.000				

Model predictor: Sex is WY/SL = General Linear Model

Response (dependent variable): dependent variable

Model (independent variable): WY

Covariates: Age and Sex for regression analysis

The fact that all age-dependent variables are also size-dependent explains why it is not difficult to predict age of intertidal based on single variable growth curves (see Section 4.2.2 for discussion). All of the size-dependent variables are also size-dependent, with the exception of BOW. However, its efficiency in predicting intertidal size, through discriminant analysis is low. Four size-dependent variables (OL, OW, and BOL) are also age-dependent, but the remaining four of all size-size variables (BOW, BLSL, BOW, and BWSL) are not. Age-dependent as well as size-dependent variables are generally located in the anterior. Only one age-dependent (BWSL) will not independent variable is a located in the middle.



Alternative relations of wild animals are presented in Tables 22–23 and Figures 9–10. Table 22 and Figure 9 show the alternative relations between body- and head size variables and the roost-roof length (RRL). Table 23 and Figure 10 show the alternative relations between roost variables and RVL. Disturbance-relatively small sample size, wild animals and females are presented together. Table 24 and Figure 11 show the alternative relations between body- and head size variables and the roost-roof length (RVL) as captive animals. Table 25 and Figure 12 show the alternative relations between roost variables and RVL as captive animals.

Table 22. Alternative relations of wild animals: Body- and head size variables

	Sex	N	R	r	t	r	Adjusted	p	%
1	wt	SVL	SVL	0.6666	0.6666		0.4602	0.674	24
2	wt	BL	Log BL	0.6430	0.7548		0.4325	0.555	26
3	wt	SVL	BL	0.2222	0.1999		0.6966	0.498	25
4	wt	SVL	BL	0.7057	0.4834		0.5051	0.568	26
5	wt	SVL	CV	0.1566	0.4756		0.6862	0.174	26
6	wt	SVL	SL	0.1566	0.1281	0.4474	0.6662	0.468	27
7	wt	SVL	CV	0.2683	0.2687	0.3439	0.6491	0.677	25
8	wt	SVL	SL	0.4874	0.4438		0.5661	0.626	26
9	wt	SVL	CV	0.6159	0.1775	0.1463	0.5652	0.664	26
10	wt	SVL	CV	0.5833	0.6833		0.5661	0.679	25
11	wt	SVL	CV	0.6430	0.6430		0.5661	0.662	26
12	wt	SVL	SVL	0.2479	0.7548		0.6966	0.495	26
13	wt	SVL	SVL	0.1887	0.6546		0.6966	0.497	25
14	wt	SVL	SL	0.7766	0.7610		0.6966	0.567	26
15	wt	SVL	LM	0.4607	0.5855		0.6966	0.558	26
16	wt	SVL	SVL	0.4408	0.7137		0.6966	0.562	26

$$Y = a + bX + cX^2 + dX^3$$

wt: males and females

Model procedure: Step → Regression → Forward Stepwise (Polynomial Regression)

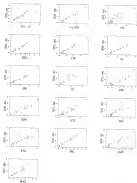
Note: the exception of BL variable was not introduced because their values all magnitude were similar and transformation did not improve results

Value excepted values included in the equation ( $p < 0.1$ ) otherwise significant ( $p$ -value  $< 0.05$ )

Quadratic element ( $X^2$ ) was included in the equation ( $p < 0.1$ ) otherwise cubic quadratic or cubic element were significant ( $p$ -value  $< 0.05$ )

With the exception of body mass (BL), log transformation did not improve the alternative equations for either wild or captive animals. Alternative equations for captive animals presented higher coefficient of determination ( $r^2$ ) than the ones for wild animals. Body- and head size variables presented a significantly higher  $r^2$  than roost variables for both wild and captive animals. They varied from 0.438 (BL) to 0.675 (CV) with body- and head size variables (Table 22) and from 0.460 (BLTX) to 0.687 (BLTX) with roost variables (Table 23) for wild

animals. For capture animals, at three years they varied from 0.93 (276) to 0.93 (287) with body- and head-size variables (Table 18) and from 0.903 (30,482) to 0.914 (30,517) with size variables. The range of SVL relative to snout-vent length can be found on the plot of Figures 9-12.



**Figure 9.** Plots of allometric relations of initial animals' body- and head-size variables (Log SVL log-transformed SVL, SVL, and SVL) to SVL (the others as cov) (see Table 12 for regression equations).

Table 12. Adequacy criteria of regression models

i	Seq	r	R	s	t	F	2	Adjusted	F <sup>2</sup>	n
1	seq1	0.91	0.79	-49.8336	131.000			0.989	0.979	24
2	seq1	0.91	0.78	1999.85	-62366.4	1.64811E	-69221.1	0.989	0.989	24
3	seq1	0.91	0.8037	79.0000	-48.1000			0.990	0.990	24
4	seq1	0.91	0.74	10.0000	238.000			0.989	0.979	24
5	seq1	0.91	0.79	14.2500	24.0000			0.990	0.989	24
6	seq1	0.91	0.91	-0.2700	902.910			0.990	0.997	24
7	seq1	0.91	0.90	0.0000	900.000			0.979	0.979	24
8	seq1	0.91	0.90	-4.1000	899.000			0.979	0.979	24
9	seq1	0.91	0.90	24.0000	899.000			0.979	0.979	24
10	seq1	0.91	0.90	-12.2500	899.000			0.979	0.979	24

$$Y = a + bX + cX^2 + dX^3$$

seq1: water and biomass

Model generation: Seq → Regression → Total Linear (Polynomial Regression)

Water component of (Seq) variables were not transformed because their values of response were similar and transformation did not improve results.

Collinearity test (V) was included in the regression (V) criterion significant (P-value < 0.05).

Quadratic component of transformed water response (V) + (V) criterion either quadratic or cubic function were significant (P-value < 0.05).

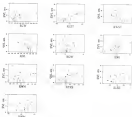


Figure 10. Fitted elements values of water and biomass Seq variables. See Table 12 for regression equation.

**Table 26.** Allometric relationships of captive sea otters (body- and head size variables).

<i>i</i>	Sex	<i>T</i>	<i>K</i>	<i>a</i>	<i>b</i>	<i>c</i>	<i>d</i>	<i>E</i> <sub>max</sub>	<i>r</i> <sup>2</sup>	<i>N</i>
1	mf	TL	SNL	1.8896	1.1137	-0.0620		0.000	0.991	126
2	mf	SNL	LogSNL	13.9050	26.8661	1.0000		0.000	0.999	126
3	m	SNL	SNL	-1.2264	1.3686	-0.0001		0.000	0.999	23
4	f	SNL	SNL	3.0843	1.3679	-0.0004	0.000000	0.000	0.999	91
5	mf	SNL	SNL	-0.1648	0.7233			0.000	0.991	126
6	mf	SNL	L <sup>3/4</sup>	-0.2642	0.7817	-0.0004		0.000	0.991	126
7	mf	SNL	SL	0.1348	0.7611	-0.0001		0.000	0.991	126
8	mf	SNL	SNL	0.1797	0.7620	-0.0001		0.000	0.991	126
9	m	SNL	SL	0.1182	-0.1094	0.1097	-0.0011	0.000	0.992	23
10	f	SNL	SL	0.1179	0.0891	0.1036	-0.0104	0.000	0.998	91
11	m	SNL	SNL	24.7948	0.0049			0.000	0.998	23
12	f	SNL	SNL	13.1000	0.0036	0.1000	-0.0110	0.000	0.998	91
13	mf	SNL	SNL	1.2060	-0.1831			0.000	0.998	126
14	mf	SNL	L <sup>3/4</sup>	17.6104	0.0100			0.000	0.991	126
15	mf	SNL	SNL	3.3504	0.4136			0.000	0.991	126
16	mf	SNL	FL	-0.0104	1.0370	-0.0017		0.000	0.992	126
17	mf	SNL	SL	-0.0761	0.8818			0.000	0.992	126
18	mf	SNL	SNL	-0.0036	0.9171	-0.0004		0.000	0.995	126
19	mf	SNL	SNL	0.0144	0.8234	0.0011		0.000	0.998	126
20	mf	SNL	SL	0.2481	0.5087	0.0001	-0.0001	0.000	0.987	91

$$T^{(i)} = 10 + 20^{(i)} - 10^{(i)}$$

Sex: m=male, f=female, mf=both sexes combined.

Model procedure: Step 1=Regression, 2=Fixed-Last Plot (Polynomial Regression).

Note the exception of item 10: regression was performed with/without that value of magnitude zero, under and without/without that regression results.

Extra chosen (0.000) included in the equation (i = 0) otherwise: regression (P-value < 0.05).

Quadratic element (c) was included in the equation (i = 4) otherwise: either quadratic or cubic element were significant (P-value < 0.05).

Extra and chosen presented separately when ANOVA for sex was significant (P-value < 0.05) for Table 13 for P-values.

The coefficients of determination of wild and captive female concerning body- and head size variables can be considered extremely high. Their sizes biological meaning is the opposite: lack of morphological variation in the patterns studied, which could be expected for captive but not for wild animals. They also considered most of the head size variables studied can be useful for predicting body length. This can be particularly interesting for the study of marine invertebrates, or more precisely mammals, in which usually only crania are preserved or found relatively intact. Therefore, the present study helps with wild individuals, which can necessary to increase the range of data or other conclusions than in other studies.

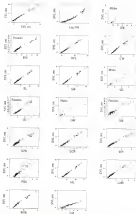


Figure 11. Plots of allometric relations of log<sub>10</sub>(RF)-log<sub>10</sub>(L) for *B. p.* and listed size variables. RF for leaf from the proximal negative values (under the water) (see Table 7) for *ANDRYA P.* (see Table 24 for regression equations).

Table 25. Allometric relations of capture per individual. Same variables

		1	2	3	4	5	6	7	8	9
										value
1	m	34%	82%	134.881	247,644					0.000
2	f	34%	82%	8932.47	24702.4					0.000
3	m	34%	82.87	2.77434	-13.188.38	1540.30				0.000
4	f	34%	82.87	1586.44	-33,071.7	38270				0.000
5	m/f	33%	82%33	-4934.89	12870	3354.74	3407.43			0.000
6	m	33%	82%	240.144	-1080.40	4834.37				0.000
7	f	33%	82%	-149.483	863.46	-4013.4	44711.3			0.000
8	m	33%	82%8	124.645	1074.1					0.000
9	f	33%	82%8	37.1133	74.9233					0.000
10	m/f	33%	82%	184,249	-8844.34	7464.4	7464.33			0.000
11	m	33%	82%8	-100.80	547.146					0.000
12	f	33%	82%8	171448	2884.7	17176.1	14346			0.000
13	m/f	33%	82%8	201.332	2177.36	1844.46				0.000
14	m/f	33%	82.88	20.1445	144.445					0.000
15	m/f	33%	82%8	144.444	1771.72	1444.74				0.000
16	m/f	33%	82.148	24.4448	34.4707					0.000

$$R^2 = 0.62, r = 0.79, r^2 = 0.62$$

Box-cox makes  $\beta$ -function,  $\alpha$ -function and function

Minimum prediction: Box-cox Regression  $\rightarrow$  Fixed Line Plot/Polynomial Regression

Whether captures of fish variables selected by model from those variables of importance were similar and log transformation did not improve results

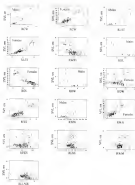
Cohen's kappa (K) was calculated  $\rightarrow$  the agreement of all reference significant (P-value  $< 0.05$ )

Quadratic equation (y) was tested in the equation  $y = ax^2 + bx + c$  where  $a$  is the quadratic coefficient,  $b$  is the linear coefficient and  $c$  is the constant term

Model and function presented separately when ANCOVA for sex was significant (P-value  $< 0.05$ , see Table 26 for F values)

Some predictions in red and white using allometric relations of same variables. Series of three regression equations are not statistically significant (P-value  $> 0.100$ ). This is the case for the following variables: BLPW, BLGS and BW for white and BLPW and BLGS for capture anomaly. First in Figure 12 - 13 help to visualize these patterns.

Results helping to estimate body size from head dimensions (the allometric equations above show that there changes during the recruitment process. First linear equations suggest changes in the proportions of the skull "stockiness" or "taperedness" in the different years. For instance, the problem of capture anomaly becomes relatively narrower in body size increases (see plot of CW in Figure 13). A similar trend expected pattern can be seen in the residuals (see plot of RW in the same figure). In both cases allometric equations are quadratic with the coefficients of the quadratic element being negative (see Table 34).



A common but equivocal slope can be perceived in the allometric growth curves of the standard length (SL) and weight (CW) in aquatic animals. A positive quadratic and a negative cubic element in the allometric equations of both cases show a period of fast relative growth in young followed by a period of slow relative growth of these regions in adult animals. However, the smaller coefficient of the linear element of the CW equation than of the SL equation represents the ontogenetic process of "disproportion" suffered by the organs during initial development of the animals. The pattern is compatible with the growth curves of these variables (see Figure 3 and section 4.2.1 for discussion).

#### 4.2.4 Sexual Dimorphism

Sexes phenotypically determined in amniotes basically by the incubation temperature of their eggs (Gall 1940; Williams et al., 1991). However, there seems to be a greater clutch effect, defined by Long and Andrews (1994) as the variance in the temperature and determination (SD) system and expressed as the clutch of origin<sup>2</sup>. There is a both lower and higher, but under a intermediate stable composition of metabolism (total PM) percent ratio to its predecessor (Long and Andrews, 1994). Alligator *osteosaurus* (Ferguson and Austin 1982, 1983) and *Crocodon crocodilus* (Long et al., 1993) were usually below or greater PM percent (phenetic) lower and higher (epigenetic) temperatures of incubation<sup>3</sup>. However, Long and Andrews (1994) show that this is incorrect.

Temperature of incubation influences morphology development and consequently metabolism period (Schmalzer et al., 1998). Incubation: it may also affect post-hatching growth rates (Austin et al., 1987 and Webb and Cooper Peters 1989), which is probably related to diet preparation and homeostatic behavior (Llorente and Ferguson 1993). Besides temperature, the following factors may possibly affect embryo development and post-hatching growth rate: egg size (Guder and Pedroni 1983), egg volume and feeding rate (Schmalzer and Chabreck 1996).

Now the factors above affect sexual dimorphism mainly relative to amniotisms, possibly because amniotisms in general do not present an evident external sexual dimorphism. Direct measurement of the genotype to the climate is described next and/or not these animals (Chabreck 1993, Allread and Long 1993).

<sup>2</sup> It is characterized by whether parent study, clutch effect a cooperative concept not only parents but also environment affect on the morphology of eggs and hatchlings due to the clutch of origin.



Although, there are some qualitative changes to detect for secondary sexual characteristics in crustaceans (McIlwain 1995, Yaman 1999, Smith 1997), few multivariate analysis were done to relate to secondary sex characteristics. Using linear discriminant analysis of head measurements and ratios, Webb and Blaney (1978) report a proportion of 97.5 to 100% correct sex classification of *Chironomus tentans* with varying body size (from 40 to 60 mm TL), as well as PVL, respectively. Ball and Pomeroy (1978) compared two methods of multivariate analysis (discriminant analysis and classification tree analysis), using variables based on the rostrum skull, antennae and mandible. They report results varying from approximately 54–75 to 92.5% of correct classification for the discriminant and from 5 to 100% for the classification tree analysis (the most conservative due to great validation procedure).

In the present study, a multivariate approach is presented in order to quantify (sexual) dimorphism in wild and captive BPC. The questions currently presented (Figure 3) are tested, and the implications of other sources of morphometry to variation (e.g., age, clutch, maternal effects, and environment) on sexual dimorphism are discussed.

#### 4.2.1 Wild animals

Possibly due to the relatively small sample size and also to the fact that all captured individuals were not wild, sexual dimorphism in wild animals was statistically detected by Wilk's Lambda, Hotelling, and Pillai's Trace of MANOVA ( $P$ -value = 0.001 only for non-transformed head measurements;  $P$ -value > 0.100 for log-transformed and ratio variables). For the present, subsequent analyses are just made with non-transformed head size variables.

Linear discriminant analysis of the head of variable presented a relatively high overall proportion of correct classification (81.79%, higher for males (88.88%), but significantly lower (62.10%) for females (Table 26). This pattern may reflect the male biased distribution of the sample (28 of 29 individuals).

PCA of the same set of non-transformed variables (Table 27) showed a considerable amount of variance in PC1 (37.4%), with 9.8% due to PC2. This represents a significant difference between sexes for the non-transformed size. PC1 presents all coefficients with similar magnitude (from 4.361 for OL to 0.183 for BCL and CW) with negative sign.

Some "visual" shape change can be interpreted as the contrast between orbital width (OW) and its association between orbital width (OL) and inter-orbital width (IOW). These variables present the largest coefficients in absolute value in the second component. Thus

increased eye openness means a possible widening of the eye orbit relative to the orbital length and interorbital width.

**Table 26: Linear discriminant analysis: Response use of wild animals. Predictors: all head-size variables. Sample size: 28 animals. Groups: omnivores and herbivores.**

Statistics of distributions with group relations

Group	Mean	Stdev
		<i>n</i>
Herbivores	26	7
Omnivores	2	1
Total <i>N</i>	28	8
SD (mean)	26	1
Response	0.100	0.100

*N* = 28    *N* Cases = 28    Response (mean) = 0.100

Linear discriminant function for group

	Constant	PC1	PC2	PC3	PC4	PC5	PC6	PC7	PC8	PC9	PC10	R <sup>2</sup>
Herbivores	-100.00	0.75	-0.08	-0.08	0.13	0.01	-0.70	0.06	0.01	0.01	0.01	0.01
Omnivores	-100.00	1.00	-0.00	0.00	0.00	0.00	-0.70	0.06	0.01	0.01	0.01	0.01

	PASS	MS	LAGS	WSS
Herbivores	0.00	0.13	0.13	0.13
Omnivores	0.00	0.13	0.04	0.04

**Table 27: Principal component analysis: Eigenanalysis of the covariance matrix of all head-size variables. Sample size: 28 animals. Only the six first principal components are presented.**

Component	PC1	PC2	PC3	PC4	PC5	PC6
Eigenvalue	11.260	0.188	0.045	0.114	0.077	0.061
Proportion	0.449	0.008	0.002	0.005	0.003	0.002
Cumulative	0.449	0.457	0.459	0.464	0.467	0.469
Variable	PC1	PC2	PC3	PC4	PC5	PC6
SkL	-0.191	-0.005	0.004	-0.026	0.001	0.000
OrW	-0.281	-0.000	-0.001	0.000	0.004	0.000
IL	-0.283	-0.000	0.000	0.000	0.000	0.000
OrW	0.100	-0.000	-0.000	0.000	-0.001	0.000
OrL	-0.000	-0.000	0.001	-0.000	0.000	0.000
OrW	-0.100	0.000	0.000	-0.000	-0.000	0.000
PC2	-0.271	-0.000	-0.000	-0.000	-0.000	0.000
LCR	-0.170	0.000	0.000	-0.000	-0.000	0.000
W	-0.170	-0.000	-0.000	0.000	0.000	-0.000
PC3	-0.270	0.000	-0.000	0.000	-0.000	0.000
IL	-0.000	-0.000	-0.000	0.000	0.000	0.000
LAGS	-0.000	0.000	-0.000	-0.000	0.000	0.000
WSS	-0.270	-0.000	-0.000	0.000	-0.000	0.000

Our study also re-examination of the natural regime, presents the contrary tendency, based on growth studies of captive animals (see Table 4 and Section 4.2.2) for discussing these principal components analysis data are two independent non-dependent variables, this may not be attributed to sexual dimorphism as this was not categorically change in the literature on mortality sampling. Another possibility is a difference between wild and captive animals (see Section 4.4.2), as the fact that wild animals captured are young, only, whereas the captive animals sample also included adult individuals. A stronger relation between DL and KPW and SVL as captive animals were represented by values equations with a slightly sigmoid shape (Table 5 and Figure 11). This could explain an approximately linear growth pattern involving young wild and adult captive animals.

PC1 shows particularly to much variance in PC1 (84.11 and 8.807 respectively). The relation between DL (-0.408) and KPW (0.122) portrayed a linear relationship after that the re-examination of the specimens in the literature is responsible for most of the shape change in this group. Although this may or may not be caused by sample composition change, it is impossible to exclude how much this problem is related to sexual dimorphism.

Table 28 presents the best subset of total test variables for the performance in wild animals. In order to keep the coefficient of determination ( $R^2$ ) relatively high, only two variables were eliminated from the models (SW and NL). This procedure resulted in a statistical improvement in the models of Wilk's  $\lambda$ , Levene's testing, and F-test a "lowest MANOVA's  $F$  value" = 0.001 for all of them). This possibly means that SW and NL were more "noisy" in the model. In other words, these variables introduce more "noise" variation in the system, not related to the morphometric differences between males and females.

Linear discriminant analysis for the best subset of transformed total test variables (Table 29) shows an improvement in the overall proportion of correct classification (8.80% and higher for males (9.88%) than for females (8.10%). The upper extent of the efficiency of the discriminant analysis corroborates the idea above.

PCA of the best subset of variables presents similar results to the analysis run for the whole set of variables, probably because only two variables were taken off the system (Table 30). PC1 accounts for 84.11% of the total variance, whereas PC2 is responsible for only 8.807. PC1 presents all coefficients with similar magnitudes (0.294 for SW and KPW and 0.109 for DL and CW) and negative signs.

Table 28. Best values of variables for sexual dimorphism in wild mango

Variable	Best values										Female	$r^2$	N
Best values	DCI	CB	SL	OL	LWR	OPW	LGR	WPI	PGL	LGR	WPI		
Best values	100.5	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	1.0	0.97	21

Statistical procedure: Test = Regression vs. Discrimination

Sample size: 21 animals

Table 29. Linear discriminant analysis. Regression test of best variables. Predictors: Best values of best variables. Sample size: 21 animals. Groups: males and females

#### Discrimination of coefficients and cross validation

Group	1	2
1	11	1
2	1	2
Total N	12	1
N Correct	12	1
Proportion	1.000	1.000
N = 13	N Correct = 13	Proportion Correct = 1.000

#### Linear discriminant function by group

	Constant	DCI	CB	SL	OL	LWR	OPW	LGR	WPI	PGL
Males	104.11	0.00	-0.11	-0.12	0.00	0.00	0.00	0.00	0.00	-0.11
Females	120.44	0.00	-0.02	0.00	0.00	0.00	0.00	0.00	0.00	0.00
	0.00	0.00								
Males	-0.00	0.00								
Females	0.00	0.00								

Table 30. Principal components analysis. Eigenanalysis of the correlation matrix of best values of best variables. Sample size: 21 animals. Only the six first principal components are presented

Component	0.325	0.287	0.198	0.188	0.077	0.058
Proportion	0.000	0.007	0.001	0.000	0.007	0.001
Discrimination	0.001	0.008	0.076	0.000	0.007	0.002
Variable	PC1	PC2	PC3	PC4	PC5	PC6
DCI	-0.100	0.079	-0.100	-0.100	0.000	0.000
CB	-0.101	-0.179	0.000	-0.100	0.000	0.000
SL	-0.101	0.000	-0.101	-0.100	-0.100	-0.100
OL	-0.100	0.000	-0.100	0.000	0.000	0.000
OPW	-0.100	0.000	-0.101	0.000	0.000	-0.100
LWR	-0.100	-0.100	0.000	0.000	0.000	0.000
LGR	-0.100	0.000	-0.100	0.000	0.000	0.000
WPI	-0.100	-0.100	0.000	-0.100	0.000	-0.100
PGL	-0.100	0.000	0.000	-0.100	0.000	0.000
LWS	-0.100	0.000	-0.100	0.000	0.000	0.000
WPI	-0.100	-0.100	0.000	-0.100	0.000	-0.100

Figure 3.1 shows the plots of wild animals against values for the two principal components. Although there seems to be some differentiation between males and females along the second component "timidity" For a complete answer, the first possible explanation for this grouping pattern would be a possible mistake on the survey method used in the field. However, there are other sources of variation in the overall morphology of these animals, such as age and sex of origin. Thus, it would be prudent to check these possibilities in the other sections of the present study.



(a) All wild animal variables



(b) First values of females variables

Figure 3.1. General description of wild animals. Plots of wild animals (PC1) against values for the two principal components (1) males, (2) females.

The influence of series of origin is analysed in Section 4.4.1. Sex-lens animals are clustered in Figure 2b in relation to their series of origin. Although there is some clear separation between samples from Dargahpur (line 1 in the figure) and the others, their distribution is more less clear than for the distribution of animals and females in the plots of Figure 1c.

Age of wild animals was estimated through the best choice of variables determined by capture marks (see Figure 4 and Section 4.2.2 for discussion). Although the score  $t$  is well designed for scoring age-effect for image effect, there is a correlation, although unaccounted, between the age (specifically increasing weakness) and in the larger series of variables for wild animals. If true, this introduces a confounding factor in several descriptive studies of correlations. In order to reduce this effect it is necessary to keep all other things equal, e.g. age, body size, etc. which is almost impossible in capture and management, just to narrow the variables included in this study. This is practically impossible in field studies and very difficult even in captivity as it is preferable not to do the last action.

#### 4.2.4.3 Capture animals

A significant difference between capture males and females was found by MANOVA ( $F$ -value >0.000 on WILK's, Lawley, Hotelling, and Pillai's Tests for the selection of the three kinds of variables: untransformed and log transformed head-size and ratio variables). This means that the variance among groups is bigger than the variance error group, which means the first question presented in Figure 3.

In order to evaluate how much individual gender can be predicted, linear discriminant analysis with cross-validation was run for the whole set of untransformed head-size variables (Table 2b), log-transformed variables (Table 2c), statistic variables (Table 2d). The overall proportion of correct classification varied from 0.124 (untransformed head-size variables) to 0.847 (log transformed head-size variables), with ratio variables presenting an intermediate result (0.427). Males presented the lowest correct classification rate (0.114) with all kinds of variables, where females varied from 0.194 (untransformed variables) to 0.880 (log transformed variables), with ratio variables again presenting an intermediate rate (0.457).

Principal components analysis of capture animals are presented in the Section 4.2.2.1. See Tables 8, 9, and 10 for PCA of untransformed head size variables, log transformed variables, and ratio variables respectively. The pattern of variables found in the principal components after eigenanalysis are discussed in the previous section and not repeated here. However, distribution of

males and females in the plots of the two first principal components are presented in Figure 1d at the end of the present study.

**Table 11** Linear discriminant analysis (Response: sex of capture records). Predictors: all head shape variables. Sample size: 10 animals. Groups: males and females

**1d) Summary of classification with cross-validation**

Group	1	2
1	13	0
2	0	45
Total N	14	45
N Correct	13	45
Proportion	0.734	1.000

N = 59. N (Correct) = 58. Misclassification Rate = 0.044

**1d) Linear discriminant function for group**

	Constant	BL	SW	SL	BR	BLR	OL	BLW	OLW	SLW	SLR
Males	148.04	2.75	1.42	1.44	1.93	1.59	12.81	4.42	11.75	1.13	
Females	120.11	2.85	1.29	-0.50	-0.28	2.51	14.93	4.37	11.37	1.65	

	PAS	ML	LMS	WLS	LRS
Males	0.23	0.13	0.00	-0.11	0.28
Females	0.35	0.14	0.13	-0.34	-0.14

Similar to the pattern observed by wild animals, a large number of variables was kept by the least subset regression for capture records in order to keep a sufficient coefficient of determination ( $r^2$ ). In other words, most of the variables had to be kept in the model in order to account for a reasonable amount of variation. However, it was consistently lower for capture records than for the wild animals (0.63 for the latter and below 0.500 for the former). Twelve standardized head shape variables (all the BL, and SL), six morphological variables (BL, OL, OLW, SL, WSL, and LMS), and seven ratio variables (BL/LT, BL/ET, BL/W, BL/S, BL/L, BL/WL, and BL/LMS) were kept in their respective sets.

The major source of variation between sexes came to be small scaled variables (PW and OL). Therefore, by manipulation of size significantly improved sexual dimorphism perception through the relevance of scaling-related among variables. The least subsets of variables (Table 10) showed a consistent difference between sexes for all length of variables and various shape/ $r^2$  values < 0.001 for WLS, Leastly, Bootstrap, and Piller's Tests of MANOVA). Linear discriminant analysis with cross-validation for the least subsets of environmental (Table 11), bi-morphological (Table 12), and ratio variables (Table 13) showed a slight improvement in the

overall proportion of correct classifications (0.821, 0.838, and 0.845, respectively). Females still perform a higher proportion of correct classifications (0.844, 0.852, and 0.863, respectively) than males (0.792 for untransformed and log transformed variables, and 0.754 for ratio variables).

**Table 12** Linear discriminant analysis: Response rate of capture animals. Predictors: all log transformed head size variables. Sample size: 98 animals. Groups: males and females

<u>Summary of classification with given relations</u>								
Group	1	2						
1	11	9						
2	4	48						
Total N	21	57						
N Correct	15	48						
Proportion	0.714	0.842						
N = 98 N Correct = 63 Proportion Correct = 0.643								
<u>Log Linear discriminant function for group</u>								
	Constant	FL1	FL2	FL3	FL4	FL5	FL6	FL7
Males	1224.9	1489.5	434.5	1842.4	-1184.4	187.8	121.1	144.8
Females	7139.4	1084.2	446.9	3823.2	-1131.9	347.7	452.1	16.1
	FL1	FL2	FL3	FL4	FL5	FL6	FL7	FL8
Males	417.1	121.9	1084.8	-627.4	1827.6	-108.1		
Females	772.1	-182.5	1853.8	-627.4	-1112.5	128.1		

**Table 13** Linear discriminant analysis: Response rate of capture animals. Predictors: all ratio variables. Sample size: 98 animals. Groups: males and females

Summary of classification with given relations								
Group	1	2						
1	11	9						
2	4	48						
Total N	21	57						
N Correct	15	48						
Proportion	0.714	0.842						
N = 98 N Correct = 63 Proportion Correct = 0.643								
Log Linear discriminant function for group								
	Constant	FL1	FL2	FL3	FL4	FL5	FL6	FL7
Males	2466.1	1817.4	1944.1	1129.7	4498.7	564.7	487.4	1746.8
Females	2940.9	-489.4	2134.7	1129.8	3542.1	772.1	484.4	-1111.5
	FL1	FL2	FL3	FL4	FL5	FL6	FL7	FL8
Males	1191.2	104.1	111.2					
Females	852.1	121.9	111.2					



Table 30. First subspace of variables for sexual development in euglycic groups

Variable	First subspace	P-value	r	n
head-ape	CW 1.047 0.01, CW 0.046 1.00, LCE 0.007 1.00, PEI 0.001 1.00, ML 0.000 1.00, WBL 0.000 1.00	0.000	-0.123	95
Log10-head*	SW 0.01 0.99, CW 0.00 1.00, WBL 0.00 1.00	0.000	-0.035	95
Body	MLT 0.0013 1.00, SW 0.00 1.00, ALAL 0.000 1.00, BLAL 0.000 1.00	0.000	-0.174	95

\* Log10-transformed head-size variables

Multiple procedure: Test → Regression → First Subspace

Principal components analysis for the first subspace of variables (Table 30) and description (Figure 14) included to visualize how variables and females are clustered apart and also how much other variation of variation may still remain in these studies. PC1 is responsible for most of the variation for both the untransformed and the log-transformed head-size variables (BLAL and BLML, respectively). PC1 coefficients for untransformed variables (see Table 30a) present similar magnitudes (from 0.259 for CW to 0.255 for CW and SW) and negative sign. PC1 coefficients for log-transformed variables (see Table 30b) are more variable (from 0.284 for CW to -0.474 for SW and ML), although they also present the same sign. A possible biological reason for these patterns is because overall growth rate in the larger differentials between males and females. In addition, log transformation may stress these morphological differences. However, these two possibilities are ultimately classified by the delineation of sexuals against the values for the two principal components (Figure 14).

Table 31. Linear discriminant analysis (discriminate use of euglycic sexuals). First subspace of head-size variables. Sample size: 95 individuals. Groups: males and females.

a) Frequency of classification in each group: individuals

Group	1	2
1	16	11
2	5	85
Total N	21	77
% Correct	16	85
Proportion	0.762	0.884

N1 = 16, N2 Correct = 16, Proportion Correct = 0.887

b) Linear discriminant functions for groups

	Constant	CW	SW	PEI	WBL	WBL	LCE	ML	BLAL
Males	81.581	-0.004	-1.152	-1.008	0.042	-0.760	-0.783	0.754	-0.015
Females	88.511	-0.004	-0.541	-0.025	0.013	-0.083	0.418	-0.002	-0.022

	ML	WBL	WBL	LCE
Males	-0.754	-0.783	0.754	-0.002
Females	-0.002	-0.002	0.418	-0.785

PC1 accounts for 8.8% and 6.6% respectively for untransformed and log-transformed first subset of variables. A contrast between DFW (4.8%) and the association between OL (8.14%) and RWR (9.25%) is also on PC1 of transformed variables (see PC1 in Table 26a). The contrast is represented by smaller magnitudes and opposite signs. This is clearly visible in the pattern expressed in Table 8, which is composed after most of the variables were kept as the first subset of transformed variables. The biological meaning of this pattern is discussed in Section 4.2.2.1 and seems to represent a relative “disagreement” on the eye-calls induced by ontogenetic development, measured by sexual dimorphism.

**Table 26** Linear discriminant analysis. Response: sex of experimentalists. Predictors: first subset of log-transformed head and eye variables. Sample size: 50 animals. Groups: males and females.

**a) Summary of classification with group and classes**

Group	1	2
1	18	6
2	3	70
Total N	24	73
N Correct	66	70
Accuracy	1.982	1.959

N = 98. MCCohen = 0.7. Agreement Cohen = 0.988

**b) Linear discriminant function for group**

	Constant	LR	LR	LR	MR	WR	LR
Males	195.48	-592.45	140.87	568.81	2.6732	21.937	-465.39
Females	362.14	-978.34	386.23	679.27	4.6838	38.177	2.9789

**Table 27** Linear discriminant analysis. Response: sex of capture scientist. Predictors: first subset of color variables. Sample size: 50 animals. Groups: males and females.

**a) Summary of classification with group and classes**

Group	1	2
1	12	5
2	6	88
Total N	21	93
N Correct	18	88
Accuracy	0.764	0.945

N = 98. MCCohen = 0.7. Agreement Cohen = 0.967

**b) Linear discriminant function for group**

	Constant	LR1	LR2	LR3	LR4	LR5	LR6	LR7
Males	-1427.1	2624.7	621.3	884.7	8.613	1047.1	897.8	225.4
Females	-1677.1	7824.8	448.6	384.8	275.2	1114.2	802.5	182.8



relative elongation to the orbit of the spot, relative reduction in the localized region of the smoothie, and relative widening of the crown infundibulum. This seems to be related to integrated development as well.



(a) All head-size variables



(b) Best subset of head-size variables



(c) All leg-structure of head-size variables



(d) Best subset of leg-structure of head-size variables



(e) All body variables



(f) Best subset of body variables

Figure 14. Small displays of eigenscatter. Plots of eigenscatter (x-axis) versus values for the two principal components (y-axis) (loading).

PCA of body variables provided, as expected, a complex result (see Table 16a). PC1 accounts for only approximately half of the total variance, while PC2 accounts for approximately one fourth of it. Continuous between relative length (RL), SV and width (RW) of head, and between relative orbital width (ROW) and inter-orbital distance (IOW) are classified by PC1

an allometric factor. They are presented both as an allometric  $\alpha$  (in all length variables) they are represented as several constant size changes, instead of just a local management of a specific aspect of the head. BOW-1 (0-11) accounts for most of the variation of PC1. When its coefficient has a negative sign, that may represent a relative reduction of this variable, which means again an “elongation” of the skull, relative described above.

Figure 14 shows the plots of capture animals against the values for the two principal components for the whole set and best subsets of environmental head size variables (Figure 14a and b), log-transformed head-size variables (Figure 14c and d), and size variables (Figure 14e and f). Although better representation of males and females can be seen in all plots, it seems to be more close with log-transformed variables than with untransformed variables, which is compatible with the results of the multivariate analyses and the best subset regressions.

However, a significant number of apparently trapped animals is older than most all plots. Comparing Figure 14 with Figure 7, it is possible to see that there is a considerable influence of age in the distribution of the animals in the plots of the two first principal components. This age effect is increased by the unbalanced distribution of one-year old and three-bladed distributions of two-year-old animals (see Appendix A, B and A IV). Additionally, age effect causes to completely related from body-size effect (see Table 21).

Tables 1 and 21 shows that there is a consistent sexual dimorphism in BAC, lengths and female presents fastest growth rates (see Tables 4 – 7 and Figures 5 and 6) and allometric relations (see Tables 19 – 20 and Figures 11 and 12) due a considerable number of variables. It is unlikely that this happens only for reptiles animals. However, in order to evaluate sexual dimorphism it is necessary to include as much as possible other sources of morphometric variation. In this study, age, body size, climatic, morphological, biogeographic, and environmental are specifically analyzed. Other sources of variation like individual and sensory status of the individuals, geographic region of the population and others, may evaluated even more the study of secondary sexual characteristics of crocodilians. Apparently, in study in the work all these factors is considered. Thus, that results may be representing many different things as “sexual dimorphism”.

Since it is practically impossible to keep “all other things being equal” in field studies experimental design using capture animals may help to solve these problems. Since sex is considered a temperature determined, and there is a considerable climatic effect, an experiment with eggs and hatchlings seems to be the best way to study sexual dimorphism. These animals from study like this, should be equally split into female and male temperatures of

considered. With the procedure, age, size, and clutch effects would be taken off the system. In addition, some treatments (for instance, involving herbicides, as suggested by Gendron et al. 1994) could be added.

An interesting comparison could also be helpful in situations such as the one described by Gendron et al. (1994). Chemical pollution was apparently affecting the primary sexual characteristics of *Aligator mississippiensis* in Lake Apogon, Florida, with definite size trends for the reproductive fish of the whole population. In that case, the analysis of secondary sexual characteristics of the individuals collected may help to understand how the pollution impacts the population. It may also help to identify the occurrence of similar problems in other areas.

### 4.2 Parent-offspring Morphometric Relations

Parent-offspring morphometric relations are considered as the present study to be the morphological manifestation between parents and offspring that can be quantified by the use of multivariate correspondence analysis. The main objective of the present study is to investigate intensive relations between parents and offspring that permit the identification of the forms by the heritable characteristics of the larva.

The first step within the study of the allometry of reproduction, when female body weight and length are related with various clutch characteristics such as number of eggs, absolute and relative mass, and egg and hatchlings formation characteristics (Rosenfeld 1981). This could be called "allometry of parental investment" (Kruuk 1994, however, the concept of parental investment is rather subjective and complex from Triversen 1972). The second step will be the investigation of the morphometric relations between parents and offspring, in such a way that the same variables presented in Figure 1 are used. For this analysis, female families and males will be considered groups to which eggs and hatchlings are related (Figure 4.3.2).

#### 4.2.1 Allometry of Reproduction

The studies of the allometry of reproduction have been correlated with fecundity data. Porgueta (1988) presents an extensive review of the reproductive biology and embryology of ornamental fish, with information on size at sexual maturity, clutch size, egg dimensions, and formation of hatchlings for many species. Porgueta's review presents a great amount of information about species-specific variation in reproductive biology. However, little information is

gross above-ground-specific variation in volume, which he depicted as functions he had no reason to suspect were "stable for most of the species, related/related to single associated aspects."

Hall (1957) used morphometric data from among *A. maculipennis* and *C. maculipennis* and associated shape characteristics to produce regression models where demographic profiles of the reproductive females could be reliably estimated on the basis of clathrate-mass characteristics. He found consistent allometric relations between female Morphometrics (BM, SVL, TFL, and tail girth) and clathrate mass and size ( $r^2$  varying from 0.140 to 0.788 and  $P < 0.05$  for all equations).

In his elegant study, Hall proposed the study of the allometry of reproduction at the population level, in order to determine which segments or sub-segments of females are actively breeding in the population. This remains to be the best approach for large populations, such as the *A. maculipennis* and *C. maculipennis*. However, small populations may (and in some circumstances have to) be approached at the individual level. In small fragments of habitat, the "whole segment of reproductive females" may be just one or a few animals. In this case, allometric relations, such as the ones described by Hall may possibly be suitable already reproductive sub-classes.

Theijunianen (1994) reviews the reproductive characteristics of *maculipennis* and presents allometric data on female size, egg mass, clathrate-mass, and volume, and morphometric allometric results. He reports that egg size, clathrate mass, and clathrate mass are positively correlated with female body size at the interspecific level. According to his interspecific (among populations) comparisons, some limited by sample size, but among six species for which data were available, a positive correlation between female size and egg mass was found in all species except *Cnemidophorus*. Theijunianen also reports a significantly positive relationship between clathrate mass and female length, but none again *C. maculipennis* was an exception. Unfortunately, Theijunianen does not use his information sources concerning the broad-mass volume, which could be suitable already some interspecifics.

The present study female-mass-tail length (TFL) and body mass (BM) are related with the follow egg characteristics: egg mass, length and width, hatching BM and SVL, and clathrate mass (number of eggs) mass, and relative body length mass. Female BM relative clathrate mass was determined by the ratio between clathrate mass and the female BM before reproduction. Data were collected in the captive breeding program of the University of São Paulo, Botucatu, Brazil, during the reproductive period of 1998 + 1999, including samples of two males from further period. Females were measured in early October. At this period of the year

females rarely store a clutch in multiple developmental of follicles (Yin et al., 1992). Hatching occurs in Minke fish from November to February, with peak in January (Verhulst 1992). Linear regression equations are presented in Table 10. Their related plots are presented in Figure 12.

Contrary to the pattern shown for the Thorp parameter for SSC, in the present study a positive correlation was found between egg mass and female SVL ( $r^2 = 0.499$  and  $P$ -value  $< 0.001$ ). A positive correlation, although not significant, was found between clutch mass and female SVL ( $r^2 = 0.461$  and  $P$ -value  $= 0.442$ ). Female SVL showed also significant positive correlations with egg length ( $r^2 = 0.241$  and  $P$ -value  $= 0.006$ ), egg width ( $r^2 = 0.179$  and  $P$ -value  $= 0.005$ ), hatching BM ( $r^2 = 0.423$  and  $P$ -value  $= 0.002$ ) and hatching SVL ( $r^2 = 0.407$  and  $P$ -value  $= 0.004$ ). No clear correlation was found between female SVL and clutch size ( $r^2 = 0.411$  and  $P$ -value  $= 0.197$ ). A negative correlation, although not highly significant, was found between female SVL and the relative clutch mass ( $r^2 = 0.194$  and  $P$ -value  $= 0.166$ ).

Female BM showed a significantly positive correlation with egg mass ( $r^2 = 0.493$  and  $P$ -value  $= 0.000$ ), egg length ( $r^2 = 0.113$  and  $P$ -value  $= 0.000$ ), egg width ( $r^2 = 0.101$  and  $P$ -value  $= 0.000$ ), hatching BM ( $r^2 = 0.609$  and  $P$ -value  $= 0.000$ ), hatching SVL ( $r^2 = 0.579$  and  $P$ -value  $= 0.000$ ), and clutch mass ( $r^2 = 0.413$  and  $P$ -value  $= 0.002$ ). No clear correlation was found between female BM and clutch size ( $r^2 = 0.124$  and  $P$ -value  $= 0.129$ ). A highly significant negative correlation was found between female BM and the relative clutch mass ( $r^2 = 0.781$  and  $P$ -value  $= 0.000$ ).

With the exception of clutch size, there seems to be a clear positive allometric relation between female body-size (or its form correlated with BM or SVL) and all measured clutch characteristics. Even though in the hatching body size, in other words, larger females tend to produce heavier (instead of more numerous) offspring. However, this seems to cost relatively less as females (even for a 20 kg female, the clutch represents approximately 1.2% of her body mass, whereas for a 40 kg female, the clutch represents only approximately 0.6%. Not surprisingly the (negative) correlation between female BM and clutch mass was the most significant allometric relation found in this study ( $r^2 = 0.733$ ).

From the physical constraints suffered by the eggs during incubation to the more weaker constraints faced by hatching, some other hatching characteristics have not been a great number of (direct) selective pressures (see Casper 1980 and Galat et al., 1990, for the effect of habitat on the survival of eggs). Although SSC is still an average derived variable of approximately 2.70 mm from the mother to the next generation (18 – 49 eggs, according to Verhulst et al., 1992



approximately the same according to Leontov (1983), there may be a selection pressure strong enough to prevent the stability of smaller clutches.

Table IV. Allometry of reproduction, least square regression.

			$\alpha$	$\beta$	$\gamma$	$\delta$	$\alpha$	$\beta$	$\gamma$	$\delta$
			value							
1	Egg	Female								
	mean (g)	SVL (mm)	33.3333	81.3482	-0.0041	0.0081	0.000	0.000	0.000	1.00
2	Egg	Female								
	"weight" (mm)	SVL (mm)	-64.0000	1.10000	-0.0000	0.0001	0.000	0.000	0.000	1.00
3	Egg	Female								
	"width" (mm)	SVL (mm)	-0.0000	1.44000	-0.0000	0.0001	0.000	0.000	0.000	1.00
4	Hatching	Female								
	SD (g)	SVL (mm)	-1.0000	0.00000	-0.0000	0.0000	-0.000	0.000	0.000	0.00
5	Hatching	Female								
	SVL (mm)	SVL (mm)	-0.0000	0.0000	-0.0000	0.0000	-0.000	0.000	0.000	0.00
6	Clutch	Female								
	size	SVL (mm)	0.0000	0.0000			0.000	0.000	0	
7	Clutch	Female								
	mean (g)	SVL (mm)	0.000	-0.000			0.000	0.000	0	
8	Relative	Female								
	clutch mass	SVL (mm)	0.0000	-0.0000			0.000	0.000	0	
9	Egg	Female								
	mean (g)	SD (g)	0.0000	0.0000	-0.0000	0.0000	0.000	0.000	0.00	0.00
10	Egg	Female								
	"length" (mm)	SD (g)	0.0000	-0.0000	0.0000	0.0000	0.000	0.000	0.00	0.00
11	Egg	Female								
	"width" (mm)	SD (g)	-0.0000	0.0000	-0.0000	0.0000	0.000	0.000	0.00	0.00
12	Hatching	Female								
	SD (g)	SD (g)	0.0000	-0.0000	0.0000		0.000	0.000	0.00	0.00
13	Hatching	Female								
	SVL (mm)	SD (g)	0.0000	-0.0000	0.0000		0.000	0.000	0.00	0.00
14	Clutch	Female								
	size	SD (g)	0.0000	0.0000			0.000	0.000	0	
15	Clutch	Female								
	mean (g)	SD (g)	0.0000	-0.0000			0.000	-0.000	0	
16	Relative	Female								
	clutch mass	SD (g)	0.000	-0.000			0.000	0.000	0	

$$Y = \alpha + \beta X_1 + \gamma X_2 + \delta X_3$$

Relative clutch mass = clutch mass (g)/ female egg mass

Mean (precision) Size = Regression on Regression

Regression = 0 (not used during)

Precision = 0 (not used during)

Hatching body size may be related to survival rate or the first winter survival rate in the immature part of the distribution of the species (Pike et al., 1994). This differential success may benefit larger females. It is tempting to believe that this happens as well established

populations with “healthy” structured age pyramids and a well established hierarchy among reproductive subgroups (Barnes et al. 2003) and in the case of small colonizing populations exhibiting depressed and unbalanced sex ratios, in which sexual processes and behavioral ecology must be completely altered. The possible consequences for the reproduction of the species under such circumstances are unknown.

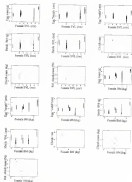


Figure 12: Plots of relationships of reproductive, see Table 14 for regression equations. Reproductive-Chick mass (Tg) = Chick mass (kg), Female body mass (kg).

Although the observed values above are statistically significant, the high variance found (see Figure 1.5) prevents their use in the assessment of individual parents from their characteristics, even in small populations. In the next section, multivariate approaches are attempted in order to accomplish this goal.

#### 4.1.2 Clutch Discrimination

There is a popular conception: a mud-dweller female can sometimes identify the individual laying her by the size and shape of an egg. This is considered here a skill that can be described, quantified, and used for field studies of small populations of mud-dwellers. The term *clutch discrimination* is used in the present study in the process of developing a technique, based on the morphology of eggs and hatchlings. Female mud-dweller hoppers (called here female hatch) and male effect are evaluated as eggs' and hatchlings' morphology. Multivariate statistical analysis (as presented in Figure 2) are used here as tools for the evaluation of the differences among groups, reduction in the number of variables, and discriminatory analysis.

Data were collected in the capture breeding program of the University of São Paulo, Piracicaba, State of São Paulo, Brazil, during the reproductive period of 1991/1992. Reproductive groups were formed by a single male with three to four females, three years before the present study, which makes primary evidence. Maturity was achieved by genital behavior participating (as described by Cox 1971 for *Crossodonta salicaria* Blais 1913 for *C. maculata* Blais 1918 for *Gerris pusillus* Kachler and Kachler 1976 for *Allygus microsphaerus* and Wolfmeier et al. 1968 for *C. dorsalis*). PCN primers provided by Glass et al. in press, for *A. microsphaerus* were used in order to establish paternity among individuals, but preliminary results showed no evidence of paternity discrimination. This was probably due to a high heterozygosity between females.

Six clutches were produced by six mating females (CL1, CL2, CL3, CL11, CL15, and CL19) and three males (CL1, CL36 and CL20). Females CL1, CL2, CL3, and CL15 are clones, while the other two females are mated to the non-related animals. Clutches were produced by the following combinations of parents (male/female): 1) (CL36/CL1), 2) (CL36/CL19), 3) (CL1/CL19), 4) (CL36/CL3), 5) (CL36/CL15), and 6) (CL19/CL1). Clutch effect is studied in this study in the genetic and phenotypic influence of the parents associated with the phenotype, influence of the incubation medium on eggs and hatchlings.

All clutches were collected during the first two days after oviposition. Eggs from all clutches were incubated in waterbaths made of dental enamel jars (as described by Partridge *et al.*, 1983) at the same temperature and relative humidity ( $24 \pm 0.5^\circ\text{C}$  and  $60 \pm 10\%$  respectively). The clutch effect on the morphology of the eggs is analysed in section 4.1.1.1. The clutch effect on hatchlings is presented and discussed in section 4.2.2. The female family effect on eggs is presented and discussed in section 4.1.1.2. The female family effect on hatchlings is analysed in section 4.2.3. Males are treated as non-factor: a significant influence on the egg's sex ratio drops. Although their effect on hatchlings is difficult to completely exclude from the female and female family effects, it is focused in section 4.1.2.1. Egg morphology is represented by the following variables: mass ( $\mu\text{g}$ ), length (major diameter,  $\mu\text{m}$ ) and "width" (minor diameter,  $\mu\text{m}$ ). Hatchlings' overall morphology is represented by the morphometric variables described in Figure 1 and Table 1. Log-transformations of eggs' measurements is also presented and discussed.

#### 4.1.1 Clutch effect on eggs

There is a significant clutch-effect on egg morphology ( $F$ -value  $> 0.001$  for Wilk's Lambda-criterion, and Pillai's Trace of MANOVA) for the whole set of untransformed and log transformed egg measurements. Log-transformations did not significantly affect linear discriminant analysis with cross-validation. For both kinds of statistical discriminant analysis based on several properties of correct classification of 97.0% (varying from 9.54 to 97.0% among clutches for the former (Table 4b) and from 0.004 to 9.54% for the later (Table 4c).

PCA of untransformed and log-transformed egg measurements is presented in Table 4d. PC1 of untransformed measurements is responsible for 9.81% of the total variance (Table 41a), while PC1 of log transformed measurements accounts for 9.15% (Table 42a). This probably means that log-transformations rendered a difference in egg overall size between clutches. PC2 of untransformed measurements is responsible for 6.12% of the total variance (second variance), PC2 of log-transformed measurements accounts for only 4.95%. This probably means, as we say later, that log-transformations makes analysis of data less sensitive to changes in egg shape.

Egg mass is the major cause of variation in PC1 for both untransformed and log transformed measurements. Followed by egg length in the former and egg width in the later. This suggests that there is a egg shape variation between clutches in terms of egg mass, which is true ( $F$ -value  $> 0.001$  in simple ANOVA). PC1 shows a clear gradient between length and width, both

logarithm of magnitude (number) and sign (positive) for the two kinds of measurements. The relation between these two linear measurements (egg length and width) explains the variance of the egg shape, which can be longer or wider, but always oval or broadly elliptical. This relation also seems to represent an individual female effect. This means that this relation must be the major source of shape variation in the eggs, which is explained by the constant observed at PC1.

**Table 46** Linear-discriminant analysis (lognormal shape). Predictors: all untransformed measurements (sample size 194 eggs). Group: cluster.

**A) Summary of Canonical variates with group relations**

Group	CL21	CL19	CL8	CL3	CL20	CL11
CL21	29	5	0	1	0	0
CL19	4	26	0	1	0	1
CL8	1	3	22	1	0	0
CL3	1	0	0	24	1	0
CL20	0	0	0	0	20	0
CL11	0	1	0	1	0	24
Total N	39	28	22	24	20	24
N Correct	29	24	20	18	20	24
Proportion	0.743	0.857	0.909	0.750	1.000	0.999

N = 194 N Correct = 158 Proportion Correct = 0.814

**B) Linear discriminant function for group**

Variable	Egg Area	Egg Length	Egg Width
CL21	1499.1	26.1	234.4
CL19	3429.1	18.4	273.9
CL8	-3379.0	-25.8	282.1
CL3	2562.9	-20.3	340.4
CL20	3660.6	24.1	273.9
CL11	-3767.1	37.3	267.7

However, as absolute terms, egg area and egg width measure the major variance of overall variation in the eggs (area/size and shape). Table 41 presents the first subset regression of log-transformed and log-transformed egg measurements. In both cases, egg area and egg width are the chosen variables ( $r^2$  varying from 0.349 to 0.374, respectively, with  $P$ -value = 0.000 for both cases). The difference between absolute  $r^2$  will disappear after the reflection in the number of variables ( $P$ -value = 0.000 for Wilk's Lambda, Hotelling, and Pillai's Trace of MANOVA).

Linear discriminant analysis of the first subset of untransformed (Table 46) and log-transformed measurements (Table 47) determines slight decrease in the overall proportion of

worst classifications (3/707 and 6/70), respectively) in relation to the whole set of measurements (0/700 for both kinds of searching) (Fig. 4) of the combined function (CL13) and not of the single CL1/CL2 kept presenting the highest proportions of correct classifications (above 8/100), similarly to the pattern derived by the discriminant analysis of the whole set of variables. The single combined function (CL13) showed intermediate results, while the other three scores presented the worst results, which is also similar to the pattern formerly presented.

**Table 67** Linear discriminant analysis: Response of both Predators (3) to egg-transformed *Macromphalina* (0/1), more (light) and smaller (dark) *Caecophora* (1/0) eggs. Group: cluster

of Response of the function with cross-validation						
Group	CL1	CL2	CL3	CL4	CL5	CL13
CL1	39	1	1	1	0	0
CL2	4	20	0	2	0	1
CL3	2	0	22	0	0	0
CL4	0	0	0	10	1	0
CL5	0	0	0	0	24	0
CL13	0	1	0	1	1	20
Total N	19	20	20	20	20	20
McNemar	20	20	20	20	20	20
Asymptotic	0.793	0.871	0.799	0.871	0.799	0.793

N = 190, McNemar = 0.0, Pearson Chi-Square = 0.798

#### CL13 new function: function by group

	Constant	Egg (dark)	Egg (light)	Egg (white)
CL1	39.012	16.1	199.7	30.011
CL2	-39.012	60.1	1.0747	60.017
CL3	39.012	20.1	199.7	30.011
CL4	-39.012	60.1	160.1	30.011
CL5	-39.012	10.1	199.6	150.1
CL13	-39.012	60.1	150.1	100.1

Squared distance between groups for both subsets of variables presented results similar to the whole set of variables, but combined function CL13 slightly closer to last scores. Functions CL1, CL2 and CL13 kept closer to each other than to any other, whereas function CL3 is the most statistical separated. *Macromphalina* followed this pattern, with the closer distances showing the highest number of correct classifications of eggs between each other. A possible biological meaning for this pattern is that there exists a genetic component on the egg form. Another possible reason is the difference of age and body size between related and unrelated females. The scores have similar body size and stage eggs, whereas unrelated females are older and bigger.

PCA of the log transformed and log transformed egg measurements (Table 4b) presented approximately equivalent results as shown by the PCA with the whole set of variables. PC1 of the non-transformed measurements accounted for 8.937 of the total variance (Table 4b) with an equal distribution of the variance between the two variables (8.780 for egg mass and egg width). However, PC1 of the log-transformed measurements accounted for a little more variance (9.101) unequally distributed through the two variables (8.543 for egg mass and 8.133 for egg width, as can be seen in Table 4b). PC1 of non-transformed measurements showed an equal balance between variables, whereas PC1 of the log-transformed variables showed egg width as the major source of variance, as exemplified in the pattern of variables captured in the principal component load principal component.

**Table 4b** Check effect on eggs. Principal components analysis of all measurements. Sample size: 194 eggs.

**4a) Eigenvalues of the correlation matrix of all non-transformed measurements**

Eigenvalue	2.3374	0.7024	0.0001
Proportion	0.841	0.249	0.010
Cumulative	0.841	0.890	1.000
Variable	PC1	PC2	PC3
Egg mass	-0.614	0.614	-0.000
Egg length	-0.612	-0.610	0.000
Egg width	-0.789	0.610	0.000

**4b) Eigenvalues of the correlation matrix of all log-transformed measurements**

Eigenvalue	9.703661	0.540288	0.000052
Proportion	0.704	0.032	0.004
Cumulative	0.704	0.736	1.000
Variable	PC1	PC2	PC3
Egg mass	-0.686	0.111	0.000
Egg length	-0.114	-0.804	-0.000
Egg width	-0.559	0.589	-0.781

The biological implications of the results above may be interpreted as follows. Overall egg mass is the major source of variation among clutches. Log-transformations may stress the role of each variable as the standard shape components of eggs from variation among clutches. On the other hand, non-transformed data seem to be more sensitive to the standard size and shape variation than to the identification of individual variables.

Table 43: First subject of correspondence for shape effect on eggs.

Variable	First subject	F value	p	df
Low measurements	egg.mass, egg.width	0.046	0.549	194
Log-transformed measurements	egg.mass, egg.width	0.008	0.934	194

Model procedure: Step 1 → Response1 → Low Masses.

Table 44: Linear discriminant analysis Response: clutch Predictors: first subject of low measurements Sample size: 194 eggs Groups: clutches

#### 4.4 Summary of discriminant analysis with low

Group	CL.05	CL.10	CL.15	CL.20	CL.25	CL.30
CL.05	20	1	1	1	0	0
CL.10	0	14	0	1	0	1
CL.15	1	9	0	4	0	0
CL.20	4	1	1	17	1	0
CL.25	0	0	0	0	26	0
CL.30	0	1	0	1	1	26
Total N	24	26	14	24	28	27
N Correct	11	16	14	14	26	26
Proportion	0.458	0.615	0.985	0.579	0.929	0.963

$N = 194$  N Correct = 136 ProportionCorrect = 0.701

#### 4.4 Linear discriminant function for group

	F value	Egg.mass	Egg.width
CL.05	0.0000	0.0	0.0000
CL.10	0.0000	-0.0	0.0000
CL.15	0.0000	0.0	0.0000
CL.20	0.0000	0.0	0.0000
CL.25	0.0000	-0.0	0.0000
CL.30	0.0000	0.0	0.0000

Petrica (1981) developed an interesting way to describe shape of bird's eggs, based on the statistical expression of their two-dimensional profile represented by a "profile mapping matrix", developed by Janssen. This seems to be useful in comparative interspecific studies. However, in comparative studies such as those taking two or three measurements of eggs seem to be less statistically powerful and precise enough to capture morphological variation in relatively discriminant shapes.

Figure 14 shows the plot of egg aspect values for the two principal components for the whole set and best subset of untransformed measurements (Figure 14a and b, respectively) and the whole set and best subset of log-transformed measurements (Figure 14c and d). It is possible to see that clutches 01 and 02 (and partially 04) are clustered apart from the others. Untransformed females are responsible for clutches 01 and 02, which explains this pattern. Clutch 04 belongs to



one of the effects that seems to be more distinct from the others, as was expressed by the separated dimensions of the discriminant analysis. Clusters CL 23 and 34, however, seemed to belong to the same group. In the next section, female body effect on egg size-dimension. Figure 11 stresses this effect and should be compared with Figure 14.

Table 45: Linear discriminant analysis. Response: cluster. Predictors: log10 values of log-transformed measurements. Sample size: 150 eggs. Groups: clusters

Efficiency of discrimination with new variables						
Group	CL23	CL34	CL35	CL36	CL37	CL38
CL23	20	1	2	1	0	0
CL34	0	18	12	1	0	1
CL35	0	0	18	4	0	0
CL36	4	3	7	15	1	0
CL37	0	0	0	0	10	1
CL38	0	1	0	0	1	18
Total N	24	20	29	24	10	24
N Correct	20	19	18	20	10	18
Proportion	0.792	0.950	0.621	0.833	1.000	0.750
N = 150, N Correct = 125, Proportion Correct = 0.833						

#### 4. Linear discriminant function for group

	Constant	log10 log	log10 width
CL23	3805.1	396.1	621.2
CL34	2962.6	450.4	483.9
CL35	3938.7	399.0	621.4
CL36	3962.9	389.1	621.3
CL37	-1208.6	389.1	619.1
CL38	3891.1	359.1	593.9

The plot on Figure 16 shows that log-transformations slightly decrease the efficiency of cluster discrimination for female-CL23, which is compatible with the results of discriminant analysis. This female's eggs are clearly separated from all the others basically by size. Graphically, this can be visualized by the fact that they can be put apart along PC2-axis only. Eggs from female CL23 (black PI) seemed to present a consistent difference in shape as related to the others, as can be seen by the preponderance of PC2-axis in their separation from the others. A consistent difference in egg shape seem to separate female CL 23 (black PI) from the others.

Considering that morphological seems to present a low genetic variability (Clerici et al., 1977; Adams et al., 1988; Loefer et al., 1987), the results shown are highly surprising and promising. Even in a small group of related females, there seems to be enough morphological

variability in discriminant abilities with significant success. It is interesting to compare these results with the ones presented in the next section (4.3.2.2 – 4.3.2.3).

**Table 16: Classification results: Principal components analysis of first subset of measurements, Sample size: 198 eggs**

a) Classification of the categories results of first subset of the measurements.

Category	PC1	PC2
Positive	0.927	0.043
Consistent	0.927	1.000
Variable	PC1	PC2
Egg-M	0.797	-0.797
Egg-S	0.797	-0.797

b) Classification of the categories results of first subset of measurements of measurements

Category	PC1	PC2
Positive	0.927	0.043
Consistent	0.927	1.000
Variable	PC1	PC2
Egg-M	0.797	-0.797
Egg-S	0.797	-0.797



(a) All the measurements



(b) First subset of the measurements



(c) All log transformed measurements



(d) First subset of log transformed measurements

**Figure 18: Effect of log on the eggs.** Pairs of egg-space values for the two principal components: Cluster (ambush-h): #1 (CL1704/CL175), #2 (CL1004/CL175), #3 (CL1434/CL175), #4 (CL1404/CL175), #5 (CL1004/CL175), and #6 (CL1404/CL175).

### 4.3.3.1 Female family effect on eggs

As described above, four of the six reproductive females are sisters, while the other two are non-related individuals. Therefore, the following female families are recognized in the present study: Family #1 (CL25), Family #2 (CL3, CL6, CL8, and CL10), and Family #3 (CL17). These maternal lineages, also called here female families, are considered (any  $\alpha$  groups, and their influence on eggs morphology is analysed according to the methodology presented in Figure 7 and described early in the text (see Section 3.1.1).

A significant difference between female families in terms of eggs morphology was found ( $P$ -value = 0.000 for Wilcoxon, Levene Homosked. and Fisher's Tests of MANOVA) for the whole set of measurements including morphological measurements. Linear discriminant analysis with cross-validation showed significant higher proportions of correct classifications for both untransformed (Table 47) and log-transformed (Table 48) measurements (0.944 and 0.988 respectively) when compared to the chance effect described in Section 4.3.2.1.

Table 47 Linear discriminant analysis: Response: female family effect on eggs. Prediction of new measurements. Sample size: 174 eggs. Groups: female families

a) Summary of classification with cross-validation			
Group	1	2	3
1	24	6	0
2	0	103	0
3	0	1	49
Total N	24	109	49
N Correct	24	103	49
Proportions	1.000	0.940	1.000
N = 178    N Correct = 177    Proportion Correct = 0.944			

b) Linear discriminant function by group				
	Constant	Age (years)	Age (years) <sup>2</sup>	Age (years) <sup>3</sup>
1	-0.0227	24.3	1.044	0.0414
2	-0.0211	21.9	1.027	0.0404
3	-0.0211	24.3	2.067	0.0214

Consistently to the pattern suggested by cluster effect analysis, families #1 and #2 can be considered closer to each other than family #3. Females (CL25 and CL33) (Families #1 and #2 respectively) seem to be the most maternal individuals, while the group of sisters (Family #2) seem to be the least maternal group. The Bayesian Goodness of fitG as found is relatively minor (see Bartolde and Santiago 1991). Consequently, there is an reliable information that

results indicate that females lay greater eggs. It is revealed in the present study that females CL23 or CL50 are not related to the females of family 62. They (CL23 and CL50) do not necessarily related to each other as well.

PCAs of the whole set of untransformed and log-transformed egg measurements were presented and discussed in the former section (Figure 4). Table 4F shows the best subset regression for the female-family effect on eggs. These results are consistent with the ones showed by the best subset of measurements for clutch effect analysis (see Table 1E), with especially higher coefficients of determination ( $R^2$ ).

**Table 4E. Least discriminant analysis. Response: female-family effect on eggs. Predictor: all log-transformed measurements. Sample size: 184 eggs (groups: female-families).**

<u>4E. Summary of discriminant analysis with group effect (eggs)</u>			
Group	1	2	3
1	26	4	4
2	4	114	4
3	0	1	40
Total N	34	129	48
McNemar	31	114	40
Proportion	0.353	0.907	1.000
N = 184    McNemar = 184    Response Correct = 61/68			
<u>4E. Least discriminant function for group</u>			
	Constant	$X_{egg}$ mean	$X_{egg}$ length
1	-2541.9	26.79	2711.9
2	2823.1	144.4	2780.0
3	7468.6	231.1	2811.2

After excluding the number of females, there is still a significant difference between female-families in relation to egg morphology, as could be expected ( $P$ -value < 0.000 for Wilk's Lambda Hotelling and Pillai's Trace of MANOVA for the best subsets of measurements). Least discriminant analysis with more variables for the best subsets of measurements (Table 4F) and log-transformed measurements (Table 5F) shows the slight difference in the proportions of correct classifications (9.40% in the former and 11.4% in the later). Discriminant analysis of the best subsets also showed similar patterns of distances between female-families to the ones showed by the whole set of measurements.

Figure 1F shows the plots of eggs' apical values for the two principal components for the whole set and best subsets of untransformed and log-transformed measurements. Comparing the

with Figure 14a the lower variance of the stronger family effect is relative to the effect of egg morphology. One again, it is possible to see that the eggs from female CL10 (i.e., Family CL1) are clustered apart from the other females by size difference (i.e., a difference explained mostly by the first principal component) whereas the eggs from female CL11 (of only 4) presents size and shape differences in relation to the others (i.e., their variance is spread over PC1 and PC2 axes). Here it is possible to see that Family CL1 is at an intermediate position between the unrelated females. Log transformations did not significantly improve data analysis. Reducing the number of measurements, as it has, did not cause any significant loss of efficiency on the analysis as well.

Table 49: Best values of measurements for family family effect on eggs

Variable	Best values	Factor	$r^2$	N
Size measurement	egg length, egg width	0.000	0.000	100
Log transformed measurements	egg length, egg width	0.000	0.000	100

Model procedure: Best  $\rightarrow$  Regression  $\rightarrow$  Best Subset

Table 50: Linear discriminant analysis. Response: family family effect on eggs. Particular best value of untransformed measurements. Sample size: 100 eggs. Groups: female-female

#### A) Summary of classification with your variables

Group	1	2	3
1	34	7	0
2	4	102	0
3	0	1	40
Total N	38	109	40
N Group	38	110	40

Proportion: 0.895 0.918 0.900

N = 178. All Correct = 160. Proportion Correct = 0.900

#### B) Linear discriminant function for group

	Constant	Factor1	egg width
1	-11.000	0.0	0.111
2	0.000	-0.0	0.011
3	-11.000	0.0	0.011

Differences between family families may possibly express a genetic factor related to maternal lineage. In this case, it is unclear what a small group of family related female family could represent if females originate from different families may possibly be determined with high efficiency. Alternatively, it change in the pattern of egg morphology could be related to spatial or displacement of maternal reproductive females. In these cases,

cluster decomposition based on female and female-family effect on the composition of the egg sample for a holistic understanding on landscape genetic structure and meta-population dynamics.

**Table 11** Linear discriminant analysis. Response: female-family effect on eggs. Predictor: first subset of log-transformed measurements (egg, chorion and width). Sample size: 198 eggs. Group: female families.

a) Summary of classification with cross validation			
Group	1	2	3
1	34	2	6
2	4	140	6
3	6	2	48
Total (n)	44	150	60
% Correct	77	93	80
Proportions	0.220	0.750	0.330

$N = 198$      $N$  Correct = 141    Proportion Correct = 0.714

b) Linear discriminant function for eggs			
	Constant	Egg width	Egg vol.
1	2004.2	1240.9	1.954.1
2	2048.6	1154.4	2094.1
3	2085.4	1223.2	2162.9



a) 1st two measurements



b) First subset of raw measurements



c) 1st two log-transformed measurements



d) First subset of log-transformed measurements

**Figure 17** Female-family effect on eggs. Physical egg aspect values for the two principal components. Number: #1 (CL24) #2 (CL5 - CL8 - CL10 - CL13) #3 (CL66)

### 4.3.3.2 Check effect on hatchlings

There was significant check effect on the sexual morphology of hatchlings ( $P$  values < 0.002 for SVL,  $\pm$  Leaning, Headling, and Pelvic + Tails of MANDVA) for the whole set of independent and log transformed head size variables and ratio variables. The correlation between hatchlings' body length and check is presented by head size and ratio variables in Table 10. Significant difference in ANCOVA means that checkers present significant differences in the allometric relation between SVL and the morphometric variable  $x$ -variable. The biological meaning for this is a significant difference in hatchlings' sexual shape between checkers.

Sex of dependent head size variables (CW, CL, CW, KPW, ML, and WBL) and size of dependent ratio variables (RLW, RCL, RAN) and BWL) showed a significant difference between checkers as covariates of SVL. This difference seems to be equally distributed between pregnant and nonpregnant, which possibly means that the whole composition of the skull varies among checkers. However, there seems to be a predominance of ratio variables (CW, CPW, KPW, RCL, RLW, RAN, and BWL) over length variables (CL, ML, and WBL). This means that hatchlings from different checkers, presenting similar body size, differ very much in terms of their relative skull width than that of their relative skull length.

Table 11. Analysis of variance (check as covariate) of SVL ( $P$  values).

Variable	SS	Check	Variable	SS	Check
SVL	0.000	0.174	CL	0.154	0.001
CW	0.000	0.000	CPW	0.011	0.001
ML	0.004	0.171	KPW	0.000	0.174
CPW	0.001	0.110	RCL	0.009	0.001
RCL	0.000	0.000	RLW	0.000	0.001
CW	0.000	0.000	RAN	0.001	0.001
KPW	0.000	0.000	BWL	0.001	0.001
RLW	0.000	0.000	WBL	0.000	0.001
WBL	0.000	0.000	ML	0.000	0.001
ML	0.000	0.001	ML:Lean	0.070	0.001
LMR	0.000	0.000			
WLR	0.001	0.001			
LM	0.000	0.174			

Statistical procedure: Stat  $\rightarrow$  ANCOVA  $\rightarrow$  General Linear Model.

Response (dependent variable): morphometric variable.

Model (independent variable): SVL.

Covariates: Check for response hatchlings.

that all children experienced under normal conditions during routine period: the non-chick infection threat is assumed to come direct from their parents. Although, parents' own moral and sanitary status is unlikely a few –day ordinary influence the overall morphology of hatchlings, a substantial portion of this variation may be associated to genetic in the wild, however, a consistent reduction in nesting care and inappropriate conditions is likely to happen, probably increasing the variance into noise.

Log-transformation did not affect the results of the linear discriminant analysis with seven variables (see Tables 11 and 14). The whole set of both untransformed and log-transformed larval size variables provided the same overall proportion of correct classification (84.8%), varying from 8.75% (Clutch #3, Female CL10) to 1.00% (Clutch #1, Female CL12). These results are even higher than the ones obtained by the whole set of both untransformed and log-transformed egg measurements (8.76%). Seven variables provided a slightly lower overall proportion of correct classification (84.1%) – varying from 0.76% to 0.80% respectively to the most cluster change (Table 11).

Spatial distances between groups based on hatchlings' overall morphology showed a slightly different pattern than the one derived by eggs' morphology. Clutch #1 (Female CL25) was the most distinct while Clutch #2 (Female CL10) was the closest to the others for all kinds of variables. Log-transformation of data resulted in the approximation of Clutch #5 (Female CL09) to the others, which did not happen with untransformed variables. Misclassifications predominantly occurred clusters from related families, similarly to what happened with eggs.

PCA of the whole set of untransformed (Table 9a) and log-transformed larval size variables (Table 10) showed similar results. PC1 represented by less than half of the total variance (37.4% for the former and 34.9% for the latter), whereas PC2 accounted for 14% (38.0%) for the former and 22.4% (39.4%) for the latter. The remaining principal components were responsible for 20.7% (39.9%) and 28.7% (38.6%) respectively for untransformed and log-transformed variables. This means that most of the variation among clusters is expressed by differences in larval shape, not size.

PC1 coefficients varied from 41.614 (JOW) to 0.363 (ML) for untransformed variables, and from 0.049 (JOS) to 0.192 (JOW) for log-transformed variables. This difference of the weight of JOW variable seems to be probably due to differences in scaling caused by log-transformation of data. In both cases, PC1 coefficients are much smaller than the ones found in egg (Percent 4.2.3) and overall morphology (4.2.4) studies. In these cases, overall differences in size were the single source of variation, which is not the case here. PC1 coefficients varied from





**Table 14: Linear discriminant analysis: Regression effects-effect on landings. Predictor all log-transformed larval size variables. Sample size: 174 landings. Group: otoliths**

**a) Summary of coefficients with zero values**

Group	CL12	CL15	CL18	CL1	CL10	CL11
CL12	24	0	0	0	0	0
CL15	0	14	0	0	0	1
CL18	0	3	14	1	0	0
CL1	1	0	0	10	0	0
CL10	0	0	0	0	20	0
CL11	0	3	0	0	0	0
Total N	31	24	20	14	20	11
N/Cases	00	00	24	10	20	0
Proportion	0.041	0.078	1.000	0.000	1.000	0.000

N = 136 N Cases = 136 Proportion Cases = 0.000

**b) Linear discriminant function for group**

	Constant	CL12	CL15	CL18	CL1	CL10	CL11	CL1	CL10	CL11	R <sup>2</sup>
CL12	-1047	1026	210	-1068	1011	000	144	1017	000		
CL15	-1064	1034	-01	-1034	-040	1030	-11	100	1000	-100	
CL18	10001	10470	-170	-1162	1004	1004	0	-100	1000	100	
CL1	-10020	10370	100	-1004	1111	1110	-10	-100	1070	-100	
CL10	-10001	10000	00	-1010	1010	1010	0	-100	1000	-100	
CL11	-10100	10100	-100	1010	1010	1010	0	-100	1010	-100	

	R <sup>2</sup>	ML	LM	WRT	LM
CL12	0.00	11000	1100	100	-1000
CL15	-0.00	11000	-1100	100	-1100
CL18	-0.00	10000	1000	100	-1000
CL1	0.00	11000	1100	100	-1100
CL10	-0.00	11000	-1100	100	-1100
CL11	-0.00	11000	1100	100	-1100

PCA of size variables presented a stronger pattern of results (Table 15). PC1 was responsible for 8.63% of the total variance, respectively higher than PC1 of size ratio variables, whereas PC2 accounted for 6.83%, and the remaining principal components represented the 0.21% of the total variance. EWST (4.74%) was the single source of variation in PC1, whereas BOW (4.93%) accounted for most of the variation expressed by PC2. There are no clear complex patterns considering that neither EWST nor BOW are non-dependent variables in landings (see Table 11). They may be related, however, to the tendency of variation between landings from different otoliths, expressed above, at which relative width of larval core is related to the relative length.

Table 14. Lower discriminant analysis: Response: clutch effect on hatchlings. Predictor: all main variables. Sample size: 129 hatchlings. Groups: absolute.

All group of Colubiferina with non-variables						
Group	CL10	CL14	CL16	CL17	CL18	CL19
CL10	34	0	0	0	0	0
CL14	0	17	0	1	0	1
CL16	0	0	11	0	1	0
CL17	1	1	0	12	0	0
CL18	0	1	2	0	10	0
CL19	0	1	1	1	0	1
Total 1st	35	19	13	14	11	11
N Cases	10	17	11	13	10	11
Percent	8.54	8.78	8.78	8.54	8.54	8.78

N = 124. N Cases = 62. Percentages sum = 100.

b) Lower discriminant analysis by group

	Group	W	CL10	W	CL14	W	CL16	W	CL17	W	CL18	W	CL19
CL10	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7
CL14	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7
CL16	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7
CL17	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7
CL18	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7
CL19	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7	136.2	-4302.7

Table 15 shows the best subset of variables for the study of clutch effect. Log-transformation did not result in a "best subset" of different variables. Log-transformation did not significantly increase the coefficient of determination ( $R^2$ ) of the model either. Since variables provided the lowest  $r^2$  (0.44) against approximately 8.78% of non-variant variables. The reduction in the number of variables did not affect the significance of the difference between CL10 class, in terms of sexual asymmetry of hatchlings (P-values > 0.00 for Wilk's, Likelihood, and Fisher's tests of MANOVA) for all kinds of variables.

**Table 16** Principal components analysis. Eigenanalysis of the correlation matrix of all load use variables. Sample size: 126 hatchlings. Only six first principal components are presented

Eigenvalue	0.2511	0.2406	0.1746	0.0475	0.0398	0.0300
Proportion	0.483	0.448	0.323	0.090	0.074	0.056
Cumulative	0.483	0.931	0.999	0.999	0.999	0.999
Variable	PC1	PC2	PC3	PC4	PC5	PC6
SL	-0.156	0.156	-0.047	0.117	0.064	-0.176
CV	-0.146	0.163	-0.131	-0.086	-0.137	0.168
SL	0.112	-0.146	-0.134	-0.114	-0.039	0.060
SW	-0.138	0.176	0.044	-0.054	-0.031	0.030
OL	-0.171	-0.154	-0.066	-0.016	-0.076	-0.066
CPW	-0.060	-0.161	-0.136	0.077	0.132	-0.156
SWP	-0.160	-0.145	-0.135	-0.146	-0.028	0.166
WFB	-0.020	-0.153	-0.130	0.117	0.064	0.164
WB	-0.080	0.064	0.160	0.116	-0.027	-0.106
PCD	-0.120	0.144	-0.080	-0.112	0.060	0.107
MS	-0.060	0.126	0.066	-0.046	0.075	-0.066
LSW	-0.160	0.164	0.160	0.147	0.080	0.166
WSD	-0.177	0.136	-0.120	-0.112	-0.106	0.166
LSR	-0.120	-0.166	0.166	0.136	0.067	-0.136

**Table 17** Principal components analysis. Eigenanalysis of the covariance matrix of all log-transformed head-use variables. Sample size: 124 hatchlings. Only six first principal components are presented

Eigenvalue	0.164460	0.160717	0.091134	0.059976	0.050497	0.007737
Proportion	0.413	0.334	0.224	0.148	0.124	0.020
Cumulative	0.413	0.747	0.971	0.991	0.997	0.999
Variable	PC1	PC2	PC3	PC4	PC5	PC6
SL	-0.160	-0.170	-0.043	-0.036	-0.067	-0.074
CV	-0.164	0.166	-0.076	0.136	0.064	-0.166
SL	-0.160	-0.070	0.137	0.271	-0.136	-0.161
SW	-0.166	-0.136	-0.046	-0.081	-0.064	-0.057
OL	-0.177	-0.067	0.136	-0.076	0.087	0.126
CPW	-0.166	-0.080	0.160	-0.137	-0.012	0.134
SWP	-0.160	0.166	-0.070	0.136	-0.067	-0.066
LSR	-0.164	-0.060	0.166	-0.020	-0.160	-0.166
WB	-0.170	-0.166	-0.076	-0.176	0.136	0.060
PCD	-0.166	-0.166	-0.010	0.070	0.136	0.166
MS	-0.164	-0.117	-0.076	0.076	-0.112	0.066
LSW	-0.136	-0.166	-0.066	-0.160	-0.066	0.177
WSD	-0.162	0.166	-0.076	0.166	-0.060	-0.134
LSR	-0.136	-0.111	-0.066	-0.020	-0.121	0.160

Linear discriminant analysis with cross-validation of the untransformed (Table 14) and log-transformed head-use variables (Table 15), and size variables (Table 12) showed a slight decrease in the overall proportion of correct classifications (approximately 0.600 for the two



**Table 48** Linear discriminant analysis. Response: clutch effort on hatchlings. Predictors: log values of head size variables. Sample size: 124 animals. Groups: clutches

*a)* Summary of discriminant analysis with cross-validation

Group	CL20	CL21	CL2	CL3	CL10	CL11
CL20	25	1	0	1	0	1
CL21	0	10	3	1	0	1
CL2	1	3	22	3	0	0
CL3	0	2	4	7	0	1
CL10	0	0	0	0	21	0
CL11	0	1	0	1	0	0
Total N	27	16	29	14	21	10
N Correct	25	10	22	7	21	0
Proportion	0.926	0.625	0.760	0.500	1.000	0.000

N = 124    N Correct = 90    Proportion Correct = 0.726

*b)* Linear discriminant functions for group

	Constant	CL2	CL3	NCL2	NCL3
CL20	-4666.0	54.3	20.2	14.2	1.3
CL21	-4111.3	81.8	20.7	4.7	3.8
CL2	-4077.7	27.5	39.0	7.8	2.3
CL3	-4088.8	68.9	32.6	4.3	4.8
CL10	-1000.2	72.8	28.8	3.9	1.4
CL11	-2143.3	72.9	28.1	3.6	4.5

The responses for the number of variables resulted in correct assignment on the squared distance between groups. Clusters from the correlated clusters are still considered the most distant for both untransformed and log transformed variables. However, a surprising correspondence between clusters 10 (Female CL10) and 11 (Female CL11) appears with cross-validation. This is compatible with the significant number of misclassifications involving these two clusters in the discriminant analysis of ratio variables.

PCA of the first subset of variables (Table 49) presented some differences in relation to the PCA with the whole set of variables. The variance composition (PC1 explained for both untransformed (Table 50a) and log transformed head size variables (Table 50b), but decreased for ratio variables (Table 50c). PC1 here accounts for a little more than half of the total variance of head size variables on hatchlings (0.341 and 0.304 for untransformed and log transformed variables, respectively). PC1 is responsible for approximately one fourth of a (0.244 for the former and 0.238 for the latter), whereas the remaining two principal components together account for approximately one fifth of the total variance in both cases. Similar to the results of PCA for the whole set of variables, the results above mean that the major source of variation between hatchlings from different clutches is the shape instead of the overall size of the head.

**Table 11: Linear discriminant analysis. Response: shock effect on loadings. Post shock, best subset of log-transformed total olive variables. Sample size: 124 records. Groups: clusters.**

*a) Summary of classification with group rotations*

Group	CL11	CL18	CL19	CL2	CL3	CL11	CL18
CL19	33	0	0	0	0	0	0
CL18	0	18	7	2	0	0	0
CL2	1	0	20	2	0	0	0
CL3	0	0	1	9	0	0	0
CL11	0	0	0	0	21	0	0
CL1	0	0	0	0	0	0	0
Total N	27	26	28	11	21	0	0
N Correct	26	18	25	9	20	0	0
Proportion	0.963	0.692	0.893	0.818	0.952	0.000	0.000

N = 124 N Correct = 106 Proportion Correct = 0.855

*b) Linear discriminant function for group*

	Constant	CL1	CL2	CL3	CL11	CL18
CL11	-10008	1950	1975	-412	-421	-421
CL18	-10008	1941	17810	344	-443	-443
CL2	-10017	2049	18888	341	-573	-573
CL3	-10039	1814	18940	380	-497	-497
CL1	-10144	1562	12191	340	-489	-489
CL19	-10169	1214	12150	340	-444	-444

**Table 12: Linear discriminant analysis. Response: shock effect on loadings. Pre-shock, best subset of olive variables. Sample size: 124 records. Groups: clusters.**

*a) Summary of classification with group rotations*

Group	CL18	CL19	CL2	CL3	CL11	CL1
CL19	12	2	0	0	0	0
CL18	2	13	2	1	0	0
CL2	0	4	19	0	1	0
CL3	1	0	0	11	2	0
CL11	0	0	0	0	14	0
CL1	0	0	1	2	2	0
Total N	17	24	20	14	20	0
N Correct	13	17	19	11	14	0
Proportion	0.765	0.708	0.950	0.786	0.700	0.000

N = 124 N Correct = 93 Proportion Correct = 0.742

*b) Linear discriminant function for group*

	Constant	CL1	CL2	CL3	CL11	CL18
CL19	-10010	1711.0	1881.0	1418.0	1792.7	1792.7
CL18	-10144	1679.0	1871.0	1488.0	1823.4	1823.4
CL2	-10010	1688.0	1828.4	1798.0	1798.1	1798.1
CL3	-10010	1683.0	1742.0	1873.0	1894.3	1894.3
CL11	-10113	1682.1	1813.1	1717.0	1870.4	1870.4
CL1	-10114	1692.7	1814.0	1801.1	1798.1	1798.1

**Table 12** Characterisation of buildings: Principal component analysis of first subset of variables (sample size: 124 examples)

*12) Eigenvalues of the covariance matrix of first subset of variables*

Eigenvalue	2.1136	0.8779	0.1775	0.0021
Proportion	0.561	0.224	0.045	0.000
Cumulative	0.561	0.786	0.831	1.000
Variable	PC1	PC2	PC3	PC4
CL	-0.134	0.088	0.138	0.000
SL	-0.033	0.087	0.035	0.004
WBL	-0.098	-0.080	0.025	-0.003
LH	-0.075	0.083	-0.060	0.004

*13) Eigenvalues of the covariance matrix of first subset of transformed first subset variables*

Eigenvalue	1.0000000	0.5000000	0.0000000	0.0000000
Proportion	0.500	0.250	0.000	0.000
Cumulative	0.500	0.750	1.000	1.000
Variable	PC1	PC2	PC3	PC4
SL	-0.000	0.000	0.000	-0.000
SL	-0.000	-0.000	0.000	0.000
WBL	-0.000	0.000	0.000	0.000
LH	-0.000	0.000	-0.000	0.000

*14) Eigenvalues of the covariance matrix of first subset of ratio variables*

Eigenvalue	1.0000000	0.5000000	0.0000000	0.0000000
Proportion	0.500	0.250	0.000	0.000
Cumulative	0.500	0.750	1.000	1.000
Variable	PC1	PC2	PC3	PC4
SL	-0.000	-0.000	0.000	-0.000
SL	-0.000	0.000	0.000	0.000
WBL	-0.000	-0.000	0.000	-0.000
SL:WBL	0.250	0.000	-0.000	-0.000

Belows load which, in 44 test, seems to be the major source of shape variation (see PC2 coefficients for WBL or Traffic risk and  $\beta_1$  and PC1 coefficient for BWI in Table 10a). This is also compatible with the patterns expressed by the clustering relations presented in Table 11.

Second second source is the crossflexion system with the relatively lowest load:  $\beta_2$  is. Therefore, interesting to discover that there is a considerable interoperative variation of this feature, even more surprisingly involving buildings. Although coordinates are usually consistent geometric features, there seems to be a consistent correlation between shaft shape and loading factor, but the interoperative implication of this correlation have not yet been studied.

Figure 18 shows the plots of buildings against values for the two principal components of the whole set and the first subset of transformed (Figure 11a up to) and log transformed



head-size variables (Figure 1b) and 4), and mass variables (Figure 1d) and 5). It is possible to see the clusters (P1 and P2) from multivariate linear (GLS) and (LIS) respectively, but clusters apart from the others although cluster P2 is not really identified. These patterns are comparable with the results showed by discriminant analysis. The female family effect on body shape morphology is discussed in the next section (3.2.2). See Figure 2 for visually compare both cluster and female family effects.

However, the plots present only PC1 and PC2 coefficients, which usually represent no more than two-thirds of the total variation. If new axes were added to the plots of Figure 1 B concerning the other principal components, the distribution of the clusters would become even more visible. However, plots were kept in two dimensions (only PC1 and PC2) using for the sake of simplicity.

Clutch effect on the cranial morphology of *h. halimae* is evident based on the results above. Log transformation of data did not significantly improve the models presented. Body variables presented again better proportions of correct classification in discriminant analysis and complex patterns in PCA. Reduction in the number of variables represents a decrease of approximately 90% in the efficiency of the discriminant analysis, but it may be useful depending on the process intended.

Clutch effect has been described by Lang and Anderson (1994) as significantly affect the pattern of temperature-size determination in *crustaceans*. In the present study, clutches have been discriminated with approximately 90% of success in an extremely homogeneous group of captive animals with controlled environment of eggs incubation. Since female *crustaceans* build up (or dig) self-clad nests, clutches are usually isolated from each other in the wild. Even nests partitioned by one or females usually show separate egg clutches. Thus, clutch discrimination may result in the consistency of phenotype among individuals.

For this regarding, clutches show an extremely significant. The establishment of phenotype among individuals in long-term studies, may be useful for the comprehension of extremely important ecological-behavioral variables, such as mating system and dispersal pattern. These variables are intrinsically related to the population structure and with population dynamics.

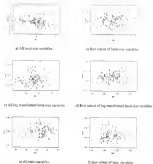


Figure 11: Change effect on head size: Plots of head size against values for the two grouped covariates (Climatic/size factors) (a) (CLIMATIC) (b) (CLIMATIC) (c) (CLIMATIC) (d) (CLIMATIC) (e) (CLIMATIC) (f) (CLIMATIC).

### 4.2.4 Female-family effect on bodylength

There is a significant female-family effect on the annual complementarity of bodylengths ( $F$  values >1.000 for Wall's, Lawley, Shrinkage, and Pillai's Tests of MANOVA) for the distinction of unimodal and bi-modalised head size variables, relative variables. The same goes between bodylengths: cranial measurements and female-family is presented in Table 4.6. Significant differences in MANOVA mean the clusters first different modalised images present significant differences body dimension relations between SVL and the complementarity variable in question. The biological meaning for this is a significant difference in bodylengths: cranial shape between female-families.

None of the cranial head size variables (DCL, CW, BL, CL, QW, HW, LCB, ML, and LMJ) bodylength of elements variables (SCN, BLST, SPWT, BCL, SPFL, SPVN, SPNL, and BLAQ) showed a significant difference between female-families at covariance of bodylengths<sup>2</sup> SVL. This is consistently apparent in the number of variables influenced by shape-effect body (however, there are also scenarios for any predominance of width- or length-variables as a predominance of the variance over the variable). This means that head length first different modalised images, presenting similar body size, tend to successively vary in skull shape as body relative width and length.

Table 4.6. Analysis of covariance: bodylength-family as covariate of bodylength<sup>2</sup> SVL.

Variable	FV	Female-family	Female	SVL	Female-family
DCL	0.000	0.000	BLW	0.000	0.000
CW	0.000	0.000	BLST	0.000	0.000
BL	0.000	0.000	SPWT	0.000	0.000
HW	0.000	0.000	QW	0.000	0.000
CL	0.000	0.000	LCB	0.000	0.000
QW	0.000	0.000	HW	0.000	0.000
HW	0.000	0.000	SPVN	0.000	0.000
LCB	0.000	0.000	SPNL	0.000	0.000
SPN	0.000	0.000	BLAQ	0.000	0.000
SPN	0.000	0.000	BLAQ	0.000	0.000
ML	0.000	0.000			
LMJ	0.000	0.000			
SCN	0.000	0.000			

Interdependent: Sex → ANOVA → 4 Covariates Linear Model

Response (dependent variable): bodylength new variable

Model (independent variable): SVL

Covariate: Female-family (to separate bodylengths)

Linear discriminant analysis with cross-validation for the whole set of both untransformed (Table 40) and log-transformed head-size variables (Table 41) showed the maximum proportion of correct classifications (100%) for female lambs. Early variation (F-test:  $P$ -value) was of 0.14. Assuming that conditions usually present a low genetic variability, and the group of animals under the study is extremely homogeneous, these results mean that multivariate analysis may be extremely sensitive to morphometric variables.

Repeat discriminant between groups showed a pattern similar to the one presented in the preceding sections (4.3.2.3). The group of related females (Family A2) seems to be more discriminating process in relation to the other unrelated females (Families H1 and H3). The same pattern was found in the analysis of egg morphology. As seen in the former sections some family variance can be detected and a constant clutch-effect. However, multivariate analysis seems to be the right means of variables for the visual morphology of hatchlings.

Table 40: Linear discriminant analysis: Response: female family effect on hatching. Predictor: all head size variables. Sample size: 124 hatchlings. Groups: clutches.

4) Summary of classification with cross-validation

Group	1	2	3
1	21	0	0
2	0	76	0
3	0	0	20
Total N	21	76	20
N Correct	21	76	20
Proportion	1.000	1.000	1.000

N = 124. N Correct = 124. Proportion Correct = 1.000

4.1) Linear discriminant function for group

	Constant	SN1	SN2	SN3	SN4	SN5	SN6	SN7	SN8	SN9	Total
1	-1322.4	22.7	16.3	-0.7	13.6	27.4	28.4	54.2	40.5	64.8	26.2
2	1444.4	22.1	17.4	0.4	26.8	28.1	18.1	59.4	57.9	28.2	22.8
3	-1543.6	22.8	22.8	-1.1	-25.2	20.5	10.3	56.3	22.8	27.8	23.8

	SN1	SN2	SN3	LM
1	18.4	10.1	1.1	26.4
2	16.3	1.4	-1.8	-11.2
3	18.2	10	-4.3	12.3

Principal components analysis of hatchlings were presented and discussed in the former section (4.3.2.3). See Tables 42 - 44 for results. Table 42 shows the first six values of variances for the analysis of female-family effect on hatchlings. Readably in the former section, log

transformations did not lead to different selection of variables. However, the best subset for the study of clutch-effect on hatchlings was different from the variables for the study of family identity. Surprisingly however, the coefficient of determination ( $r^2$ ) of the best subsets of variables for the former were higher than for the later.

**Table 66.** Linear discriminant analysis. Response: Family-family effect on hatching. Features: all log-transformed head-size variables. Sample size: 124 hatchlings. Groups: clutches

*Discriminatory of classification with cross-validation*

Group	1	2	3
1	17	9	0
2	9	18	4
3	4	9	21
Total/ln	27	34	25
McLennan	27	34	25
Proposed	1 000	1 000	1 000

$N = 124$ .  $K$  Correct = 124. Proposed-Correct = 1 000

*ln Linear discriminant function for group*

	Constant	PC1	PC2	ln	lnW	lnL	lnW	lnHL	lnCL	lnSL	lnTL
1	-60.84	1.1717	0.07	-0.46	-0.71	0.03	0.68	0.01	0.04	-0.21	0.04
2	64.97	1.0102	0.08	-0.07	0.68	0.01	0.0	0.02	0.04	0.00	0.41
3	102.13	1.0170	0.05	-0.44	0.44	0.01	0.0	0.02	0.00	0.00	-0.17

	lnL	lnW	lnSL	lnTL
1	0.00	-0.00	0.00	0.00
2	0.00	0.00	0.00	0.00
3	0.00	0.00	0.00	0.00

The reduction in the number of variables (Table 66) did not affect the significance of the difference between family-families for all kinds of variables ( $P$ -value < 0.000 for Wyle's, Levene-Hausding, and Pillai's Tests of MANOVA). Linear discriminant analysis with cross-validation of the best subsets of transformed (Table 67) and/or transformed head size variables (Table 70), and size variables (Table 71) demands slight decrease in the overall proportion of correct classification ( $P$  97%–975% and 94.1% respectively). Once again, log-transformations did not improve analytical data, and size variables performed better efficiency than head size variables.

PCA of the best subsets of variables is presented in Table 72. PC 1 is responsible for most of the variation in data. It varies from 8.53% for the size variables to 6.44% for the log-transformed variables, with intermediate variables presenting an intermediate value of 6.41%.

These PC1 values are significantly higher than the ones presented by the first principal component for the first subsets of variables in the study of clutch effect.

**Table 17.** Linear discriminant analysis: Response family-family effect on hatchling. Predicted all nine variables (Sample size: 114-hatchlings). Group: cluster

*a) Summary of classification with group-variables*

Group	1	2	3
1	26	1	1
2	1	47	1
3	0	0	19
Total N	27	48	20
N Correct	26	47	19
Percentage	0.963	0.979	0.950

$N = 124$ ,  $N$  Correct = 111, Accuracy (Group) = 0.943

*b) Linear discriminant function for group*

	Constant	PC1	PC2	PC3	PC4	PC5	PC6	PC7	PC8	PC9	PC10	PC11
1	-1421.0	641.7	1011.0	126.7	3251.1	1487.1	143.6	1461.9	1755.2	2666.0		
2	-4004.4	434.2	1471.1	784.0	1784.2	1571.3	143.7	-1082.2	1198.1	2534.9		
3	2833.0	177.0	1611.1	476.1	2753.6	1241.1	243.5	1061.4	1664.1	1476.2		

	Wilk's	$\eta^2$ (1, N)
1	0.072	0.004
2	0.049	0.004
3	0.043	0.004

**Table 18.** First subsets of variables by family-family effect on hatchlings

Variable	Test statistic	F value	p	N
Head size	1.91, 54, 55, WSS	1.566	0.001	124
Log transformed head size	CW, 55, 55, WSS	1.007	0.002	124
Body	0.477, 505, 506, 111, 109	2.022	0.001	124

Statistical properties: Head  $\rightarrow$  Regression,  $\rightarrow$  First Subset

A possible biological meaning for these results is that overall skull size may be the major source of variation among hatchlings from different maternal lineages, whereas skull shape seems to be the major variation between relatives of the same female family. Skull size is strongly correlated with body size. Thus, the results above may imply that hatchlings from different maternal lineages tend to present correlated variation in body size. Body size, in its turn, may be strongly related to survival rate of hatchlings during the first winter, and consequently to the relative reproductive success of the female. Relative reproductive success, in its turn, is directly

related to fitness (Kruke 1992, 2002). Conversely, parental strategies could interact with possible ranking hierarchy among females (Long 1967). Differential reproductive success may be related to the female social ranking position. Mated and female small therefore be related to the maintenance of social rank, which could also be related to sex determination, analogously to the pattern described by Simpson and Simpson (1962) for human siblings.

**Table 11** Linear discriminant analysis: Response female family effect on hatchlings. Posterior best subset of log-transformed head size variables. Sample size: 128 animals. Groups: clutches

*A Summary of classification with cross-validation*

Group	1	2	3
1	26	1	0
2	1	74	4
3	0	1	20
Total N	27	76	24
N Groups	26	74	24
Proportion	0.553	0.973	1.000

N = 128 N Correct = 121 Response Correct = 0.939

*K-Learn discriminant function by group*

	Constant	C/N	SA	AS	WPA
1	-183.22	37.25	4.71	34.28	-49.13
2	-454.16	37.14	1.66	30.99	-10.63
3	-454.48	37.32	1.62	32.46	-10.63

**Table 12** Linear discriminant analysis: Response female family effect on hatchlings. Posterior best subset of log-transformed head size variables. Sample size: 128 animals. Groups: clutches

*A Summary of classification with cross-validation*

Group	1	2	3
1	26	1	0
2	1	74	0
3	0	1	20
Total N	27	76	21
N Groups	26	74	20
Proportion	0.553	0.973	1.000

N = 124 N Correct = 121 Response Correct = 0.936

*K-Learn discriminant function by group*

	Constant	C/N	SA	AS	WPA
1	-157.60	44.944	-43.3	3379.3	-188.4
2	-1054.0	44.243	30.9	3026.5	540.4
3	-1027.4	44.413	45.3	3552.7	-188.3

**Table 11** Linear discriminant analysis: Response: female-female effect on hatchlings. Predictors: log values of four variables. Sample size: 124 animals. Groups: clinches

**a) Summary of classification results (log-transformed)**

Group	1	2	3
1	25	1	0
2	2	59	3
3	8	14	59
Total N	35	74	62
N Correct	25	59	59
Proportion	0.714	0.797	0.952

N = 124. McFadden = 0.12. Adjusted Likelihood = 14.23

**b) Linear discriminant function for group**

	Constant	BL1	BL2	BL3	BL4
1	1811.5	333.7	3693.8	1562.2	1768.3
2	-1493.8	5483.1	-648.8	1867.3	1322.8
3	7734.8	2559.2	-2228.8	1495.1	5683.4

**Table 12** Female-female effect on hatchlings. Principal component analysis of the log values of variables. Sample size: 124 animals

**a) Eigenvalues of the correlation matrix of log values of four variables**

Eigenvalue	1.000	0.294	0.198	0.194
Proportion	0.403	0.119	0.081	0.080
Cumulative	0.403	0.522	0.603	0.683
Variable	PC1	PC2	PC3	PC4
BL1	0.597	0.229	-0.133	0.758
BL2	0.548	-0.270	-0.758	0.014
BL3	0.333	0.470	-0.891	-0.508
BL4	0.333	0.858	0.099	-0.451

**b) Eigenvalues of the correlation matrix of log values of log-transformed four variables**

Eigenvalue	0.997965	0.000433	0.000002	0.000000
Proportion	0.998	0.000	0.000	0.000
Cumulative	0.998	0.999	0.999	1.000
Variable	PC1	PC2	PC3	PC4
Log	-0.043	0.000	-0.000	-0.000
Lat	-0.798	-0.000	0.000	0.000
Lal	-0.998	-0.000	-0.000	0.000
Lrr	-0.532	0.000	-0.000	0.000

**c) Eigenvalues of the correlation matrix of log values of six variables**

Eigenvalue	0.999247	0.000753	0.000000	0.000000	0.000000
Proportion	0.999	0.000	0.000	0.000	0.000
Cumulative	0.999	0.999	0.999	0.999	1.000
Variable	PC1	PC2	PC3	PC4	PC5
BL1	-0.002	-0.000	-0.000	-0.000	0.000
BL2	-0.000	-0.000	-0.000	-0.000	-0.000
BL3	0.000	-0.000	-0.000	-0.000	-0.000
BL4	0.000	0.000	-0.000	-0.000	0.000



Figure 18 shows the plots of loadings against values for the two principal components of the whole set and loadings of arcsinh-transformed (plots a, b) and log-transformed head size variables (plots c, d), and size variables (plots e, f). It is possible to see how arcsinh-transformed data, unlike head size data, are more multi-modal. Great again, log-transformations did not improve consistency of number and size variables were less effective than head size variables. It is interesting to compare these results with clutch effect in Figure 13 and apparent clutch effect on hatchlings shown in Figure 20.

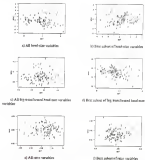


Figure 18. Female clutch effect on the hatchlings. Plots of loadings against values for the two principal components. Variables: (a) CL1(2), (b) CL1(3), CL3, CL4, CL5, CL6, CL7(3), (c) CL1(2),

### 4.5.2.5 Main effect on hatchlings

Maternal effect on hatchlings could not be completely isolated from clutch and female family effect because the two different types of large sample test used provided an experimental design, which was impossible to obtain with the available samples. However, as an exploratory exercise the "supposed" main effect is here presented and discussed. That effect seems to be significant ( $P$  value  $< 0.005$  for both  $\chi^2$ , Levene, Bartlett, and Fisher's Permutation MANOVA). Lower discriminant analysis with cross-validation was presented in Table 73 for the whole set of non-transformed head size variables, in Table 74 for the log-transformed head size variables, and in Table 75 for mass variables.

Table 73 Lower discriminant analysis: Response: main effect on hatchling. Partition of head size variables. Sample size: 11 hatchlings. Group: males.

#### 3) Summary of classification with cross-validation

Group	CL283	CL39	CL1
CL283	40	1	1
CL39	1	40	0
CL1	0	4	20
Total N	40	44	20
N Correct	40	44	20
Proportion	1.000	1.000	1.000

N = 104 N Correct = 104 Proportion Correct = 1.000

#### 4) Lower discriminant analysis by group

	Count	CL1	CL39	CL1	CL39	CL1	CL39	CL1	CL39	N%
CL283	40 (38.5)	38.50	-1.93	-18.20	-1.93	-33.20	-4.13	30.75	20.25	21.11
CL39	44 (40.4)	20.04	19.92	18.46	-0.90	19.09	4.43	23.66	21.07	24.09
CL1	20 (19.3)	21.24	0.75	-14.49	1.41	13.08	8.61	11.14	20.28	11.75

	FVS	MS	1.683	9.038	1.42
CL283	27.36	0.75	30.31	1.03	1.01
CL39	34.01	1.09	27.73	11.17	0.40
CL1	32.79	0.70	20.43	11.43	1.41

Log-transformed data again did not improve the results of the lower discriminant analysis with cross-validation. The proportions of correct classification varied from 0.914 for the six variables to 0.947 for both non-transformed and log-transformed head size variables, which is still considerably high. Male CL1 score of the male CL283. However, squared distances between groups show that male CL39, although isolated, seems to be in an intermediate position between them.

Principal component analysis of headlings are presented and discussed in a later section (4.3.3.3). See Tables 34–36 for results. Figure 36 shows the plots of headlings against values for the two principal components for instantaneous (Figure 36a) and log transformed head size variables (Figure 36b) and size variables (Figure 36c). Comparing these plots to the ones presented in Figure 15, it is evident that headlings are clustered according to female family and not to the male, in other words, discriminant analysis could not exclude male effect from female-family and clutch effect, resulting in an erroneous interpretation of its results. This finding can be justified by the principal components analysis. This shows that male effects on overall morphology of headlings are consistently less marked than parent-based changes. The discrimination of its overall magnitude is attributable to experimental design.

Table 34: Linear discriminant analysis: Response: male effect on headling; Prediction of log-transformed head size variables. Sample size: 114 headlings. Groups: males

a) Summary of classification with prior probabilities

Group	CL100	CL50	CL1
CL100	88	9	1
CL50	7	40	3
CL1	1	4	34
Total	96	48	38
McNemar	88	40	34
Response	0.814	0.833	0.891

N = 124    McNemar<sup>2</sup> = 82    Response Error = 0.047

b) Linear discriminant analysis by group

	Constant	CL1	CL5	CL10	CL15	CL20	CL25	CL30	CL35	CL40	CL45
CL100	-1.067	0.042	8	-0.04	0.07	0.06	0.06	0.07	-0.04	-0.02	-0.02
CL50	-1.043	0.008	0.8	-0.01	0.01	0.009	-0.008	0.01	0.01	0.01	0.02
CL1	-0.009	0.003	0.01	-0.001	0.001	0.001	0.001	0.001	0.001	0.001	-0.001

	P < .05	W	LR	Wald	LR
CL100	0.00	0.00	0.00	0.00	0.00
CL50	0.00	0.00	0.00	0.00	0.00
CL1	0.00	0.00	0.00	0.00	0.00

### 4.3 Morphometrics: Analysis of Wild Animals

The field data where wild animals were captured for the present study are described in Section 3.4. All of them are located in the Bar-Central region of São Paulo State, Brazil (Figure 3). They are all connected to subwatersheds of Tietê River (the most river of the State of São Paulo) which runs westward and across almost the whole State. *Pithecheilus* and *Pomocidichneumon*

original nucleus of *Proculiteris* group) one of the more north subnuclei of *Ilheus* River. *Pinimã* was merged nucleus and *Chaparrão* is an artificial pond both being situated in *Assaí* Creek, a subnuclei of *Proculiteris* River. These field sites are located no further than 15 km from each other. *Chaparrão* is the more recent field site located approximately 150 km far from the others. It is a group of artificial ponds connected to *Pedernegado* Creek, a south subnuclei of *Ilheus* River.

Table 71: Linear discriminant analysis: Regression model effect on breeding. Posterior: all sites variables. Sample size: 124 (breeding) Groups: males.

Group	CL300	CL30	CL1
CL300	48	3	1
CL30	4	20	3
CL1	1	4	23
TotalN	48	48	28
N Cases	48	27	26
Posterior	0.188	0.379	0.183
N = 124 N Cases = 97 Regression Cases = 104			

	Constant	B*W	B*H	B*LT	B*V	W*W	B*V	B*V	B*V
CL300	-1049.4	408.7	0.000.2	408.6	-0.78.9	1.423.9	351.4	-0.000.0	0.000.1
CL30	-408.1.0	408.0	0.000.7	408.6	-0.88.0	1.394.9	350.3	-0.134.0	0.000.1
CL1	-2723.6	407.0	0.000.3	408.0	-0.789.3	1.399.2	351.3	1.191.0	0.000.4

	R <sup>2</sup> (N)	F (W)	R <sup>2</sup> (H)
CL300	0.000.4	0.00.4	0.000.0
CL30	0.000.6	0.00.0	0.000.0
CL1	0.000.7	0.00.0	0.000.1

Studies of morphometric geographic variation involving *Proculiteris* have traditionally been related with taxonomy of similar taxa (Black 1924, Peabody and Leitch de Carvalho 1941, Leitch de Carvalho 1973, Mazon 1943, Hall 1950, Mazon and Sogari in press), although some studies have also been related to intraspecific variation (Black 1924, Gordon 1971, Smith and Mazon 1978, Mazon 1982, Aguiar 1994, Hall and Petrus 1994).

In the present study morphometric analysis of wild animals was carried out with the objective of establishing differences between the subpopulations studied (Section 4.4.1), and also between wild and captive animals (Section 4.4.2). Groups are studied as described in Figure 1.



a) Kopf-Höhe (cm)



b) Körpergröße (cm)



c) Körpergewicht (kg)

Figure 30: Apparent static effect on head height: Plot of head height against values for the first principal component. Model: (a) (CL200) (b) (CL170) (c) (CL1)

#### 4.4.1 Differences among subpopulations

There is a significant difference among subpopulations from the area studied in relation to the general morphology of both skulls ( $P$ -value < 0.001 for  $WBL$ ,  $L$ , Lowering Hopling, and  $PBL$ ;  $\chi^2$  Test; MANN- $\chi^2$  for all kinds of variables). However, linear discriminant analysis with cross-validation for auto-transformed (Table VI) and log-transformed head size variables (Table VII), and size variables (Table VIII) showed consistently low overall proportion of correct classifications. Log-transformations did not significantly increase model efficiency (0.455 against 0.421 for auto-transformed data) and size variables performed the lowest rate of correct classifications (0.482).

The squared distance between groups generated in the pattern independently of the kind of data transformation (e.g., log transformation). However, skull sizes (Table IX) and  $P$ -values were generated in the next domain size between each other, which does not correspond to their geographic location. In addition,  $P$ -value distances between were generated as the direct size between each other, which does not correspond either to their geographic location. These 6 reasons for this unexpected pattern are discussed below.

PCA of the whole set of head size variables is presented and discussed in Section 4.2.1 (see Table IV). PCA of log-transformed head size variables (Table VI) showed similar patterns, with PC1 as the major source of variation (81.6% for the linear and 81.9% for the linear). PC1 accounted for approximately 80% of the total variation in both cases.  $QW$  and  $EW$  were respectively the major sources of variation in PC1 for the non-transformed and log-transformed variables, although the contrast between variables in the linear is not as clear as in the linear.

PCA for size variables (Table VII) showed a different pattern, with PC1 accounting for less than half and PC2 for approximately one fourth of the total variation.  $ROW$  is the major source of variation in PC1, which seems to be compatible with the pattern showed by PCA for head size variables.  $EW$  and  $EW$  are the major source of variation in PC2, which possibly represents a contrast between the relative width of skull and the relative interorbital length. This contrast is similar to the size variability around the orbit, as discussed in Section 4.2.1.1.

Table III shows the best subsets of variables for the morphometric analysis of the differences among subpopulations.  $WBL$  variables performed in the best subsets of both auto-transformed and log-transformed head size variables, whereas length and width variables contributed in the best subset of size variables. Variables located long in the variation and width is represented in the best subsets, which possibly means that differences among subpopulations involve overall size and shape variations in the whole skull of otomys.

Table 16 Linear discriminant analysis: Response rates of target Population of least-cost variables: Group 1 (study case 1) Folia-Groenle 2 *Prunella arvensis* 3 *Prunella* 4 *Chamaecrista* and 5 *Drosera* (Sample size: 20 animals)

Classification of Calculated variables with cross-validation					
Group	1	2	3	4	5
1	0	1	0	0	0
2	2	9	1	0	0
3	0	1	1	1	0
4	1	0	0	1	2
5	0	1	0	0	1
Total N	3	12	2	2	3
N Correct	0	9	1	1	1
Proportion	0.000	0.750	0.500	0.500	0.333
N = 20 N Correct = 10 Proportion Correct = 0.500					

Table 17 Linear discriminant function for group

Constant	PC1	PC2	PC3	PC4	PC5	PC6	PC7	PC8	PC9	PC10
1	-285.26	11.09	11.20	-28.90	1.29	1.80	32.75	15.44	47.11	82.58
2	-405.64	39.76	-84.72	25.84	8.28	-10.89	19.23	12.04	68.92	82.29
3	-573.28	13.23	11.69	23.68	38.12	4.60	20.83	10.93	31.34	75.14
4	-595.84	14.07	-19.80	54.44	2.74	1.75	24.54	41.09	23.77	75.86
5	-279.80	13.26	-41.28	32.07	7.11	-4.22	28.76	11.34	32.07	89.11

	PC1	PC2	PC3	PC4
1	-1.76	1.28	-0.80	1.04
2	0.48	1.40	-1.00	-1.31
3	1.05	0.19	-0.97	1.23
4	-1.76	1.89	-1.25	-1.34
5	0.28	1.28	-1.10	0.11

A significant difference among subpopulations still persists after reduction in the number of variables ( $P$  value < 0.001 for Wall's, Likelihood, and Fisher's Tests of MANOVA) for the best subset of untransformed and log-transformed least-cost variables and error variables. Linear discriminant analysis with cross validation showed an overall improvement in the rate of correct classification for the three levels of variables (Table 12–14).

The spread distance between groups showed a similar pattern for the least-cost variables (both untransformed and log-transformed), with Folia-Groenle and *Prunella* as the most distinct sites as before the reduction in the number of variables. However, the best subset of error variables presented three two sites as the closest ones. This paradoxical pattern is possibly related to the small sample size. Individuals' age and gender may also influence the results shown, since these sources of variation could not be related.

Table 17 Linear discriminant analysis: Between sites of origin. Predictors: all log-transformed head size variables. Groups: study sites (1) Lake George, 2) Potomac stream, 3) Potomac 4) Chesapeake, and 5) Delaware. Sample size: 29 stomachs.

**5. Summary of classification with error reduction**

Origin	1	2	3	4	5
1	1	0	0	0	0
2	0	1	1	0	0
3	0	0	0	0	1
4	0	0	1	0	0
Total N	4	10	2	2	1
Success	1	9	1	0	1
Accuracy	0.774	0.700	0.500	0.000	0.500

N = 20, N(Lake) = 10, Potomac Group = 0.000

**5. Linear discriminant functions by group**

Group	LD1	LD2	LD3	LD4	LD5	LD6	LD7	LD8
1	-402.7	1002.8	7606.4	1236.0	1841.0	1381.9	10420.1	722.4
2	-4021.9	1381.2	-7504.4	3421.0	1543.7	1319.3	2380.3	-794.0
3	1811.7	7619.4	-3591.9	-1242.0	1021.0	1286.2	1381.9	861.4
4	-4233.4	5617.1	-7524.0	2729.0	1481.6	-1336.4	2084.0	-481.9
5	-4194.0	1375.8	-3341.3	2875.0	1780.1	-1481.3	2481.4	717.1

	Wp	Wd	Wt	WHL	WHL
1	11271	271.2	24717	1302.3	-792.4
2	1384.8	881.1	1711.0	1201.4	-481.9
3	2532.6	819.0	1481.4	1028.1	-481.9
4	1447.7	807.1	2181.7	1236.1	1204.0
5	1446.8	789.1	2001.0	1481.3	-475.1

Principal component analysis of the ten subsets of variables are presented in Table 18. PC1 accounts for more of the total variance for all landmark variables, 0.546 for untransformed head size variables, 0.562 for log-transformed head size variables, and 0.400 for total variance. This states that we overall account for more of the size variance of variation among stomachs from different areas of origin, which is similar to the pattern of geographic variation described by Johnson (1985) for North American wolves. However, this may be related to clonal environmental influence.

Figure 20 shows the plots of total stomach size of origin against values for the two principal components. It is possible to see that individuals from Delaware are clustered apart, which is reasonable as the geographic location of that study site. However, the other sites are not well differentiated by the plots. Log-transformations did not significantly improve interpretation of the principal component analysis. Basic variables showed an unclear pattern of distribution of subpopulations.



Table 76. Linear discriminant analysis: Regression rates of origin. Predictor: all nine variables. Groups: study sites (1: Falsa/Groenle; 2: Forno de dentro; 3: Passosul; 4: Champagneul; and 5: Darglieve). Sample size: 29 animals

*assignment of individuals with cross-validation*

Group	1	2	3	4	5
1	4	2	0	1	0
2	1	8	2	1	1
3	0	1	0	0	1
4	0	0	0	0	1
5	0	1	0	0	0
Total N	5	12	2	2	3
N Correct	4	8	0	1	1
Percentage	80%	66.7%	0%	50%	33.3%

N = 29. N Correct = 14. Percentage Correct = 48.3%

*12 Linear discriminant functions by group*

	Constant	BCW	BCZ	BAW	BCV	BCW	BCV	BCZ	BCW
1	-1811.1	231.2	1382.4	1048.1	1941.2	1755.3	1909	-4854.7	-181.2
2	2798.0	184.0	1123.6	2664.4	4880.0	1780.1	198.3	-4888.4	-487.0
3	-1832.2	-426.7	1507.8	2716.2	1861.2	1890.4	811.7	-4833.8	-1831.0
4	1831.7	54.5	1098.0	1886.4	1878.0	1875.1	-4203.8	-4761.2	-1884.4
5	1817.9	-284.8	1122.1	2628.7	4952.7	1755.3	-4713	-4824.2	-1817.9

	BCW	BCZ
1	8338.1	783.9
2	8438.4	-713.0
3	8132.1	149.4
4	8684.8	873.0
5	8138.1	761.1

Table 78. Principal-component analysis: Eigenanalysis of the covariance matrix of all log-transformed head size variables. Sample size: 29 animals. Only first six principal components are presented

Eigenvalue	0.000711	0.000700	0.000676	0.000668	0.000633	0.000611
Percentage	8.18%	8.16%	8.12%	8.07%	7.98%	8.04%
Cumulative	8.18%	16.34%	24.46%	32.53%	40.51%	48.55%
Variable	PC1	PC2	PC3	PC4	PC5	PC6
BCZ	-0.181	0.124	0.058	-0.040	0.048	0.144
CV	-0.284	-0.083	-0.020	0.028	0.099	0.086
GL	-0.136	0.287	0.108	-0.175	-0.113	0.141
BAW	-0.064	-0.064	-0.170	0.080	0.121	0.201
BCV	-0.084	0.084	-0.020	-0.040	0.174	-0.170
BCW	-0.113	0.100	0.043	0.080	0.123	-0.087
BCW	-0.121	-0.020	0.048	0.028	0.080	-0.090
BCZ	-0.064	-0.020	0.040	-0.210	-0.186	-0.021
BCV	-0.113	-0.114	-0.040	-0.020	-0.084	-0.081
BCZ	-0.020	0.283	0.084	0.087	-0.020	-0.204
BCV	-0.041	0.060	-0.101	-0.027	-0.142	0.201
BCW	-0.040	0.100	0.074	-0.170	-0.090	0.171
BCZ	-0.144	-0.021	-0.070	0.040	0.111	0.024

**Table 10: Principal component analysis: Eigenanalysis of the covariance matrix of all zero variables. Sample size: 20 animals. Only first six principal components are presented**

Component	Eigenvalue	%var	%cum	Component	Eigenvalue	%var	%cum
Arterial	0.432	0.273	0.274	1.341	0.042	0.026	0.698
Cardiac	0.333	0.207	0.481	0.933	0.043	0.026	0.754
Variable	PC1	PC2	PC3	PC4	PC5	PC6	PC7
Wt (kg)	-0.023	-0.011	0.004	0.003	-0.004	0.000	0.000
SLST	-0.189	0.000	0.000	0.000	-0.179	-0.003	-0.003
CRST	0.000	-0.000	-0.000	-0.000	0.000	-0.000	-0.000
BDL	0.100	0.073	0.001	0.003	-0.100	0.000	0.000
BCR	-0.000	-0.000	-0.000	0.000	-0.000	-0.000	-0.000
TRST	-0.000	-0.000	0.001	-0.000	0.000	0.000	0.000
CRST	-0.000	-0.000	0.000	0.000	-0.000	-0.000	-0.000
SPSL	-0.000	-0.000	0.000	0.000	-0.000	-0.000	-0.000
BCR	-0.000	0.000	-0.000	0.000	0.000	0.000	0.000
TRST	0.000	-0.000	0.000	0.000	0.000	0.000	0.000

**Table 11: Best values of morphological variables for differentiating gender subcategories**

Variable	Best values	F value	p	N
Heart size	0.00, 0.00, 0.00, 0.00	0.000	0.000	20
Bay-transformed heart size	0.00, 0.00, 0.00, 0.00	0.000	0.000	20
Ratio	0.00, 0.00, 0.00, 0.00	0.000	0.000	20

Linear procedure Sex vs. Regression on Best Values

**Table 12: Linear discriminant analysis: Regression score of origin. Posterior best values of best size variables. Group: study site (1: Polje Crno, 2: Polje Crno, 3: Polje Crno, 4: Polje Crno, 5: Polje Crno, 6: Polje Crno, 7: Polje Crno, 8: Polje Crno, 9: Polje Crno, 10: Polje Crno, 11: Polje Crno, 12: Polje Crno, 13: Polje Crno, 14: Polje Crno, 15: Polje Crno, 16: Polje Crno, 17: Polje Crno, 18: Polje Crno, 19: Polje Crno, 20: Polje Crno)**

discrimination of classification with zero variables						
Group	1	2	3	4	5	6
1	0	0	0	0	0	0
2	0	0	0	0	0	0
3	0	0	0	0	0	0
4	0	0	0	0	0	0
5	0	0	0	0	0	0
6	0	0	0	0	0	0
Total N	0	0	0	0	0	0
Mean	0	0	0	0	0	0
Regression	0.000	0.000	0.000	0.000	0.000	0.000

N = 20, 20 Cases, 0.00, Properties Correct = 0.00

**Table 13: Linear discriminant analysis for group**

Variable	1	2	3	4	5	6
1	-0.000	-0.000	-0.000	-0.000	-0.000	-0.000
2	-0.000	-0.000	-0.000	-0.000	-0.000	-0.000
3	-0.000	-0.000	-0.000	-0.000	-0.000	-0.000
4	-0.000	-0.000	-0.000	-0.000	-0.000	-0.000

Table 10: Linear discriminant analysis: Response rates of origin. Posterior best values of log transformed baseline variables. Groups: study-sites (3: Folia-Densely, 2: Folia-Ar, 4: Archa), Potential: 4: Chlorophyll, and 3: Droughtless. Sample size: 29 animals

*Summary of classification with cross-validation*

Group	0	1	2	3	4
3	0	1	0	0	1
2	0	10	0	0	0
4	0	0	0	0	0
3	0	0	0	1	0
4	0	1	0	1	2
Total N	0	12	0	1	3
No (Group)	4	10	1	1	7
Proportion	0.067	0.833	0.166	0.143	0.429

N = 29    N-Excluded = 21    Proportion Correct = 0.724

*Log linear discriminant function for group*

	Constant	LOG	LOG	LOG	LOG
3	-442.0	184.1	20.1	180.1	-41.0
2	175.0	-454.2	-100.0	80.0	147.0
4	587.0	-489.0	84.0	110.0	-89.0
3	-173.0	-150.0	-18.0	300.1	-89.0
4	-687.4	-833.1	33.2	179.2	120.4

Table 11: Linear discriminant analysis: Response rates of origin. Posterior best values of log transformed baseline variables. Groups: study-sites (1: Folia-Densely, 2: Folia-Archa), Potential: 4: Chlorophyll, and 3: Droughtless. Sample size: 29 animals

*Summary of classification with cross-validation*

Group	0	1	2	3	4
1	7	0	0	0	0
2	1	0	0	0	1
1	1	1	1	0	0
2	0	1	1	0	0
4	0	0	0	0	0
Total N	9	2	2	0	1
No (Group)	7	0	1	0	2
Proportion	0.778	1.000	1.000	0.000	0.400

N = 20    N-Excluded = 10    Proportion Correct = 0.700

*Log linear discriminant function for group*

	Constant	LOG	LOG	LOG	LOG
1	-753.7	1337.1	-82.1	544.4	1370.4
2	-131.2	1307.0	-89.2	544.0	1309.7
1	-753.7	1330.3	14.0	100.7	1009.7
2	-884.0	1024.1	104.0	542.0	1007.4
4	-451.0	1040.4	-100.0	100.0	1004.1

It is interesting to compare Figure 11 with Figure 8 in which the estimated age of wild animals presented a surprising result. It is possible that age – not size of organs – is the major source of variation between wild animals captured. Sexual dimorphism is not as clear as wild as it is in captive animals, possibly because the wild sample is young, and males are used. In addition, other environmental factors, such as food availability, may also have influenced the results shown.

Table 41. Principal components analysis for the first subset of variables for differences among subpopulations.

*a) Eigenanalysis of the covariance matrix of first subset of first-size variables*

Eigenvalue	1.7837	0.1849	0.0079	0.0019
Proportion	0.748	0.087	0.003	0.001
Cumulative	0.748	0.835	0.837	0.838
Variable	PC1	PC2	PC3	PC4
SW	-0.008	-0.000	-0.001	0.000
OW	-0.000	0.000	-0.000	-0.000
OCW	-0.000	-0.000	-0.000	-0.000
WIS	-0.007	0.000	0.000	-0.002

*b) Eigenanalysis of the covariance matrix of first subset of head-size variables*

Eigenvalue	0.000000	0.000000	0.000000	0.000000
Proportion	0.000	0.000	0.000	0.000
Cumulative	0.000	0.000	0.000	0.000
Variable	PC1	PC2	PC3	PC4
SW	-0.000	0.000	0.000	-0.000
OW	-0.000	-0.000	0.000	0.000
OCW	-0.000	-0.000	-0.000	0.000
WIS	-0.000	0.000	0.000	0.000

*c) Eigenanalysis of the covariance matrix of first subset of ratio variables*

Eigenvalue	0.000000	0.000000	0.000000	0.000000
Proportion	0.000	0.000	0.000	0.000
Cumulative	0.000	0.000	0.000	0.000
Variable	PC1	PC2	PC3	PC4
OCW	-0.000	0.000	-0.000	-0.000
OW/OW	0.000	0.000	-0.000	0.000
OCW	0.000	-0.000	-0.000	0.000
WIS	0.000	0.000	-0.000	-0.000

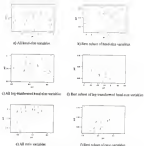


Figure 21. Morphometric differences among subpopulations. Plots of "valid regions" (area of range space) values for the two principal components (1: Anter. Diameter, 2: Postorb. diam.; 3: Femur, 4: Tibia, and 5: Dorsiflexor).

#### 4.4.2 Differences between Wild and Captive Animals

Behavioral changes are usually observed in wild animals captured in captivity. In 2004–2005, to evaluate capture efficiency of reintroduction program, three animals have to be "released" before released (Bou 1993). All morphological characteristics caused by capture stressors are frequently assessed as equal, they have been rarely quantified or even described in literature.

Ellison (1989) reports morphometric differences between wild and captive American alligators. According to him, the morphological changes caused by the captive environment could be deleterious for released animals. However, Ellap et al. (1992) reports that farm-released juveniles presented better performance than wild animals both in terms of growth rates and gonadal body condition.

Crocodilians are assumed to be opportunistic predators with an opportunistic role of parents in the guidance of hatchlings or if young about how to catch food, although in some circumstances, they may present distinctive methods of prey capture (Lang 1987). This apparently "innate" capacity seems to explain the results reported by Ellap et al. (1992). However, if on one hand, captivity seems not to affect crocodilian feeding behavior, plasticity, on the other hand, it may affect their reproductive capacity due to their differentiation in stress (Lowe 1987).

Evil eyes in Brazilian *Pseudocaiman* (United States) are easily recognized by local people because of their shape: long-headed eyespots, which gives them the same appearance as Brazilian Portuguese. This morphological characteristic, possibly related with feeding behavior, is more likely related to a high phenotypic plasticity due to genetic variation since the animals require a post- in the first instar. However, the most interesting signal of this pattern is that it provides a simple identification of this category of individuals, which is the object of the present study concerning wild and captive young WIC. For this purpose, capture individuals of the same range of body length (SVL) were compared with wild animals in terms of the morphometric variables showed in Figure 1 and described in Table 1 (Section 3.1). The questions and statistical methods showed in Figure 2 (Section 3.2) are presented and discussed below.

There is a significant difference between wild and captive animals in terms of their sexual complementarity ( $P$ -value < 0.005 for Wilk's, Levene-Browning, and Pillai's Tests of MANOVA) for the whole set of untransformed and log-transformed of size variables, and ratio-variables. The comparison between animals body-length (SVL) and complementarity is presented by head size and ratio variables in Table 6. Two of thirteen head-size variables (BL, LxS), and five of eleven ratio-variables (BL/W, BL/ST, BL/H, BL/SL, BL/M) expressed significant differences between wild and captive animals in covariance of SVL ( $P$ -value < 0.100).

Significant differences in MANOVA mean that animals from different environment present significant differences in the allocation resources between SVL and the morphometric variable in question. The biological meaning for this is a significant difference in sexual shape between wild and captive animals.

Captive animals presented higher average values for all significant variables described Table 8 (below- $P$  values = 0.01 in the analysis of variance- ANOVA). This means that captive animals present relatively broader snout and maxilla, longer mandibular symphysis (broader snout), and wider intermaxillaries. This is possibly not related to biomechanical stress and stress-tolerance level, as described by Ramirez (1980) for maxilla, since there is no any evidence recently of ground food. Instead, it is more likely caused by the food presentation of captive rodents, whereas in the captive animals longer for most of the time. The skull of a mouse is mainly strong to create a considerable pressure on its two lower mandible-cum-bone (lower mandible) described above.

Table 8. Analysis of variance: Intermaxillary (mid and upper)/an maxilla of WTG.

Variable	L <sub>1</sub>	Experiment	Variable	L <sub>1</sub>	Experiment
ICL	0.000	0.251	SLW	0.000	0.000
CM	0.000	0.270	SLST	0.000	0.207
SL	0.000	0.000	SMST	0.000	0.000
SM	0.000	0.000	BCL	0.000	0.000
CL	0.000	0.000	SLTW	0.000	0.000
OM	0.000	0.000	SLWT	0.000	0.000
ICWT	0.000	0.000	SLWS	0.000	0.000
LCS	0.000	0.000	SLWS	0.000	0.000
WS	0.000	0.000	SLWS	0.000	0.000
PCS	0.000	0.000	SLWS	0.000	0.000
LMS	0.000	0.000			
WMS	0.000	0.000			
LS	0.000	0.000			

Intermaxillary: mid = ANOVA in General Linear Model

Experiment (log ratios variables) symmetrical variables

Model (symmetrical variables) 100

Observation: Intermaxillary (mid and upper)/an maxilla and upper jaw length

The results above are somewhat similar to those described by Myers (1980) for captive *Arvicola elegans*. Many found that captive rodents present relatively shorter and broader snout. In the present study captive WTG presented wider a broader snout than wild individuals whereas no difference was found in terms of the relative snout-length between wild and captive animals. As mentioned by Myers, different biomechanical processes possibly underlie distinct results on the studying of the skull of wild and captive animals. Considerable size smaller average in mandible/broader maxilla (snout)/an maxilla (WTG).

Linear discriminant analysis with cross-validation for the whole set of body measurements (Table 8) and log-transformed head size variables (Table 9), and size variables

(Table B5) showed relatively high proportion of correct classification. Log transformation did not significantly improve results. After variables presented similar results to nontransformed base line variables. This is contrast to studies in the studies of log-transformation and correct classification. This can be explained by the analysis of covariance presented above (Table B6).

Table B7 Linear discriminant analysis: Response measurement: Prediction of log-transformed variables: Sample size: 100 animals: Groups: wild and captive animals

Summary of classification with log-transformed variables

Group	Captive	Wild
Captive	66	4
Wild	10	20
Total N	76	24
N Correct	66	23
Proportion	0.869	0.958

N = 100 N Correct = 89 Proportion Correct = 0.89

Multiple discriminant function for group

	Coefficient	SE	W	SE	DF	DF	DF	DF	DF	DF
Captive	-0.010	0.001	-0.01	0.00	1.00	1.00	1.00	1.00	1.00	1.00
Wild	0.000	0.00	0.00	0.00	1.00	1.00	1.00	1.00	1.00	1.00

	W	SE	W	SE
Captive	-0.01	0.00	-0.01	0.00
Wild	-0.01	0.01	-0.01	0.00

Table B8 Linear discriminant analysis: Response measurement: Prediction of log-transformed variables: Sample size: 100 animals: Groups: wild and captive animals

Summary of classification with non-transformed

Group	Captive	Wild
Captive	66	4
Wild	1	20
Total N	71	24
N Correct	66	23
Proportion	0.931	0.958

N = 95 N Correct = 89 Proportion Correct = 0.93

Multiple discriminant function for group

	Coefficient	SE	W	SE	DF	DF	DF	DF	DF	DF
Captive	0.000	0.000	0.00	0.00	1.00	1.00	1.00	1.00	1.00	1.00
Wild	0.000	0.000	0.00	0.00	1.00	1.00	1.00	1.00	1.00	1.00

	W	SE	W	SE
Captive	0.00	0.00	0.00	0.00
Wild	0.00	0.00	0.00	0.00



**Table 93** Linear discriminant analysis: Response: environment. Predictors: all-site variables. Sample size: 161 animals. Groups: n=81 and capture periods

in frequency of classification with zero variation

Group	Capture	NOR
Captive	65	4
Wild	16	13
Total	79	17
McNemar	65	14
Significance	0.002	0.001

$N = 161$ ,  $N \text{ Cases} = 161$ ,  $Frequency/Total = 0.164$

DT Linear discriminant function for group

	Constant	NOR	SE	SE	SE	SE	SE	SE	SE
Captive	-1042.5	144.3	609.1	154.3	1488.2	121.3	220.1	-1292.5	184.7
Wild	-1754.9	-184.7	609.1	154.3	1471.3	126.8	262.5	-1889.9	141.2

	SE	SE
Captive	611.3	121.3
Wild	144.3	184.7

PCA of untransformed (Table 95) and log-transformed head-size variables (Table 97) presented similar results. PC1 accounted for most of the total variance in both cases (31.9% for the former and 21.9% for the latter). Log-transformation was responsible for a significant concentration of variance on the first remaining principal component (PC2). It responded for more or much of the variation contained in PC2 of untransformed variables (2.9% against 0.9%). In this case, therefore, log-transformation improved separability of the model to possess shape variation between groups.

PC1 coefficients of untransformed variables are similar in magnitude and sign, varying from -0.287 for NOR to -0.251 for CW. PC1 coefficients of log-transformed variables are also similar in sign, but presented a larger range, varying from 0.134 for CW to 0.362 for NOR. This seems that most of the internal morphometric variation between individuals in the overall case of skull.

PC2 of untransformed variables showed a clear contrast between CW (-0.606) and NOR (0.762), whereas CW alone responds for more of the variance of PC2 for log-transformed variables (0.449). These patterns are similar to the results of principal components analyses of the whole sets of variables for captive and wild animals captured (see Tables 8.9 and 25). They are also similar to the PCA of the two subsets of variables for sexual dimorphism in captive animals (see Figure 10). However, they slightly vary from the patterns showed by the PCA for

the best subsets of variables for age determination using any method. A post-hoc test for the better gender and age effects are not related from environmental effects in the present study, although gender effect seems to be stronger in this case, most of animals are young.

**Table 14: Principal components analysis: Eigenanalysis of the correlation matrix of all local-area variables. Sample size: 75 wild animals. Only six first principal components are presented**

Component	1 (20%)	2 (16%)	3 (14%)	4 (10%)	5 (10%)	6 (14%)
Proportion	0.405	0.307	0.251	0.165	0.141	0.208
Cumulative	0.405	0.712	0.963	0.978	0.995	0.999
Variable	PC1	PC2	PC3	PC4	PC5	PC6
OC	-0.289	-0.664	-0.177	0.236	-0.134	0.111
OW	-0.194	-0.582	-0.134	-0.388	0.173	-0.127
SL	-0.183	-0.604	-0.411	0.329	-0.208	0.048
SW	-0.264	-0.664	-0.183	0.263	-0.043	-0.177
OL	-0.184	-0.587	-0.477	0.134	-0.389	-0.138
OW	-0.238	-0.544	-0.563	0.144	-0.367	-0.048
OCW	-0.167	-0.545	-0.352	0.147	-0.331	-0.073
LCS	-0.177	-0.608	0.075	0.131	-0.388	0.131
WS	-0.178	-0.664	-0.166	-0.432	0.046	-0.078
PS	-0.175	-0.612	0.211	-0.138	-0.298	0.038
SL	-0.188	-0.606	-0.408	0.338	-0.207	-0.092
LWS	-0.278	-0.637	-0.087	-0.322	-0.248	0.107
WSL	-0.282	-0.634	-0.134	-0.327	-0.247	-0.214

**Table 15: Principal components analysis: Eigenanalysis of the covariance matrix of all log-transformed local-area variables. Sample size: 75 wild animals. Only six first principal components are presented**

Component	1 (46.6%)	2 (24.4%)	3 (16.1%)	4 (10.7%)	5 (9.0%)	6 (10.2%)
Proportion	0.466	0.244	0.161	0.107	0.090	0.102
Cumulative	0.466	0.710	0.871	0.978	0.997	0.999
Variable	PC1	PC2	PC3	PC4	PC5	PC6
LWS	-0.208	-0.588	0.183	0.133	0.136	-0.157
OW	-0.187	-0.564	-0.264	-0.264	-0.049	0.186
SL	-0.222	-0.603	-0.388	0.175	0.054	-0.134
SW	-0.208	-0.583	-0.264	-0.268	-0.066	0.077
OL	-0.188	-0.534	-0.322	-0.343	0.175	0.023
OW	-0.194	-0.588	-0.265	-0.270	-0.088	0.054
OCW	-0.262	-0.564	-0.388	-0.268	-0.073	-0.077
LCS	-0.175	-0.688	0.168	0.163	0.164	-0.073
WS	-0.213	-0.583	-0.268	-0.268	-0.043	0.167
PS	-0.163	-0.544	-0.379	-0.237	0.028	-0.076
SL	-0.184	-0.568	-0.383	0.171	0.088	-0.124
LWS	-0.198	-0.544	-0.434	-0.164	0.078	-0.038
WSL	-0.178	-0.575	-0.364	-0.167	0.077	-0.076

**Table VI** Principal components analysis. Eigenanalysis of the covariance matrix of all variables. Sample size: 25 wild animals. Only the first principal components are presented

Variables	1. 568 (72)	2. 353 (57)	3. 260 (34)	4. 164 (15)	5. 140 (10)	6. 100 (8)
Population	0.344	0.266	0.184	0.001	0.000	-0.012
Cardioides	0.369	-0.070	-0.028	0.080	0.004	-0.016
Forams	0.01	0.01	0.01	0.01	0.01	0.01
BCW	0.001	0.062	0.038	-0.017	-0.003	-0.010
BAW	0.026	0.012	0.009	-0.001	-0.007	-0.008
BWWT	-0.798	-0.001	0.000	0.000	-0.000	0.000
BCL	0.000	0.000	-0.000	-0.000	0.000	0.000
BCLW	0.000	0.000	0.000	-0.000	-0.000	-0.000
BWL	0.000	0.000	0.000	0.000	0.000	0.000
BWBL	0.000	0.000	0.000	-0.000	-0.000	-0.000
BAWBL	0.000	0.000	0.000	-0.000	-0.000	-0.000
BAWBL	0.000	0.000	0.000	-0.000	-0.000	-0.000
BAWBL	0.000	0.000	0.000	-0.000	-0.000	-0.000
BAWBL	0.000	0.000	0.000	-0.000	-0.000	-0.000

PCA of nine variables shows a more complex pattern: with PC1 and PC2 accounting for respectively 8.56% and 5.56% of the total variance. BWWT (0.798) is the major source of variation in PC1, whereas BCLW (0.062) is the major source of variation on PC2. Both values are non-dependent (see Table 2). Coefficients of PC1 and PC2 vary both in sign and magnitude. There is a clear contrast in PC1 between BWWT and the remainder between BCLW (p. 24) and BW (p. 44). This possibly represents a simplification of the mean (BWWT with negative sign) as reference for type 1 region of the shell (BCLW and BWL with positive sign). This is compatible with the results of ANCOVA (Table III) and subsequent ANCOVA, in which wild animals present a relatively narrower mean.

Table VI presents the first subsets of variables for the study of morphometric differences between wild and captive animals. Log transformation did not significantly affect the model. Both untransformed and log transformed data presented the same first subsets of variables: equally divided into width- and length-variables. The first subset of nine variables, however, included only width-variables, which is compatible to the ANCOVA results (Table III), discussed above.

**Table VII** First subsets of morphometric variables for comparisons (wild and captive)

Species	First subsets	F value	p	n
Wild-cat	1/1W, 1/2W, 3/3W, 3/4W	0.008	0.560	124
Log transformed wild-cat	1/1W, 1/2W, 3/3W, 3/4W	0.008	0.524	124
Sum	BWWT, BCL, BCLW, BCLBL	0.000	0.143	111

Model presented: test on logarithms = Bonferroni

There is still a significant difference in the model homogeneity of wild and capture animals in relation to the number of variables ( $P$  values < 0.000 for Wilk's, Lawley-Hotelling, and Pillai's Tests of MANOVA, for different values of variables). Linear discriminant analyses with cross-validation for the first subset of both transformed (Table 14) and log-transformed land-use variables (Table 15) and time variables (Table 16) presented even higher proportions of correct classifications. Log transformation did not significantly improve results. Both variables presented intermediate results (89% corrects, respectively 93.10%) and log-transformed land-use variables (9.11%). This is unusually high for other variables, as they in former sections, but can be explained by the results of the analysis of homogeneity (Table 14).

Table 14. Linear discriminant analysis. Response environment. Predictors: first subset of land-use variables. Sample size: 103 animals. Groups: wild and capture animals.

a) Summary of classification with cross-validation

Group	Captured	Wild
Captured	85	3
Wild	1	20
Total N	86	23
N Correct	85	20
Proportions	0.978	0.870

N = 103 N Correct = 95 Proportions Correct = 0.923

b) Linear discriminant function for group

	Coefficient	CV	LCV	WC	WC
Capture	-20.144	-21.641	0.110	1.657	-4.334
Wild	20.144	18.426	-0.087	-1.657	4.334

Table 15. Linear discriminant analysis. Response environment. Predictors: first subset of log-transformed land-use variables. Sample size: 103 animals. Groups: wild and capture animals.

a) Summary of classification with cross-validation

Group	Captured	Wild
Captured	85	3
Wild	1	22
Total N	86	25
N Correct	86	22
Proportions	0.978	0.880

N = 111 N Correct = 107 Proportions Correct = 0.963

b) Linear discriminant function for group

	Coefficient	CV	LCV	WC	WC
Capture	-1.474.4	-1471.9	1.041.3	-1704.2	104.3
Wild	1.474.4	1471.9	-1.041.3	1704.2	-104.3

Table 16: Latent structural analysis: Asymptotic covariance matrix. First column of table variables. Sample size: 182 animals. Groups: wild and captive groups.

(f) Summary of latent variables with their solutions

Group	Latent	Wald
Captive	6.7	1
Wild	7	29
Total N	14	29
N Varied	67	20
Determinant	1.581	0.042

$N = 182$      $N$  Censored = 15     $\chi^2$  Probability (chi-sq) = 0.000

(g) Latent structure: loadings for group

	Captive	B.WET	B.WT	B.WD	B.LM
Captive	-0.82 (8)	0.86 (1)	0.93 (0)	0.93 (0)	0.93 (0)
Wild	-0.71 (9)	0.88 (0)	0.93 (0)	0.93 (0)	0.93 (0)

PCA of the first subspace of untransformed (Table 15a) and log-transformed morphological variables (Table 15b) present similar results. PC1 accounts for most of the variation in both groups (WET) for the captive and B.WD for the free). With the reduction in the number of variables, log transformation no longer is responsible for a significant proportion of variation in PC2 (3–23% versus 4–16% in PC2 for untransformed variables).

PCA of both untransformed and log-transformed variables show a clear contrast between LCR and B.WD. B.WD also contributes significantly for the variation after log transformation of data. In these variables the management in the eye region of the head is unique to its width – probably more associated with morphologic processes and sexual dimorphism – is not present. This may be a good reflective character for processes were significantly isolated from the analysis of the environmental effect, after reducing the number of variables. A relationship between the width of head and the length of head from base, represented by WM and LCR, seems to be the major source of variation between wild and captive animals.

PCA of the first column of table variables (Table 16) presented as usual a first principal component of variation in PC1, and a significant source of variation in PC2 and the remaining principal components (3–18%, 8.34%, and 4.14% respectively). The major source of variation in PC1 is B.WET (8.78%), which seems to complement the pattern described above. However, the relative interaction with B.WD also contributes significantly for the total amount of variation contained in PC1. No contrast is perceivable in PC2, but its major source of variation is B.WD, which is a clear independent morphometric variable (see Table 6). This probably explains a

shape" component in PC2 and also means that the "size" component in PC1 is primarily due to BWST.

**Table IV** | *Loadings between wild and captive animals: Principal components analysis of best subsets of variables. Sample size = 103 animals*

**ii) Eigenvalues of the principal scores of the best subset of best-subset variables**

Component	1 (73%)	2 (14%)	3 (2%)	4 (0%)
Proportion	0.730	0.040	0.010	0.020
Cumulative	0.730	0.770	0.780	0.800
Variable	PC1	PC2	PC3	PC4
CR	-0.512	0.138	0.007	-0.036
SCB	-0.495	-0.125	0.081	0.271
SW	-0.495	-0.078	0.112	0.066
ML	-0.258	-0.104	-0.054	0.077

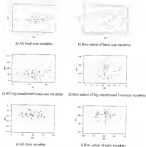
**iii) Eigenvalues of the principal scores of the best subset of the transformed best-subset variables**

Component	1 (50.4%)	2 (40.4%)	3 (3.4%)	4 (5.8%)
Proportion	0.504	0.404	0.034	0.058
Cumulative	0.504	0.908	0.942	1.000
Variable	PC1	PC2	PC3	PC4
Low	-0.503	0.000	-0.503	0.000
Like	-0.183	-0.568	-0.000	-0.277
Low	-0.508	0.000	-0.503	-0.000
Low	0.503	-0.503	0.000	-0.000

**iv) Eigenvalues of the principal scores of the best subset of raw variables**

Component	1 (60.000%)	2 (32.000%)	3 (0.000%)	4 (0.000%)
Proportion	0.600	0.320	0.000	0.000
Cumulative	0.600	0.920	0.920	0.920
Variable	PC1	PC2	PC3	PC4
SWST	0.700	-0.700	-0.000	0.000
SW	-0.700	-0.000	0.000	0.000
CR	-0.100	-0.100	0.000	0.000
SCB	0.000	-0.000	0.000	0.000

Figure 12 shows the plots of individual scores versus against values for the two principal components. It is clear that the use of the best subset of variables obtained for comparing between wild and captive animals (compare plot (a) and (c) respectively with a and c). It is also clear that this variation is basically due to shape differences between wild and captive animals (compare plot (b) and (d) that PC2 scores the importance of variation). This means that wild and captive animals with similar skull shape present a significant difference in the skull shape.



**Figure 22** Differences between related species across: (Row) all variables/first subset of variables; (Column) all size variables/first subset of size variables. (a) head, (c) leg, (e) size.

## CHAPTER 5 SUMMARY AND CONCLUSIONS

Three concrete questions are asked in the present morphometric study of the long-tailed coelacanth (*C. concolor* Latimer), Is there a significant morphometric variation caused by sexual-effect, Age-effect, sex-age, size of origin, and environment? Can the group of variables chosen to which a specific individual belongs be predicted? Lastly, is it possible to reduce the number of variables without significantly losing efficiency?

Cranial and body morphometric variables were recorded from light and wet mounts. Linear and log-transformations of data are presented and discussed. Multivariate linear discriminant analysis with various features, PCA, and least-squares regression are used in order to answer the questions above. Exploratory studies of relative growth and efficiency of reproduction are used to develop models of age estimation, sexual dimorphism, and clutch-size estimation. Comparing changes along major variation in upper jaw/fin rays which may be associated with sexual sexual signaling, sexual maturity, is the sense that these are the regions of the body mostly exposed above surface of the water. The patterns of sexual relative growth found suggests a possible relationship between sexual behavior and morphometric changes.

Energy for sexual dimorphism studies, and also for the specific case of body mass (BM) or growth curves and subsequent efficiency, log-transformations always significantly improve data analysis in the present study. Ratio-variables, as there were, were consistently less efficient in detecting differences between groups than head-size variables and consistently produced smaller patterns of relative distribution in principal-component analyses. The "shape" information contribution ratio may be substituted by the better quantification of shape and size variation of principal-component analyses of linear measurements. Table 5B presents the best subsets of cranial morphometric variables for age estimation, sexual dimorphism, clutch-size estimation, non-linear k-sample (plants theory) effect, geographic variation (effect of size of sample) and environmental effect (differences between wild and captive animals).

Multivariate morphometric models are consistently more efficient than univariate growth curves to determine age of animals. Finest head length (FHL) or body mass (BM) alone is superior (BIC) or less-efficient than other single head measurements, like dorsal-cranial length (DCL).





BAC have positive correlation between female body size (BM and SVL) and egg size (area, width, and length), hatchlings body size (BM and SVL), and clutch mass. There was significant negative correlation between female body size (BM and SVL) and the relative clutch mass (relative to female BM). The percentage in clutch mass drops from approximately 10%, when female BM increases from 20 to 40 kg. However, no significant correlation was found between hatchlings size (BM and SVL) and clutch size.

There are consistent clutch effect on egg shape and size, and hatchlings cranial morphology. Based on test to detect egg measurements (area, width, and length) as increasing order of importance, it is possible to detect female eggs from different clutches with consistently high inter-clutch (99%), area is homogeneous groups of related reproductive females. Similarly it is possible to discriminate hatchlings from different clutches with area higher inter-clutch (88%) with the relative set and approximately 10% with related set of four "best" cranial morphometric variables).

Comparative clutch effect, individual changes have even more consistent effect on the shape and size of eggs, and on the cranial morphology of hatchlings. Even in a small group of related reproductive females, it was possible to discriminate clutches/eggs and hatchlings from different female/families with relatively 100% chance (approximately 94% for eggs, and three for hatchlings, with first subset of four cranial morphometric variables). Clutch discrimination based on clutch effect and female family effect (individual changes) can be used to establish percentage survival in females in small wild populations and captive breeding programs. For the first time, these conclusions be used in morphological over-exploited freshwater tropical pond.

There is a significant cranial morphometric variation among animals from different geographic sites of origin. The present study sites are not necessarily far from each other, and are interconnectivity waterways. The possible morphometric differences between these subpopulations may be related to some level of reproductive isolation, which may related to female related factors (reproductive). Claves actually possess relatively broader snout and anterior longer mandibular symphysis, located anterior and wider interorbital space that will contribute possibly due to less overlapping caused by the fast progression of capture-mechanism rather than difference in feeding habit.

Although positive morphometric led to significant relevance on the graphed representation of being reproductive form, individual male-specific morphometrics can still be extremely useful for discrimination studies. Sex-only interspecific variation – the traditional

field of morphometrics – for the intraspecific morphometric variation can be successfully related by multivariate morphometrics. In the present study, multivariate analyses (morphometric analysis) are used to evaluate the effect of (age, ontogenetic growth, sexual dimorphism, clutch effect, maternal lineage – geographic variation, and environment in RAC) morphology.

The morphometric model developed in the present study might represent an appropriate (and effective) set of methods for the metapopulation modeling and conservation of the local extant colonies in São Paulo, Brazil. The conservation of local populations in São Paulo may be essential to provide geographic isolation between southern and southern populations of the species according to overall geographic distribution.

The assessment of individual's sex and age might lead to the establishment of the population's sex ratio and age structure. The most behavior of juvenile among individuals based on clutch characteristics might lead to the assessment of nesting system and dispersal pattern. Consequently, it might result in a better understanding of the relationship (local-population) evolutionary process of autochthonous and non-native behavior of the species in São Paulo as a metapopulation context.

Understanding how an endangered species respond to anthropogenic pressure may help us to predict other species responses in similar situations. In a certain sense, what is happening in São Paulo today, both in terms of the human population wildlife species and habitats and their response to this impact, as likely, is happens in the future in other regions of Brazil. In other words, understanding how some wild species "behave" or how well or might help to learn how to live with them.

**Appendix A**  
**Descriptive Statistics of**  
**Metropolitan Variables**



Table 1. *Continued*

Variable	Mean	SD	n	Level
Age	38.1	10.2	38	1
Gender	1.0	0.0	38	1
Marital status	1.0	0.0	38	1
Education	1.0	0.0	38	1
Occupation	1.0	0.0	38	1
Income	1.0	0.0	38	1
Health status	1.0	0.0	38	1
Stress level	1.0	0.0	38	1
Life satisfaction	1.0	0.0	38	1
Self-esteem	1.0	0.0	38	1
Resilience	1.0	0.0	38	1
Optimism	1.0	0.0	38	1
Emotional stability	1.0	0.0	38	1
Neuroticism	1.0	0.0	38	1
Extraversion	1.0	0.0	38	1
Agreeableness	1.0	0.0	38	1
Conscientiousness	1.0	0.0	38	1
Openness to experience	1.0	0.0	38	1
Intelligence	1.0	0.0	38	1
Memory	1.0	0.0	38	1
Attention	1.0	0.0	38	1
Perception	1.0	0.0	38	1
Reasoning	1.0	0.0	38	1
Problem-solving	1.0	0.0	38	1
Decision-making	1.0	0.0	38	1
Communication	1.0	0.0	38	1
Interpersonal skills	1.0	0.0	38	1
Leadership	1.0	0.0	38	1
Teamwork	1.0	0.0	38	1
Conflict resolution	1.0	0.0	38	1
Stress management	1.0	0.0	38	1
Emotional regulation	1.0	0.0	38	1
Self-regulation	1.0	0.0	38	1
Goal setting	1.0	0.0	38	1
Time management	1.0	0.0	38	1
Organization	1.0	0.0	38	1
Productivity	1.0	0.0	38	1
Efficiency	1.0	0.0	38	1
Quality of work	1.0	0.0	38	1
Job satisfaction	1.0	0.0	38	1
Work-life balance	1.0	0.0	38	1
Overall well-being	1.0	0.0	38	1

Table 2. *Continued*

Variable	Mean	SD	n	Level
Age	38.1	10.2	38	1
Gender	1.0	0.0	38	1
Marital status	1.0	0.0	38	1
Education	1.0	0.0	38	1
Occupation	1.0	0.0	38	1
Income	1.0	0.0	38	1
Health status	1.0	0.0	38	1
Stress level	1.0	0.0	38	1
Life satisfaction	1.0	0.0	38	1
Self-esteem	1.0	0.0	38	1
Resilience	1.0	0.0	38	1
Optimism	1.0	0.0	38	1
Emotional stability	1.0	0.0	38	1
Neuroticism	1.0	0.0	38	1
Extraversion	1.0	0.0	38	1
Agreeableness	1.0	0.0	38	1
Conscientiousness	1.0	0.0	38	1
Openness to experience	1.0	0.0	38	1
Intelligence	1.0	0.0	38	1
Memory	1.0	0.0	38	1
Attention	1.0	0.0	38	1
Perception	1.0	0.0	38	1
Reasoning	1.0	0.0	38	1
Problem-solving	1.0	0.0	38	1
Decision-making	1.0	0.0	38	1
Communication	1.0	0.0	38	1
Interpersonal skills	1.0	0.0	38	1
Leadership	1.0	0.0	38	1
Teamwork	1.0	0.0	38	1
Conflict resolution	1.0	0.0	38	1
Stress management	1.0	0.0	38	1
Emotional regulation	1.0	0.0	38	1
Self-regulation	1.0	0.0	38	1
Goal setting	1.0	0.0	38	1
Time management	1.0	0.0	38	1
Organization	1.0	0.0	38	1
Productivity	1.0	0.0	38	1
Efficiency	1.0	0.0	38	1
Quality of work	1.0	0.0	38	1
Job satisfaction	1.0	0.0	38	1
Work-life balance	1.0	0.0	38	1
Overall well-being	1.0	0.0	38	1







**APPENDIX B**  
**CORRELATION MATRIX (PEARSON)**  
**OF MEASURED/STUDY VARIABLES**

[illegible]





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I certify that I have read this study and that in my opinion it conforms to acceptable standards of scholarly presentation and is fully adequate, in scope and quality, as a dissertation for the degree of Doctor of Philosophy.

  
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